# Growth and Convergence across the U.S.: Evidence from County-Level Data\*

Matthew J. Higgins College of Management Georgia Institute of Technology Atlanta, GA 30308-0520 mhiggins@yakko.mgt.gatech.edu Tel: 404.894.4368, Fax: 404.894.6030

Daniel Levy Department of Economics, Bar-Ilan University Ramat-Gan 52900, Israel Levyda@mail.biu.ac.il Tel: +972-3-531-8331, Fax: +972-3-535-3180

Andrew T. Young Department of Economics, University of Mississippi University, MS 38677-1848 Atyoung@olemiss.edu Tel: 662-915-5829, Fax: 662-915-6943

Last Revision: August 8<sup>th</sup>, 2005 JEL Codes: O40, O11, O18, O51, R11, H50, and H70 Key Words: Economic Growth, Conditional Convergence, and County-Level Data

Forthcoming in: Review of Economics and Statistics, 2006.

<sup>&</sup>lt;sup>\*</sup> We thank two anonymous referees for constructive comments, and the editor, Dani Rodrik for guidance. We are grateful also to Paul Evans for answering our questions and for commenting on the previous version of the paper, and to Jordan Rappaport for kindly sharing with us his data and computer codes and for providing helpful suggestions throughout the study. Participants at the University of Mississippi seminar series provided helpful discussion on this research. Finally, we thank Bob Barsky, Bob Chirinko, Eric Hallerberg, Jesse Hamner, Nazrul Islam, Joy Mazumdar, Paul Rubin, Jacques Silber, and Som Somanathan for helpful conversations. All errors are ours.

# Growth and Convergence across the U.S.: Evidence from County-Level Data

1

<u>Abstract</u>: We use U.S. county data (3,058 observations) and 41 conditioning variables to study growth and convergence. Using OLS and 3SLS-IV we report on the full sample and metro, non-metro, and 5 regional samples: (1) OLS yields convergence rates around 2 percent; 3SLS yields 6–8 percent; (2) convergence rates vary (e.g., the Southern rate is 2.5 times the Northeastern rate); (3) federal, state and local government negatively correlates with growth; (4) the relationship between educational attainment and growth is nonlinear; and (5) finance, insurance & real estate industry and entertainment industry positively correlates with growth while education employment negatively correlates.

Last Revision: August 8th, 2005

JEL Codes: O40, O11, O18, O51, R11, H50, and H70

Key Words: Economic Growth, Conditional Convergence, and County-Level Data

# I. Introduction

We study growth determination and measure the speed of income convergence within the U.S. In so doing, we make four contributions to the empirical economic growth literature.<sup>1</sup> First, we have assembled unusually rich county-level data. In contrast to 100–150 observations (typical for existing cross-country, -state, and -regional data sets) our data contain 3,058 observations. The U.S. county data are collected by a single institution using uniform variable definitions. Also, there is no exchange rate variation between the counties and the price variation across counties is smaller than across countries. Furthermore, U.S. counties are far more homogeneous than countries.<sup>2</sup>

Second, the large number of cross-sectional observations allows us to study not only the full sample, but also regional groups (Northeast, Great Lakes, West, Plains and South), and metro and non-metro groups to control for possible cross-regional heterogeneity. Heterogeneity can exist in convergence parameters and also parameters governing the effect of conditioning variables on the level of the balanced growth path.

Third, we use 41 different conditioning variables to assess the empirical relevance of various determinants of balanced growth path positions. Previous cross-country studies, taken together, have considered as many as 90 different variables as potential growth determinants (Durlauf and Quah, 1999; Durlauf, 2001). As Brock and Durlauf (2001, p.7) emphasize, however, there are "at best about 120 countries' data available for analysis in cross-sections [and therefore] it is far from obvious how to formulate firm

<sup>&</sup>lt;sup>1</sup> The seminal studies in this area are Barro and Sala-i-Martin (1991, 1992) and Mankiw et al. (1992). Quah (1996) and Sala-i-Martin (1996) survey some of the literature that followed. For a more recent survey, see Brock and Durlauf (2001).

<sup>&</sup>lt;sup>2</sup> Many of these virtues are embodied in state-level data used by, Barro and Sala-i-Martin (1991) and Evans (1997a). However, state-level data sacrifices the large number of observations that we have.

inferences about any particular explanation of growth." Given our large number of crosssectional observations, we can use the full set of conditioning variables included in our data and still obtain precise estimates of the coefficients.

Fourth, in estimation we employ a cross-sectional variant of Evans' (1997a, 1997b) 3-stage least squares (3SLS) approach, as well as ordinary Least squares (OLS). Evans (1997b) shows that for the consistency of OLS estimates the data must satisfy highly implausible conditions. He proposes a 3SLS-instrumental variables (IV) method, which produces consistent estimates.

This paper is organized as follows. In section II we discuss the econometric model. In section III we describe the data. In section IV we present the findings regarding the conditional convergence rates, followed by the findings regarding balanced growth path determinants in section V. We conclude in section VI.

## II. Econometric Model Specification and Estimation

The neoclassical growth model implies that  $\hat{y}(t) = \hat{y}(0)e^{-Bt} + \hat{y}^*(1 - e^{-Bt})$ , where  $\hat{y}$  is log of income per effective unit of labor, *t* is the time period, and *B* is a nonlinear function of various parameters (population growth rate, preference parameters, etc.). *B* governs the speed of adjustment to the steady state while  $\hat{y}^*$  denotes the steady state. Thus, the average growth rate of income per unit of labor between dates 0 and *T* is

(1) 
$$\frac{1}{T} \left( y(T) - y(0) \right) = z + \left( \frac{1 - e^{-BT}}{T} \right) \left( \hat{y}^* - \hat{y}(0) \right),$$

where z is the exogenous rate of technological progress and B measures the sensitivity of the average growth rate to the gap between the steady state and the initial value. Since

effective unit of labor (L) is assumed to equal  $Le^{zt}$ , we have  $\hat{y}(0) = y(0)$ .

Growth regressions are obtained by fitting to the cross-sectional data the equation

(2) 
$$g_n = \alpha + \beta y_{n0} + \gamma x_n + v_n,$$

where  $g_n$  is the average growth rate of per capita income for economy *n* between years 0 and *T* [i.e., (y(T) - y(0))/T],  $\alpha$  is a constant representing *z*,  $\beta = (1 - e^{-BT})/T$ ,  $x_n$  is a vector of variables that control for cross-economy heterogeneity in determinants of the steady-state,  $\hat{y}^*$ ,  $\gamma$  is a vector of coefficients, and  $v_n$  is a zero mean-finite variance error.

OLS can then be used to infer the values of  $\beta$  and  $\gamma$  in (2) by regressing the growth rate on initial values of per capita income and other conditioning variables. However, Evans (1997b) shows that for the consistency of OLS estimates, the data must satisfy highly implausible conditions, and argues that plausible departures from them can produce large biases. Specifically, he shows that unless (i) the dynamic structures of the economies studied have identical AR(1) representations, (ii) every economy affects every other economy symmetrically, and (iii) conditioning variables control for all permanent cross-economy differences, the OLS estimates of the speed of convergence are inconsistent—they are biased downwards.<sup>3</sup>

Evans (1997b) proposes a 3SLS-IV approach, which produces consistent estimates. The first- and second-stages involve using instrumental variables (IVs) to estimate the regression

(3) 
$$\Delta g_n = \omega + \beta \Delta y_{n0} + \eta_n,$$

<sup>&</sup>lt;sup>3</sup> These results, due to Evans (1997b), are included in an appendix, available from the authors.

where  $\Delta g_n = [(y_{n,T} - y_{n,0})/T] - [(y_{n,T-1} - y_{n,-1})/T]$ ,  $\Delta y_{n0} = y_{n0} - y_{n,-1}$ ,  $\omega$  and  $\beta$  are parameters, and  $\eta_n$  is the error. As instruments we use the lagged values of  $x_n$  variables.<sup>4</sup> Given our sample period, we define  $\Delta g_n = [(y_{n,1998} - y_{n,1970})/T] - [(y_{n,1997} - y_{n,1969})/T]$ .

We use  $\beta^*$ , the estimate from (3), to construct the variable  $\pi_n = g_n - \beta^* y_{n0}$ , which is regressed on the vector  $x_n$ . Thus, the third-stage regression is of the form

(4) 
$$\pi_n = \tau + \gamma x_n + \varepsilon_n,$$

where  $\tau$  and  $\gamma$  are parameters and  $\varepsilon_n$  is the error. OLS yields a consistent estimator,  $\gamma^*$ .

Of note,  $\gamma$  is not technically the partial effects of  $x_n$  variables on the heights of the balanced growth paths. Those partial effects are functions of  $\beta$  as well as  $\gamma$ . However, if the neoclassical (exogenous) growth hypothesis is true ( $\beta < 0$ ), then signs of elements of  $\gamma$  will be the same as those of the partial effects of given  $x_n$  elements. As well, given the assumption that  $\beta$  is identical across economies, the magnitude of  $\gamma$  elements *relative* to one another expresses the magnitudes of the partial effects *relative* to one another. Thus, while  $\gamma^*$  does not allow for precise quantitative statements about the effects of given conditioning variables on balanced growth paths, it does allow for statements about the sign of such effects, as well as how important those effects are relative to each other.

To summarize, we use a three-stage procedure. In the first- and second-stages, we

<sup>&</sup>lt;sup>4</sup> These are the 1969 values of conditioning variables with the exception of Metro Area, Water Area, and Land Area. See the data appendix for details and Table 1 for a list of our conditioning variables.

difference out any uncontrolled form of heterogeneity to eliminate omitted variable bias.<sup>5</sup> In the third stage, the estimate of  $\beta$  is used to recreate the component of a growth regression that would be related to conditioning variables. This component is regressed on a constant and the conditioning variables, in "un-differenced" form, to estimate the partial correlations of conditioning variables with the growth rate. This procedure ensures that none of the information contained in the levels of the conditioning variables is lost.<sup>6</sup>

We use a Hausman test to determine the appropriateness of the IV approach. Two tests – the first run on the  $\beta^*$  values and the second on the entire model – yielded *m* values of 134.6 and 1236.6 respectively. Both tests reject the null at the 1-percent level, suggesting that the OLS estimates are inconsistent, confirming the importance of using  $\overline{}^5$  Derivation of (3) assumes constancy of the conditioning variables, allowing them to be differenced out. Nazrul Islam has noted that while this might hold for, e.g., an index of democracy for an international sample over 15 years, some of the county-level conditioning variables could potentially vary. To make sure that this did not introduce significant omitted variable bias, we ran the three IV regressions for the full U.S., metro U.S. and non-metro U.S. with differenced values of all conditioning variables included as regressors. All point estimates of  $\beta$  from the modified IV regressions fell within the 95 percent confidence intervals of the Evans method IV estimates. As well, if the  $\beta$  estimates are not significantly affected then neither are the third-stage results.

<sup>6</sup> Following a referee's suggestion, we have estimated the model using a panel-GMM method as well. However, the resulting estimates, which we generated using the method of Caselli, et al. (1996), did not make much sense. We believe the main reason for the failure of the panel-GMM approach is that it may be ill-suited for our data because our sample does not form a "true panel." Although we have over 3,000 crosssectional observations, over time we only have 3 time series observations (the 1970, 1980, and 1990 decennial Census data) and it appears that it is not enough to carry the level-information forward after the variables are differenced, which is necessary for implementing panel-GMM estimation. This is a point on which Barro (1997, p. 37) has criticized panel data methods. As they rely on time series information, the conditioning variables are differenced. However, the conditioning variables often vary slowly over time so that the most important information is in the levels.

7

the IV method for addressing the potential endogeneity of conditioning variables.<sup>7</sup>

To account for a possible spatial correlation between the error terms of the counties located in a proximity with each other, we follow Rappaport and Sachs (2003) by reporting a generalization of the Huber-White heteroskedastic-consistent estimator based on Rappaport's (1999) implementation of Conley's (1999) correction to obtain standard errors that are robust to such a spatial correlation. Rappaport and Sachs specify a cutoff distance,  $\vec{d}$ , and assume that the covariance between the errors of two counties is zero, if the Euclidean distance between the counties centers exceeds  $\vec{d}$ . Otherwise, they impose declining weight structure on the covariance by defining a distance function  $g(d_{ij}) = 1 - (d_{ij}/200)^2$ , where  $d_{ij}$  is the distance between the centers of counties *i* and *j*, and assuming that  $E(\varepsilon_i \varepsilon_j) = g(d_{ij})\rho_{ij}$ , where  $\hat{\rho}_{ij} = e_i e_j$ , and  $g(d_{ij}) = 1$  for  $d_{ij} = 0$ ,

$$g(d_{ii}) = 0$$
 for  $d_{ii} > \overline{d}$ , and  $g'(d_{ii}) \le 0$  for  $d_{ii} \le \overline{d}$ . Thus, Rappaport's (1999)

implementation of the correction assumes that the covariance between the error terms falls off quadratically as the distance between the counties increases to  $\overline{d} = 200$ km. The corrected standard errors are used in calculating the confidence intervals reported under the CR (Conley-Rappaport) column in Tables 2 and 3.<sup>8</sup> In sum, we present three sets of

<sup>&</sup>lt;sup>7</sup> It may be argued that some of the variables we use, such as educational variables are endogenous, reflecting perhaps institutional and cultural features that lead to demand for various levels of schooling in various counties. While this might be the case, we believe the problem is unlikely to be severe. This is because in the model we estimate, the RHS variables are temporally prior to the regressor. Also, we use instrumental variables to resolve whatever endogeneity problem might still be there. Finally, we used Hausman test to check for—and confirmed—the appropriateness of the instrumental variables approach. <sup>8</sup> We are grateful to Jordan Rappaport for sharing with us his computer codes and for helping us in implementing the CR correction.

estimates: OLS, CR-OLS, and 3SLS.9

#### III. U.S. County-Level Data

The data we use are drawn from several sources but the majority comes from the Bureau of Economic Analysis Regional Economic Information System (BEA-REIS) and US Census data sets. The BEA-REIS data are largely based on the 1970, 1980 and 1990 decennial Census files; the 1972, 1977, 1982 and 1987 Census of Governments; and the Census Bureau's City and County Book from various years. We exclude military personnel from the measurements of both personal income and population.

Our data contain 3,058 county-level observations. The large number of observations allows us to explore possible heterogeneity across the U.S regions by splitting the data into two sets of sub-samples. The first set separates the data into 867 metro and 2,191 non-metro counties (Figure 1).<sup>10</sup> The second set separates the data into five regions: Northeast, Great Lakes, West, Plains, and South. Given the large sample, the sub-sample analysis sacrifices little in terms of degrees of freedom. As an additional control we include state dummies in all regressions.

We use the BEA's measure of personal income, which along with county population gives per capita income. We adjust it to be net of government transfers and

<sup>10</sup> Metro counties are those that contain cities with populations of 100,000 or more, or border such counties.

<sup>&</sup>lt;sup>9</sup> The OLS and the CR-OLS point estimates are the same; only the standard errors differ. The actual significance of the CR correction appears to vary across the regions. According to the figures in Tables 2 and 3, the CR standard errors are sometimes higher, in comparison to the OLS standard errors. But often, they are not different than, or are even smaller than the OLS standard errors, which is consistent with Conley's (1999) conclusion that spatial correlation does not necessarily increase the standard errors. We shall note that the CR correction was not implemented within the 3SLS framework because the statistical properties of the resulting estimators are not known.

express it in 1992 dollars. Natural logs of real per capita income are used throughout. In addition to initial, we utilize 36 demographic conditioning variables, listed in Table 1.<sup>11</sup>

#### IV. Analysis of Convergence Rates

Table 2 reports the asymptotic conditional convergence rate estimates along with their 95 percent confidence intervals for all three estimation methods (OLS, CR-OLS, and 3SLS) across all different sub-samples considered. Following Evans (1997b, p. 16), we use the expression,  $c = 1 - (1 + T\beta)^{1/T}$ , to infer the asymptotic rate of convergence from the estimates of  $\beta$ .<sup>12</sup> The confidence intervals are obtained by computing the endpoints,  $\beta \pm 1.96^*$ (s.e.( $\beta$ )), and plugging them into  $c = 1 - (1 + T\beta)^{1/T}$ .

According to Table 2, the estimated conditional convergence rate using 3SLS is 6.58 percent, significant at the 1-percent level. Compare this to 2.40 percent using OLS, also significant at 1-percent level. The difference between the two estimates is over 250 percent, suggesting that OLS introduces a substantial bias. Therefore, we primarily focus below on the 3SLS estimation results.

<sup>&</sup>lt;sup>11</sup> An appendix at the end of the paper describes the data in more detail.

<sup>&</sup>lt;sup>12</sup> The estimates of  $\beta$ , the coefficient on the log of 1970 real per capita income, are not reported here to save space. Also, the estimates of  $\beta$  when conditioning variables are excluded (which represents the hypothesis of absolute convergence) are much smaller in absolute value (e.g. –0.0068 (OLS) compared to –0.0173 (OLS) and –0.0344 (3SLS) for the entire sample) than their conditional counterparts, suggesting that the balanced growth paths vary across counties and, therefore, the determinants of the balanced growth paths need be conditioned upon. The unconditional  $\beta$  estimates, while smaller, are still negative and significant at the 1-percent level for the full sample, and for either the metro or the non-metro samples. Thus, we cannot reject absolute convergence, but only conclude that it is very slow. E.g., the point estimate of –0.0068 implies an absolute convergence rate of about 0.7 percent. Thus, counties close half the present gap between themselves and the wealthiest county in just under a century.

Metro- and non-metro counties yield similar results. For the metro-counties, 3SLS yields a convergence rate of 8.34 percent compared to 1.67 percent with OLS. The analogous numbers are 7.16 percent and 2.73 percent for the non-metro counties (all significant at 1-percent level). The overlap between the (narrow) 3SLS confidence intervals suggests that any difference between the metro and non-metro counties is small. Thus, a consistent estimate of the rate of convergence across the U.S. counties is in the range of 6 to 8 percent.<sup>13</sup>

Considering the variation in convergence rates by region, we find that it is lowest in Northeast (4.88 percent) and Plains (4.96 percent), followed by 5.01 percent in Great Lakes. The highest rates are found in the West (7.24 percent) and South (11.49 percent).<sup>14</sup>

Comparing the results for regional samples broken down into metro and nonmetro samples, we find the biggest difference in the West where the convergence rate point estimate of the metro counties is 13.93 percent, in contrast to 8.46 percent for the non-metro counties. We find a substantial difference in the South also, but in the opposite direction: 11.80 percent at the non-metro versus 7.57 percent at the metro counties. In the

<sup>&</sup>lt;sup>13</sup> Panel data estimation methods that difference the variables to remove fixed effects tend to report higher convergence rate estimates. E.g., Islam (1995), using international data, reports estimates of 4–5 percent (even higher for OECD countries). Barro (1997) shows that, because many conditioning variables remain stable, differencing them tends to emphasize measurement error over the correct information contained in the level, biasing convergence rate estimates upward. This argument, however, does not apply here for two reasons. First, Evans' (1997b) method produces consistent estimates while OLS without differencing does not. Second, the IV regressions of the 3SLS do not include conditionals—differenced or otherwise—so we do not emphasize measurement error in the regressions.

<sup>&</sup>lt;sup>14</sup> We find no correlation between regional convergence rates and the average rate of economic growth, suggesting that the "conditional" in conditional convergence is important, e.g., Northeast's balanced growth path may be high enough that it continued to grow faster than poorer regions with higher convergence rates.

remaining regions, the difference is less remarkable. In the Great Lakes the convergence rate of metro and non-metro counties are 6.61 percent and 5.25 percent respectively; in the Northeast, 5.06 percent and 4.91 percent respectively; and in the Plains, 5.18 percent and 5.71 percent respectively.

#### V. Analysis of Balanced Growth Path Determinants

Since we do not reject the conditional convergence hypothesis, the effects of conditioning variables are interpreted as *influences on the height of an individual economy's balanced growth path*.<sup>15</sup> By this interpretation, coefficients indicate the correlation of variables with income growth indirectly via the position of the balanced growth path. Given that position, the average growth rate increases (if the balanced growth path is higher) or decreases (if it is lower) as a result of the deviation of the economy from its individual balanced growth path and the convergence effect.

We now focus on these indirect effects of the conditioning variables on balanced growth paths. The variables we discuss are grouped into educational variables, government employment variables and industry variables (see Table 1). More detailed results are included in an appendix, which is available upon request.<sup>16</sup>

<sup>&</sup>lt;sup>15</sup> If convergence were to be rejected then the coefficients would be better interpreted as influences on individual economies' balanced growth *rates* (Evans 1997b).

<sup>&</sup>lt;sup>16</sup> Levine and Renelt (1992) show that cross-country regressions may not be robust to small changes in the conditioning variable set. In particular, "broad array of fiscal-expenditure variables..., are not robustly correlated with growth" (p. 943). The 3SLS method theoretically yields consistent estimators regardless of the variables included. Further, after running 3SLS regressions for the full sample with all conditioning variables, we ran the regressions without the conditioning variables that initially had coefficient estimates of less than 0.0000 in absolute value and found that the remaining coefficients remained stable. Thus, the 3SLS method seems to help us avoid Levine and Renelt's criticisms in theory and in practice.

### A. Educational Attainment

Our data include eight variables measuring educational attainment. Here we focus on 3SLS results for three of them: the percent of the population with (i) high school diploma, (ii) some college education, and (iii) bachelor degree or more (see Table 3). To save space, only the full 3,058 county results appear in Table 3.<sup>17</sup>

The coefficient for the population achieving (but not surpassing) a high school diploma is about 0.0091 percent, significant at the 1-percent level.<sup>18</sup> The results for the percent of the population with some college, but not a bachelor's degree, are more surprising. The coefficient is –0.0014, but it is not statistically significant. The sign of the coefficient is positive for metro (0.0009) and non-metro (0.0032) counties. It is in neither case significant, however. Compare this to the (perhaps-less-surprising) coefficient for the percent of the population with at least a bachelor degree, 0.0701, significant at the 1-percent level. A possible interpretation of these findings concerns the opportunity cost of education. College education ostensibly involves a benefit, in the form of increased skills, but it also involves a cost in the form of wages foregone. The results might lead one to believe that a college education of four years represents a positive net return, while the net return on a two-year degree is questionable.<sup>19</sup>

<sup>&</sup>lt;sup>17</sup> The results from the other samples are available on request.

<sup>&</sup>lt;sup>18</sup> The coefficient on the percent of the population with 11 years of education or less is –0.0204, significant at the 1-percent level. This is not surprising. The greater percent of an economy's population without minimal skills—not to mention discipline and socialization—necessary for a high school diploma, the lower the balanced growth path.

<sup>&</sup>lt;sup>19</sup> Kane and Rouse (1995) and Surette (1997) find that the return to 2-year degree is positive, at about 4–6 and 7–10 percent, respectively. Neither of these studies, however, uses county data. In addition, they do not take into account the *social* return, which our estimates presumably do. They look at individuals' costs and

One reason for a "bachelor degree or more" positive estimated effect is that "college towns," i.e. counties where a university or college is an exceedingly large percent of the population, bias the results. College towns have a disproportional number of advanced degree-holding individuals, and may, therefore, have higher incomes. However, we attempt to control for this by including a "college town" dummy variable. We take any college or university that had total enrollment (at a single campus) of 10,000 or more and calculate the ratio of enrollment to its county's 1970 population. The county's dummy is assigned a value of 1 if this ratio was at least 0.10 and a value of 0 otherwise.<sup>20</sup>

Comparing the metro and non-metro counties, we find that the coefficient on the "bachelor degree or more" variable for the metro and non-metro counties is 0.1151 and 0.0554, respectively, both significant at the 1-percent level. Thus, it appears that bachelor degree or more-level of attainment in the metro area has a considerably larger effect on a balanced growth path than the same attainment in the non-metro area.

### B. Size of the Public Sector

Our data include variables capturing the size of the public sector at three levels of

benefits while we consider their effect on the balanced growth path. What we might be seeing in our results, therefore, is a questionable social return to an associate degree. This is potentially an important finding for policy-makers. As Kane and Rouse (1995, p. 600n) note, "20 percent of Federal Pell Grants, 10 percent of Guaranteed Student Loans, and over 20 percent of state expenditures for post-secondary education, go to community colleges." If the social return to college education that does not end with a bachelor degree is not positive, then the subsidies must be reconsidered or restructured as to encourage a bachelor degree or more as the final outcome. Alternatively, the some college coefficient may primarily represent the effect of college dropouts who ultimately obtain no degree at all.

<sup>20</sup> We checked for robustness to a cutoff value of 0.05 also but found no noticeable change in the results.

government. These are the percents of a county's population employed by (i) the federal government, (ii) the state government, and (iii) the local government.

The issue of whether or not government fosters or hinders economic growth has been explored widely. See, for example, Aschauer (1989), Barro (1991), Easterly and Rebelo (1993), Evans and Karras (1994), and Folster and Henrekson (2001). These studies, however, use government expenditure variables to capture the size and the scope of government activities. We, in contrast, use the percent of a county's population employed by the federal, state, and local governments.

These variables offer several advantages. First, they allow us to explore how the relationship between government and growth differs at the three levels of decentralization. For example, a reasonable belief may be that local governments can more closely ascertain and respond to the needs of their constituents. The productivity of government may be expected to decrease as it gets more centralized. We can address such a hypothesis whereas previous studies could not.

Second, the use of three measures of government activity helps us avoid the problems of interpreting coefficients across geographical units when externalities are present, e.g., a state government may operate educational institutes (at a cost detectable in a growth regression) only to have many of the students, upon graduation, leave to live and work in other states (creating benefits not detectable in growth regressions). In general one would expect externalities to be less important for state than federal government, and even less important for local than state government. As another example, a negative coefficient on the federal government measure might be questioned because the federal services are spread across the nation, while a negative coefficient on a local government measure is immune to such a suspicion.

Third, the variables measuring the percent of population employed allow for a

fundamentally different and complementary way of conceptualizing the extent of government's involvement in the economy. The percent of a population employed by government can be interpreted as a stock of government activities producing a flow of services, while government expenditures are the flow of services. Moreover, the percent of a population employed gives a direct perspective on to what extent government is *involved*, i.e. how much of labor force is directed by government, rather than simply how much government spends.<sup>21</sup>

Table 3 summarizes the estimation results for the full sample. We find a negative and statistically significant partial correlation between the percent of the population employed in the public sector and the rate of growth, regardless of whether one considers federal, state or local government. Moreover, there is no clear pattern of a less negative partial correlation at increasingly decentralized levels. The coefficients for the federal, state and local employee percent of the population variables are –0.0226, –0.0177, and – 0.0198 respectively, all significant at the 1-percent level.

However, the relationship might be nonlinear, e.g., government to a certain extent might be good, but then becomes a negative influence as it expands further. To check this, we run the 3SLS regressions for the full U.S. sample with both linear and quadratic terms,  $\gamma_{fl}F + \gamma_{sl}S + \gamma_{ll}L + \gamma_{fs}(F)^2 + \gamma_{ss}(S)^2 + \gamma_{ls}(L)^2$ , where *F*, *S*, and *L* are the percent of population employed by federal, state, and local governments, respectively.

With the quadratic terms, the marginal effect of, e.g., the federal government variable on the average growth rate is given by  $\partial g / \partial F = \gamma_{fl} + 2\gamma_{fs}(F)$ . Thus, a positive coefficient on the linear term and a negative on the quadratic term imply that the marginal effect of *F* on *g* is positive until a level of *F* where the second term exceeds the first.

<sup>&</sup>lt;sup>21</sup> Of course, these are not mutually exclusive. E.g., government *spends* on wages so that part of the labor force is *involved* in government actions. This overlap makes the two types of variables complementary.

The estimation results with the quadratic terms included do not conform to the above. For federal, state and local government variables entered linearly, the estimates are negative and significant, as in the original regressions. For the quadratic variables, only the federal government coefficient is significant and positive. Using the estimated figures, significant at the 1 percent and 5 percent levels, respectively, we obtain  $\partial g / \partial F = -0.0331 + 2(0.0477)(F)$ , which, after setting equal to zero, implies that marginal additions to *F* are negatively correlated with *g* for *F* values up until 0.35 (until the government employs over 35 percent of the population), and then marginal additions are positively correlated with *g*. The overall partial correlation between *F* and *g* would not be positive until *F* exceeded 0.60 (60 percent of the population). Such *F* values, however, are unreasonable for the U.S. and make little sense.<sup>22</sup> For realistic values, federal government appears negatively correlated with growth.

# C. Industry Composition Effects

We have 16 industry-level variables, measuring the percent of the population employed in a given industry.<sup>23</sup> Interpreting correlations between these variables and income growth is difficult and we stress that interpretations below are of a speculative nature. We focus only on three industries that appear to have significant effects, and about which we feel our speculations are plausible. (See Table 3.)

*i.* Finance, Insurance and Real Estate Services

<sup>&</sup>lt;sup>22</sup> Only 9 out of 3,058 counties even have F values of at least 0.30. Also note that military incomes are excluded from our personal income data.

<sup>&</sup>lt;sup>23</sup> These industries are agriculture, communications, construction, finance, insurance and real estate, manufacturing of durables, manufacturing of non-durables, mining, retail, business and repair services, educational services, professional and related services, health services, personal services, entertainment and recreational services, transportation services, and wholesale trade.

We find a positive correlation between the percent of the population employed in finance, insurance and real estate services and economic growth across U.S. counties. The coefficient estimate for the entire sample is 0.0731, significant at the 1-percent level. The correlation is similar whether one considers the metro (0.0600) or non-metro (0.0699) sub-sample (not displayed in Table 3), both significant at the 1-percent level. A possible reason for this finding is the link between financial intermediation and economic growth, as reported by Rousseau and Wachtel (1998) who document quantitatively important links between financial intensity and per capita output level in five OECD countries.<sup>24</sup>

#### *ii.* Educational Services

Unlike educational attainment, the percent of population providing educational services has a negative effect on growth, -0.0445, significant at 1-percent level. The coefficient is negative also at both metro counties with the estimate of -0.0577, significant at the 1-percent level, and at the non-metro counties, with the estimate of -0.0335, significant at the 5-percent level (not displayed in Table 3).

One explanation for this correlation is that the benefits of education are not entirely internalized by the providing county.<sup>25</sup> E.g., many college graduates do not remain within the county where their colleges are located. The finding discussed above that educational attainment is positively correlated with growth—is silent as to where a county's population accumulated that stock. Tamura (1991, p. 523) argues that labor

<sup>&</sup>lt;sup>24</sup> Our findings may be interpreted also as offering empirical support to the models of Greenwood and Jovanovic (1990) and King and Levine (1993), in which financial development promotes economic growth. A broad survey of both theoretical and empirical analysis of the link between finance and growth is provided by Levine (2004).

<sup>&</sup>lt;sup>25</sup> Another explanation for this finding is a possible bureaucratic over-expansion of the public school systems as suggested by Marlow (2001), and frequently mentioned in media discussions.

moves "... to areas where the external effect is operative." Individuals may attend a college in counties where human capital is easier to acquire, and then move to other counties. This would be particularly true for metro counties where the majority of colleges and universities are located. Indeed, we find that the negative relationship between the percent of population employed in educational services and economic growth is stronger for the metro counties, in the aggregated data as well as for each region of the US.<sup>26</sup>

#### iii. Entertainment and Recreational Services

The effect of this variable on economic growth is positive, 0.0335, significant at the 5-percent level, and is larger in metro counties, with estimates of 0.0670 (significant at the 10-percent level) and 0.0229 (not significant), for metro and non-metro counties, respectively. This is potentially important. To put it in perspective, it is larger (in absolute value) than the effect of the public sector size variables. Also, Costa (1997) reports that, as a percent of households' budgets, recreation expenses rose from 1.9 percent in 1890 to

<sup>&</sup>lt;sup>26</sup> Again, the metro/non-metro and regional results are available upon request. One may argue that, even in the absence of externalities, we would expect the partial effect of educational services provision to be negative because educational attainment is already controlled for. In other words, employment in provision is the cost and attainment is the benefit. There are two possible responses to this argument. First, attainment variables are initial stocks and so the educational provision variable is also an initial stock. The *flows* from those stocks, that we would expect to contribute to growth over time, are services from human capital and new human capital creation respectively. Of course, these two flows are likely correlated. Second, we may not expect a negative coefficient on educational provision because it may proxy, as a measure of input intensity, for educational quality that is not captured in simple attainment stocks. However, it is well known that variation in conventional measures of resources devoted to education (e.g. per student spending, teacher to student ratios and teacher experience/education) generally does a poor job of accounting for variation in student achievement, e.g. Hanushek (1996). Our negative coefficient estimate is not inconsistent with this regularity.

4.5 percent in 1950 and then to 5.6 percent in 1991. Entertainment and recreation services, thus, comprise an increasingly large segment of the U.S. economy. The above finding might be capturing the increase in economic activity that is fostered by the presence of gambling casinos and professional sports teams and their stadiums.<sup>27</sup>

#### VI. Conclusion and Caveats

We use county-level data from 3,058 U.S. counties to study economic growth and measure the speed of convergence. County-level data are valuable for studying convergence because they form a sample with substantial homogeneity and mobility of resources and technology without sacrificing the benefits of a large number of crosssectional units. We use 41 different conditioning variables to capture cross-county heterogeneity and to assess how the variables affect the balanced growth paths. We report OLS and 3SLS-IV estimates for the entire data set as well as for its subsets, which include metro and non-metro counties, and counties grouped into five regions.

We find that while OLS yields estimates of the asymptotic convergence rate just above 2 percent, the 3SLS method consistently estimates a convergence rate between 6 and 8 percent. This difference is economically significant: it represents a difference in the half-life of the gap between present levels of income and the balanced growth path of 32– 33 years versus 12–13 years, respectively. We also find that the convergence rates are quite variable: the Southern counties converge more than two and half times faster than the counties in the Northeast. In addition, we find that the size of the public sector at *all* 

<sup>&</sup>lt;sup>27</sup> Siegfried and Zimbalist (2000) report that by 2005 there will be 95 professional sports stadiums having been constructed since 1990, with more than \$27.1 billion spent on them. Eadington (1999) notes that gross gaming revenues had reached \$540 billion in 1997. Anderson, Arthur and Co. (1997) and Walker and Jackson (1998) also show that introduction of casino industries can stimulate economic growth.

levels (federal, state and local) is negatively correlated with economic growth. Further, the relationship between educational attainment and economic growth is nonlinear: it is positive for up to high school, insignificant or even negative for between high school and associates degree levels, and then positive for further years of schooling. Finally, a large presence of finance, insurance and real estate industry, and entertainment industry is positively correlated with economic growth while the percent of a county's population employed in the education industry is negatively correlated with economic growth.

We should stress that the coefficients estimated on our conditioning variables are, strictly speaking, only partial correlations between those variables and a county growth rates. Given the validity of the neoclassical model as a useful approximation to reality, they can be interpreted as the effects of the conditioning variables on balanced growth path positions. However, they are at least interesting as a summarization of associations between U.S. county growth rates and a broad set of county demographic measures

An interesting issue that was brought up by one of the referees is the question of the applicability of the neoclassical growth model framework to such "open" economies as the US counties. We agree that the neo-classical growth model may not be the most suitable framework for thinking about growth in a cross-section of US counties given their extraordinary degree of "openness."

A way around this problem has been proposed recently by Rappaport (1999, 2005) who offers a version of the neoclassical growth model for studying "local growth," where by local is meant small open economic units comprising a larger entity, such as counties comprising the U.S. The distinguishing characteristic of small open economies such as U.S. counties is the extraordinary mobility of labor. The question, then, is how does labor mobility affect convergence? Rappaport (1999, 2005) expands the standard neoclassical growth model to allow for labor mobility and demonstrates that the model

21

predicts conditional convergence. Indeed, that is what we find here.<sup>28</sup>

We shall note also that all convergence rate estimates above are based, ultimately, on a specification from the neoclassical growth model. That model is of a closed economy and convergence is a phenomenon based entirely on diminishing returns to accumulated capital—specifically capital accumulated *from the economy's own savings.*<sup>29</sup> However, across U.S. counties, especially within a given state, there is considerable capital mobility. Perfect capital mobility would predict immediate equalization of returns and instantaneous convergence, but<sup>30</sup> convergence rates less than 100 percent may still obtain for open economies in the presence of adjustment costs as well as imperfect capital markets (Levy, 2000 and 2004). Barro et al. (1995) demonstrate that gradual convergence will occur if "capital is only partially mobile" because "borrowing is possible to finance accumulation of physical capital, but not accumulation of human capital" (p. 104). If human capital accounts for a significant share of income (e.g., Barro and Sala-i-Martin, 1992, and Mankiw, et al., 1992, suggest about 1/2) then this would account for gradual convergence.<sup>31</sup> As well, Barro and Sala-i-Martin (1997) present a model where

<sup>28</sup> Further, Rappaport (1999, 2005) finds that that convergence can either be accelerated (by a positive effect of out-migration on wages) or slowed (by a resultant disincentive for capital accumulation) depending on relative changes in marginal products. Rappaport's (2005) analysis suggests that at relatively low levels of income the later effect dominates.

<sup>29</sup> This can include both physical and, in the case of the so-called augmented Solow model (e.g. Mankiw et al, 1992), human capital.

<sup>30</sup> Since we are looking at *per capita* income convergence we need not address issues relating to labor mobility in the same way as capital mobility.

<sup>31</sup> At the U.S. state level, there is evidence that even financing of physical capital from one state to another is not perfect. Driscoll (2004) found that state-specific variation in deposits has a large and statistically significant effect on state-specific loans. The intuition is that some firms that do not regard bank loans and forms of direct finance as perfect substitutes and out-of-state bank lending is not prevalent. U.S. Federal technological diffusion occurs across economies through imitation of the leader's technologies (which is cheaper than innovation). With increasing costs to imitation (e.g. easier ideas to copy are copied first) gradual convergence will occur. Any of the above assumptions can imply (gradual) conditional convergence and place reality somewhere between the convergence rate of a closed economy and the instant convergence of an open economy with perfect capital markets.

Future research could explore interactions of initial income and schooling variables. For example, schooling may affect the ability of an economy to converge. Similar hypotheses could be made concerning government variables. Another avenue for future research could also consider the possibility of a structural relationship between government expenditures and growth as suggested by Slemrod (1995).

regulation, until recently, restricted out-of-state bank lending and, even as late as 1994, more than 70 percent of bank assets were in the control of within-state entities (Berger, et al., 1995).

# 23 Data Appendix

# A. Construction of Metro and Non-Metro County Level Data

A population size of 100,000 was chosen as the minimum threshold for metro counties for three reasons.<sup>32</sup> First, the data available was limited with respect to reporting smaller city sizes. Second, the BEA uses the 100,000-figure as the minimum necessary for classifying a locality as a county for the purpose of processing the county (or county-equivalent) source data.<sup>33</sup> Third, it was felt that cities with smaller populations would not provide the spillover effects into the surrounding counties needed to justify the decision rule. Note that these populations are of the actual cities and they do not include the populations in the surrounding metropolitan areas. For example, the population for the city of Atlanta is only the population within city limits and not Fulton County – the county where Atlanta resides. Additionally, this decision rule extends beyond state boundaries. For example, Cincinnati is located in southwestern Ohio. The Cincinnati metro area, however, extends well beyond southwestern Ohio into northern Kentucky and southeastern Indiana. Therefore, when the metro counties are viewed on aggregate it is without regard to state boundaries.

This decision rule also errs on the side of conservatism. It may be the case that metropolitan areas with very large populations expand out beyond what our classification would indicate. However, the majority of the overall population for those metropolitan areas has been captured. Additionally, by erring on the side of conservatism we can be more confident that the metro counties are more homogenous than they might otherwise

<sup>&</sup>lt;sup>32</sup> In order to determine which cities had populations over 100,000, we used Census Bureau publication SU-99-1, "Population Estimates for Cities with Populations of 100,000 and Greater."

<sup>&</sup>lt;sup>33</sup> See "Local Area Personal Income, 1969–1992," Bureau of Economic Analysis, Regional Accounts Data, February 2, 2001, p. 1.

be. For example, since we are unable to further sub-divide counties, the farthest reaches of a metropolitan area may contain a county where only a small portion of the population would be classified as belonging to that metropolitan area. If we were to include that entire county as a metro county we would be incorrectly classifying the entire county.

Our decision to err on the side of conservatism might impact our final results in the following way. The metro county analysis results will be slightly understated since it may be excluding small populations on the outskirts of metropolitan areas and our nonmetro county analysis results may be slightly over stated for the exact opposite reason—it will be including a population that should otherwise be categorized as metro.

That is why we chose not to use the metropolitan statistical areas (MSA), as defined by the Office of Management and Budget (OMB). An example will help demonstrate the difference. The MSA for Atlanta, GA, as defined by the OMB consists of the following 20 counties: Barrow, Bartow, Carroll, Cherokee, Clayton, Cobb, Coweta, DeKalb, Douglas, Fayette, Forsyth, Fulton, Gwinnett, Henry, Newton, Paulding, Pickens, Rockdale, Spalding, and Walton.<sup>34</sup> Our metro classification for Atlanta consists of the following 10 counties: Carroll, Cherokee, Clayton, Cobb, Coweta, DeKalb, Fayette, Forsyth, Fulton and Gwinnett. The 10 counties included in our metro region contain the largest portion of the metropolitan area, in terms of population. It should be noted that our metro classification contains most of the same MSAs as the OMB's classification. The counties that constitute those regions, however, are different, as demonstrated above. As previously noted, our classification tends to have fewer counties attached to a particular metropolitan area providing, we believe, a more homogenous population.

#### B. Construction of the Regional County Level Data

<sup>&</sup>lt;sup>34</sup> The city of Atlanta is located in Fulton County.

To perform the second set of sub-sample analysis, we have separated the sample into five regional subgroups: Northeast, Great Lakes, Western, Plains, and Southeastern states. The limiting constraint on further increasing the regions was the number of counties within some of the States. For example, a few of the states in the Northeast Region have less than ten counties. Given the number of independent variables, it was necessary to increase the size of the regions in order to increase the overall number of observations. An attempt was made to group states that were closely related to each other as much as possible in terms of their economic and socio-economic characteristics.<sup>35</sup>

Given the data constraints, it was necessary to use an interpolation procedure for some variables.<sup>36</sup> In this study we cover the 1970–1998 period. However, in order to implement the Evans' (1997a, 1997b) 3SLS estimation method as described in section 2, we needed to have available data values for 1969 and 1997. We used a linear interpolation method to generate these missing observations. It should be noted that none

<sup>&</sup>lt;sup>35</sup> The regional breakdown of the individual states is as follows (the figures in parentheses indicate the number of counties in the given state). **NORTHEAST**: Maine (16), New Hampshire (10), Vermont (14), Massachusetts (14), Connecticut (8), Rhode Island (5), Delaware (3), Washington, DC (1), Maryland (24), New Jersey (20), New York (62), and Pennsylvania (67); **GREAT LAKES**: Illinois (102), Indiana (92), Michigan (83), Ohio (88), and Wisconsin (70); **WEST**: Alaska (9), California (58), Hawaii (4), Nevada (17), Oregon (36), Washington (39), Arizona (9), New Mexico (32), Oklahoma (77), and Texas (254); **PLAINS**: Iowa (99), Kansas (106), Minnesota (87), Missouri (113), Nebraska (93), North Dakota (53), South Dakota (66), Colorado (63), Idaho (44), Montana (56), Utah (29), and Wyoming (23); and **SOUTH**: North Carolina (100), South Carolina (45), Georgia (159), Florida (67), Tennessee (95), Alabama (67), Mississippi (82), Louisiana (64), Arkansas (74), Kentucky (120), Virginia (84), and West Virginia (55). This yields a total of 3,058 counties. (The original sample contained 3,066 counties but 8 counties were excluded for lack of data.)

<sup>&</sup>lt;sup>36</sup> Given the cross-section nature of our data, the use of interpolation does not cause problems of the type reported by Dezhbakhsh and Levy (1994), who focus on the periodic properties of interpolated time series.

of the data relating to income and population variables were generated by this method, as they were available from BEA-REIS on a yearly basis for the entire period covered. The Census data variables, which were available in 1970, 1980 and 1990, were interpolated in order to generate the 1969, 1997, and 1998 values.

# C. Measurement of Per Capita Income

Because of the critical importance of the income variable for the study of growth and convergence, we want to address its measurement in some detail. Two options were available to us for the construction of the county-level per capita income variable: (1) Census Bureau database, and (2) BEA-REIS database.

Income information collected by the Census Bureau for states and counties is prepared decennially from the "long-form" sample conducted as part of the overall population census (BEA, 1994). This money income information is based on the selfreported values by Census Survey respondents. An advantage of the Census Bureau's data is that they are reported and recorded by place of residence. These data, however, are available only for the "benchmark" years, i.e., the years in which the decennial Census survey is conducted.

The second source for this data, and the one chosen for this project, is personal income as measured by the Bureau of Economic Analysis (BEA).<sup>37</sup> The definitions that are used for the components of personal income for the county estimates are essentially the same as those used for the national estimates. For example, the BEA defines "personal income" as the sum of wage and salary disbursements, other labor income, proprietors' income (with inventory valuation and capital consumption adjustments),

<sup>&</sup>lt;sup>37</sup> The data and their measurement methods are described in detail in "Local Area Personal Income, 1969– 1992" published by the BEA under the Regional Accounts Data, February 2, 2001.

rental income (with capital consumption adjustment), personal dividend income and personal interest income. (BEA, 1994) "Wage and salary disbursements' are measurements of pre-tax income paid to employees. "Other labor income" consists of payments by employers to employee benefit plans. "Proprietors' income" is divided into two separate components—farm and non-farm. Per capita income is defined as the ratio of this personal income measure to the population of an area.<sup>38</sup>

The BEA compiles data from several different sources in order to derive this personal income measure. Some of the data used to prepare the components of personal income are reported and recorded by place of work rather than place of residence. Therefore, the initial estimates of these components are on a place-of-work basis. Consequently, these initial place-of-work estimates are adjusted so that they will be on a place-of-residence basis and so that the income of the recipients whose place of residence differs from their place of work will be correctly assigned to their county of residence.

As a result, a place of residence adjustment is made to the data. This adjustment is made for inter-county commuters and border workers utilizing journey-to-work (JTW) data collected by Census. For the county estimates, the income of individuals who commute between counties is important in every multi-county metropolitan area and in many non-metropolitan areas. The residence adjustment estimate for a county is calculated as the total inflows of the income subject to adjustment to county *i* from

<sup>&</sup>lt;sup>38</sup> The BEA's estimates of personal income reflect the revised national estimates of personal income that resulted from the 1991 comprehensive revision and the 1992 annual revisions of the national income and product accounts. The revised national estimates were incorporated into the local area estimates of personal income as part of a comprehensive revision in May 1993. In addition, the estimates incorporate source data that were note available in time to be used in the comprehensive revisions. For details of these revisions, see "Local Area Personal Income: Estimates for 1990–92 and Revisions to the Estimates for 1981–91," *Survey of Current Business* 74 (April 1994), 127–129.

county *j* minus the total outflows of the income subject to adjustment from county *i* to county *j*. The estimates of the inflow and outflow data are prepared at the Standard Industrial Classification (SIC) level and are calculated from the JTW data on the number of wage and salary workers and on their average wages by county of work for each county of residence from the Population Census.

Obviously, metro areas and the surrounding counties will have a higher proportion of "cross-county" commuters. By using our classification system for metro counties we alleviate any problems that might arise with the BEA's adjustment process as we are grouping the metro counties into one single observation unit. Moreover, the classification we have in place should pick up the majority of cross-county commuters.

- Anderson, Arthur & Co., *Economic Impacts of Casino Gaming in the United States*, *Volume 2: Micro Study* (Washington, D.C.: American Gaming Assoc., 1997).
- Aschauer, D. A., "Is Public Expenditure Productive?" *Journal of Monetary Economics* 23 (1989), 177-200.
- Barro, R. J., "Economic Growth in a Cross Section of Countries," *Quarterly Journal of Economics* 106 (1991), 407–443.
- Barro, R. J., *Determinants of Economic Growth: A Cross-Country Empirical Study* (Cambridge, MA: MIT Press, 1997).
- Barro, R. J., N. G. Mankiw, and X. Sala-Martin, "Capital Mobility in Neoclassical Models of Growth," *American Economic Review* 85 (1995), 103–115.
- Barro, R. J. and X. Sala-I-Martin, "Convergence across States and Regions," *Brookings Papers on Economic Activity* 1 (1991), 107–158.
- Barro, R. J. and X. Sala-i-Martin, "Convergence," *Journal of Political Economy* 100 (1992), 223–251.
- Barro, R. J. and X. Sala-i-Martin, "Technological Diffusion, Convergence, and Growth," *Journal of Economic Growth* 2 (1997), 1–27.
- Berger, A. N., A. Kashyap, and T. H. Hannan, "The Transformation of the U.S. Banking Industry: What a Long, Strange Trip It's Been," *Brookings Papers on Economic Activity* 2 (1995), 55–218.
- Brock, W. A. and S. N. Durlauf, "Growth Empirics and Reality," University of Wisconsin, Madison, Working Paper 2024 (2001).
- Caselli, F., G. Esquivel, and F. Lefort, "Reopening the Convergence Debate: A New Look at Cross-Country Growth Empirics," *Journal of Economic Growth* 1 (1996),

Conley, T. J., "GMM Estimation with Cross Sectional Dependence," *Journal of Econometrics* 92 (1999), 1–45.

Costa, D., "Less of a Luxury: The Rise of Recreation since 1888," NBER WP 6054 (1997).

Dezhbakhsh, H. and D. Levy, "Periodic Properties of Interpolated Time Series," *Economics Letters* 44 (1994), 221–228.

- Driscol, J. C., "Does Bank Lending Affect Output? Evidence from the U.S. States," Journal of Monetary Economics 51 (2004), 451–471.
- Durlauf, S. N., "Manifesto for a Growth Econometrics," *Journal of Econometrics* 100 (2001), 65–69.
- Durlauf, S. N. and D. Quah, "The New Empirics of Economic Growth," in J. Taylor andM. Woodford, eds, *Handbook of Macroeconomics* (N. Holland: Elsevier, 1999).
- Eadington, W. R., "The Economics of Casino Gambling," *Journal of Economic Perspectives* 13 (1999), 173–192.
- Easterly, W. and Sergio R., "Fiscal Policy and Economic Growth," *Journal of Monetary Economics* 32 (1993), 417–458.
- Evans, P., "How Fast Do Economies Converge?" *Review of Economics and Statistics* 36 (1997), 219–225.
- Evans, P., "Consistent Estimation of Growth Regressions," manuscript, available at http://economics.sbs.ohiostate.edu/pevans/pevans.html (1997).
- Evans, P. and G. Karras, "Are Government Activities Productive? Evidence from a Panel of US States," *Review of Economics and Statistics* 76 (1994), 1–11.
- Folster, S. and M. Henrekson, "Growth Effects of Government Expenditure and Taxation in Rich Countries," *European Economic Review* 45 (2001), 1501–1520.

Greenwood, J. and B. Jovanovic, "Financial Development, Growth, and the Distribution

- Hanushek, E. A., "Measuring Investment in Education," Journal of Economic Perspectives 10 (1996), 9–30.
- Islam, N., "Growth Empirics: A Panel Data Approach," *Quarterly Journal of Economics* 110 (1995), 1127–1170.
- Kane, T. J. and C. E. Rouse, "Labor-Market Returns to Two- and Four-Year College," American Economic Review 85 (1995), 600-614.
- King, R. G. and R. Levine, "Finance and Growth: Schumpeter Might be Right," Quarterly Journal of Economics 108 (1993), 717–737.
- Levine, R. and D. Renelt, "A Sensitivity Analysis of Cross-Country Growth Regressions," American Economic Review 82 (1992), 942–963.
- Levine, R., "Finance and Growth: Theory and Evidence," NBER W.P. #10766 (2004).
- Levy, D., "Investment-Saving Comovement and Capital Mobility: Evidence from Century Long US Time Series," Review of Economic Dynamics. 3 (2000), 100–136.
- Levy, D., "Is the Feldstein-Horioka Puzzle Really a Puzzle?" in Aspects of Globalization: Macroeconomic and Capital Market Linkages in the Integrated World Economy, edited by C. Tsoukis, G. M. Agiomirgianankis, and T. Biswas (London: Kluwer Academic Publishers, 2004), 49–66.
- Mankiw, N. G., D. Romer, and D. N. Weil, "A Contribution to the Empirics of Economic Growth, "Quarterly Journal of Economics 107 (1992), 407-437.
- Marlow, M. L., "Bureaucracy and Student Performance in U.S. Public Schools," Applied *Economics* 33 (2001), 1341–1350.

Quah, D. T., "Empirics for Economic Growth and Convergence," European Economic

Review 40 (1996), 1353-1375.

Rappaport, J., "Local Growth Empirics," CID W.P. #23, Harvard University (1999).

- Rappaport, J., "How Does Labor Mobility Affect Income Convergence?" Journal of Economic Dynamics and Control 29 (2005), 567-581.
- Rappaport, J., and J. D. Sachs, "The United States as a Coastal Nation," *Journal of Economic Growth* 8 (2003), 5–46.
- Rousseau, P. L. and P. Watchel, "Financial Intermediation and Economic Performance: Historical Evidence from Five Industrialized Countries," *Journal of Money, Credit and Banking* 30 (1998), 657–678.
- Sala-i-Martin, X., "Regional Cohesion: Evidence and Theories of Regional Growth and Convergence," *European Economic Review* 40 (1996), 1325–1352.
- Siegfried, J. and A. Zimbalist, "The Economics of Sports Facilities and Their Communities," *Journal of Economic Perspectives* 14 (2000), 95–114.
- Slemrod, J., "What Do Cross-Country Studies Teach Us about Government Involvement, Prosperity, and Economic Growth," *Brookings Papers on Economic Activity* 2 (1995), 373–431.
- Surette, B. J., "The Effects of Two-Year College on the Labor Market and Schooling Experiences of Young Men," Board of Governors of the Federal Reserve System, Finance and Economics Discussion Series, No. 1997/44 (1997).
- Tamura, R., "Income Convergence in an Endogenous Growth Model," Journal of Political Economy 99 (1991), 522–540.
- Walker, D. M. and J. D. Jackson, "New Goods and Economic Growth: Evidence from Legalized Gambling," *Review of Regional Studies* 28 (1998), 47–69.

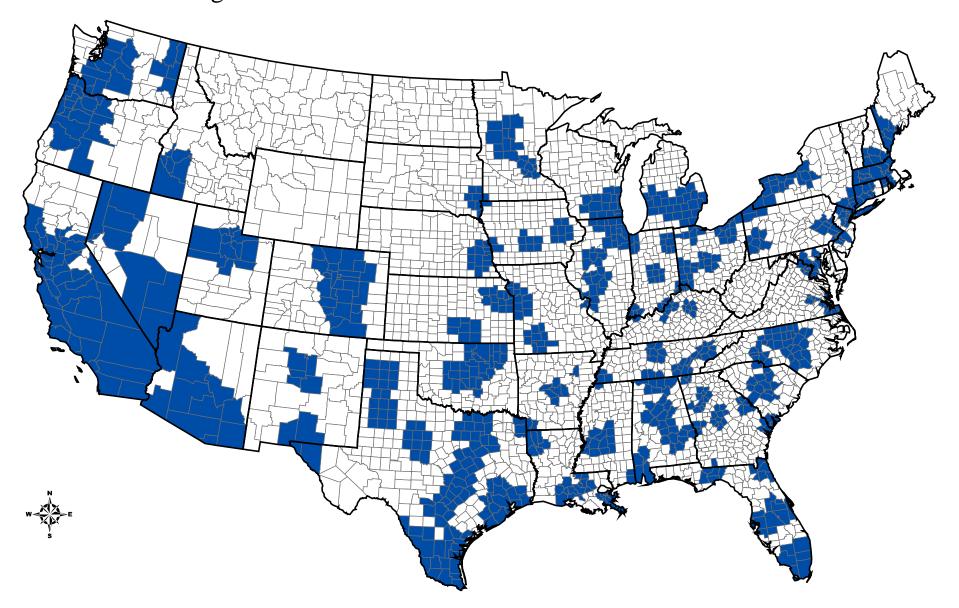


Figure 1: Metro & Non-Metro Counties - Continental U.S.

Note:

- (2) Hawaii has 1 metro county that contains Honolulu
- (3) Metro counties are shaded blue

<sup>(1)</sup> Alaska has 3 metro counties including and surrounding the city of Anchorage

TABLE 1.—VARIABLE DEFINITIONS AND THEIR SOURCE	TABLE 1V	<b>VARIABLE</b>	DEFINITIONS	AND	THEIR SOURCE
--	----------	-----------------	-------------	-----	--------------

Variable	Definition	Period	Source
Income	Per Capita Personal Income (excluding transfer	1969–1998	BEA
	payments)		
Land area per capita	Land area in km <sup>2</sup> /population	1970-1990	Census
Water area per capita	Water area in km <sup>2</sup> /population	1970-1990	Census
Age: 5-13 years	Percent of 5–13 year olds in the population	1970-1990	Census
Age: 14-17 years	Percent of 14–17 year olds in the population	1970-1990	Census
Age: 18-64 years	Percent of 18–64 year olds in the population	1970-1990	Census
Age: 65+	Percent of 65+ olds	1970-1990	Census
Blacks	Percent of Blacks	1970-1990	Census
Hispanic	Percent of Hispanics	1970-1990	Census
Education: 9-11 years	Percent of population with 11 years education or less	1970-1990	Census
Education: H.S. diploma	Percent of population with high school diploma	1970-1990	Census
Education: Some college	Percent of population with some college education	1970-1990	Census
Education: Bachelor +	Percent of population with bachelor degree or above	1970-1990	Census
Education: Public elementary	Number of students enrolled in public elementary	1970-1990	Census
	schools	1970 1990	Consus
Education: Public nursery	Number of students enrolled in public nurseries	1970-1990	Census
Education: Private elementary	Number of students enrolled in private elementary	1970-1990	Census
Education: Thrate clementary	schools	1970 1990	Consus
Education: Private nursery	Number of students enrolled in private nurseries	1970-1990	Census
Housing	Median house value	1970-1990	Census
Poverty	Percent of the population below the poverty line	1970-1990	Census
Federal government employment	Percent of population employed by the federal	1969-1998	BEA
rederal government employment	government in the county	1909-1998	DEA
State government employment	Percent of population employed by the state	1969-1998	BEA
	government in the county	10.00 1000	
Local government employment	Percent of population employed by the local government in the county	1969-1998	BEA
Self-employment	Percent of population self-employed	1970-1990	Census
Agriculture	Percent of population employed in agriculture	1970-1990	Census
Communications	Percent of population employed in agriculture Percent of population employed in communications	1970-1990	Census
Construction	Percent of population employed in construction	1970-1990	Census
Entertainment & Recreational	Percent of population employed in construction Percent of population employed in entertainment &	1970-1990	Census
Services	recreational services	1970-1990	Cellsus
Finance, insurance & real estate	Percent of population employed in finance, insurance,	1970-1990	Census
Manufacturing: durables	and real estate Percent of population employed in Manufacturing of	1970-1990	Census
Wanutacturing. durables	durables	1770-1770	Census
Manufacturing: non-durables	Percent of population employed in manufacturing of non-durables	1970-1990	Census
Mining		1070 1000	Census
Mining	Percent of population employed in mining	1970-1990	
Retail	Percent of population employed in retail trade	1970-1990	Census
Transportation	Percent of the population employed in transportation	1070 1000	C
Business & repair services	Percent of population employed in business and repair services	1970-1990	Census
Educational services	Percent of population employed in education services	1970-1990	Census
Professional related services	Percent of population employed in education services	1970-1990	Census
rolessional fetaled services	services	1970 1990	Census
Health services	Percent of population employed in health services	1970-1990	Census
Personal services	Percent of population employed in personal services	1970-1990	Census
Wholesale trade	Percent of population employed in wholesale trade	1970-1990	Census
College Town	Dummy Variable: 1 if the county had a college or	1970	National Cente
	university enrollment to population ratio greater than		for Educationa
	or equal to 5% and 0 otherwise.		Statistics
Metro area 1970	Dummy Variable: 1 if the county was in a metro area	1970	Census
mous area 1770	in 1970, and 0 otherwise	1770	Centous

All BEA variables are available annually from 1969 to 1998. All Census variables are gathered from the 1970, 1980 & 1990 Census tapes. Values for 1969 were obtained via the interpolation method as discussed in the data section.

Region	Area	Number of Counties	OLS Estimates and 95% C.I.	C-R OLS and 95% C.I.	3SLS Estimates and 95% C.I.
United States	All counties	3,058	0.0239 (0.0224, 0.0255)	0.0239 (0.0213, 0.0267)	0.0658 (0.0632, 0.0981)
Unites States	Metro counties	867	0.0167 (0.0142, 0.0194)	0.0167 (0.0137, 0.0201)	0.0834 (0.0749, 0.0944)
United States	Non-metro counties	2,191	0.0273 (0.0253, 0.0293)	0.0273 (0.0244, 0.0304)	0.0716 (0.0620, 0.0844)
Great Lakes	All counties	435	0.0219 (0.0179, 0.0264)	0.0219 (0.0189, 0.0252)	0.0501 (0.0446, 0.0564)
Great Lakes	Metro counties	140	0.0126 (0.0064, 0.0201)	0.0126 (0.0107, 0.0147)	0.0661 (0.0549, 0.0818)
Great Lakes	Non-metro counties	295	0.0282 (0.0223, 0.0353)	0.0282 (0.0241, 0.0330)	0.0525 (0.0447, 0.0626)
New England	All counties	244	-0.0031 (-0.0068, 0.0008)	-0.0031 (-0.0061, -0.0001)	0.0488 (0.0402, 0.0606)
New England	Metro counties	90	0.0018 (-0.0056, 0.0113)	0.0018 (-0.0007, 0.0047)	0.0506 (0.0364, 0.0756)
New England	Non-metro counties	154	0.0028 (-0.0024, 0.0091)	0.0028 (-0.0002, 0.0062)	0.0491 (0.0404, 0.0602)
Plains	All counties	832	0.0249 (0.0221, 0.0281)	0.0249 (0.0214, 0.0289)	0.0496 (0.0345, 0.0775)
Plains	Metro counties	143	0.0095 (0.0049, 0.0148)	0.0095 (0.0066, 0.0127)	0.0518 (0.0404, 0.0692)
Plains	Non-metro counties	689	0.0255 (0.0223, 0.0289)	0.0255 (0.0216, 0.0298)	0.0571 (0.0410, 0.0847)
Southern	All counties	1,009	0.0232 (0.0205, 0.0261)	0.0232 (0.0205, 0.0261)	0.1149 (0.0985, 0.1451)
Southern	Metro counties	252	0.0218 (0.0159, 0.0287)	0.0218 (0.0170, 0.0272)	0.0757 (0.0672, 0.0865)
Southern	Non-metro counties	757	0.0229 (0.0199, 0.0262)	0.0229 (0.0203, 0.0258)	0.1180 (0.0975, 0.1693)
Western	All counties	538	0.0318 (0.0276, 0.0365)	0.0318 (0.0273, 0.0369)	0.0724 (0.0590, 0.0953)
Western	Metro counties	242	0.0128 (0.0084, 0.0178)	0.0128 (0.0091, 0.0169)	0.1393 (0.1003, 0.1563)
Western	Non-metro counties	296	0.0356 (0.0296, 0.0428)	0.0356 (0.0308, 0.0 411)	0.0846 (0.0694, 0.1105)

Asymptotic convergence rates and 95% confidence intervals are calculated following Evans (1997a). The asymptotic convergence rate ( $\rho$ ) is determined by substituting the  $\beta$  from equation (3) into the equation  $\rho = 1 - [1 + (T \times \beta)]^{1/T}$ . The calculation of the 95% confidence interval follows two steps. First, we obtain new end points by computing  $\beta \pm$  its standard error. Second, these new values are substituted into the above equation. "C-R" abbreviates "Conley-Rappaport" and denotes standard errors obtained using Rappaport's (1999) implementation of Conley's (1999) correction for possible cross-county spatial correlation. See section 2 for details.

Variables	OLS	C-R OLS	3SLS
Educational Attainment			
High School Diploma	0.0007 (0.0028)	0.0007 (0.0052)	0.0091 (0.0029) <sup>a</sup>
Some College Education	-0.0107 (0.0056) <sup>c</sup>	-0.0107 (0.0089)	-0.0014 (0.0061)
Bachelor Degree or Higher	$0.0424 (0.0058)^{a}$	0.0424 (0.0108) <sup>a</sup>	0.0701 (0.0061) <sup>a</sup>
Government Employment			
Federal	-0.0145 (0.0048) <sup>a</sup>	-0.0145 (0.0046) <sup>a</sup>	$-0.0226 (0.0051)^{2}$
State	-0.0041 (0.0037)	-0.0041 (0.0045)	-0.0177 (0.0040)
Local	-0.0211 (0.0048) <sup>a</sup>	-0.0211 (0.0079) <sup>a</sup>	-0.0198 (0.0052) <sup>2</sup>
Industry Composition			
Finance, Insurance & Real Estate	$0.0632 (0.0117)^{a}$	$0.0632 (0.0233)^{a}$	0.0731 (0.0125) <sup>a</sup>
Education Services	$-0.0257 (0.0082)^{a}$	$-0.0257 (0.0060)^{a}$	-0.0445 (0.0087)
Entertainment & Recreational	0.0272 (0.0154) <sup>c</sup>	0.0272 (0.0230)	0.0335 (0.0166) <sup>b</sup>
Services			

TABLE 3.—ANALYSIS OF GROWTH: THE EFFECT OF SELECT VARIABLES, ENTIRE U.S.

Standard errors are reported in parentheses. "C-R" abbreviates "Conley-Rappaport" and denotes standard errors obtained using Rappaport's (1999) implementation of Conley's (1999) correction for possible cross-county spatial correlation. See section 2 for details. "a" denotes significance at the 1% level; "b" denotes significance at the 5% level; "c"denotes significance at the 10% level.

### **Referee's Appendix**

#### I. Inconsistency of OLS Estimates

The method of ordinary least squares (OLS) could be used to infer the values of  $\beta$  and  $\gamma$  in equation (2). However, Evans (1997b) states that the OLS estimates obtained from (2) are unlikely to be consistent.<sup>1</sup> In order to demonstrate this inconsistency, Evans first specifies a general autoregressive moving average (ARMA) data-generating process for  $y_{nt}$ :

(1A) 
$$y_{nt} - a_t = \delta_n + \lambda_n (y_{n,t-1} - a_{t-1}) + \sum_{i=1}^q \theta_{ni} \varepsilon_{n,t-i}$$

with

(2A) 
$$\delta_n = \kappa + \xi'_n x_n + \omega_n$$

where  $\varepsilon_{nt}$  is a zero-mean, covariance stationary error process independently distributed over time and across economies. The error term,  $\varepsilon_{nt}$ , is uncorrelated with  $x_n$ ,  $\lambda_n$  is an autoregressive parameter which lies on (0,1], and  $\theta_{n0}...\theta_{nq}$  satisfy the restriction  $\theta_{n0} = 1$ . As such,  $y_{nt} - a_t$  will also have an autoregressive representation and will be covariance stationary if  $\lambda_n < 1$  or difference stationary if  $\lambda_n = 1$ . The common time-specific effect experienced by every economy is represented by the term  $a_t$ . Evans assumes that  $\Delta a_t$  is covariance stationary and independent of  $\varepsilon_{nt}$ .

The common trend  $a_t$  for all the *y* variables will be the sole catalyst of economic growth in all economies if  $\lambda_n < 1$ . In this case, growth is exogenous and economies would follow a balanced-growth path. If  $\lambda_n = 1$ , on the other hand, then economy *n* will grow endogenously since  $y_{nt}$  diverges from  $a_t$  and the *y* variables of all remaining economies. The parameter  $\delta_n$ controls for the relative height of economy *n*'s balanced growth path if all the  $\lambda$  s are less than one. If  $\lambda_n = 1$ , then  $\delta_n$  controls for economy *n*'s relative growth rate. The error term  $\omega_n$ measures the portion of  $\delta_n$  that is not explained by  $x_n$ . This error term is assumed to be

<sup>&</sup>lt;sup>1</sup> This appendix borrows heavily from Evans (1997b), which can be consulted for further details. It is not intended for publication.

uncorrelated with  $x_n$ . The inequality  $\lambda_n < 1$  will hold for an economy described by the neoclassical growth model.

Solving equation (1A) backward from year T to year 0, substituting from equation (2A), and rearranging produces

(3A)  

$$g_{n} = \alpha_{n} + \beta_{n} y_{n0} + \gamma_{n}' x_{n} - \frac{\beta_{n} \omega_{n}}{1 - \gamma_{n}} + \frac{1}{T} \sum_{i=0}^{T-1} \lambda_{n}^{i} \left( \sum_{j=0}^{\min[i,q]} \lambda_{n}^{-j} \theta_{nj} \right) \varepsilon_{n,T-i}$$

$$+ \left( \frac{\lambda_{n}^{T}}{T} \right)_{i=0}^{q-1} \lambda_{n}^{i} \left( \sum_{j=i+1}^{q} \lambda_{n}^{-j} \theta_{nj} \right) \varepsilon_{n,-i}$$

where 
$$\beta_n = \frac{\lambda_n^T - 1}{T}$$
,  $\gamma_n = \frac{-\beta_n \xi_n}{1 - \lambda_n}$ , and  $\alpha_n = \frac{a_T - a_0}{T - \beta_n \left(\frac{a_0 + \kappa}{1 - \lambda_n}\right)}$ . If  $\beta_n < 0$ , then economy *n*

grows exogenously  $(\lambda_n < 1)$ . On the other hand, if  $\beta_n = 0$ , then economy *n* grows endogenously  $(\lambda_n = 1)$ .

Now consider a special case in which every intercept  $\delta_n$  is completely explained by the county characteristics included in  $x_n$  ( $\omega_n = 0, \forall n$ ) and every series  $y_{nt} - a_t$  is a first-order autoregression (q = 0). Under these restrictions equation (3A) reduces to:

(4A) 
$$g_n = \alpha_n + \beta_n y_{n0} + \gamma'_n x_n + \frac{1}{T} \sum_{i=0}^{T-1} \lambda_n^i \varepsilon_{n,T-i}$$

The estimator for  $\hat{\beta}$  can then be obtained in two steps. First, regress  $y_{n0}$  on an intercept and  $x_n$ to obtain the residual  $r_n$  and then regress  $g_n$  on  $r_n$ . (This is simply the OLS estimator of  $\beta$ .) Each term in  $\frac{1}{T} \sum_{i=0}^{T-1} \lambda_n^i \varepsilon_{n,T-i}$  is uncorrelated with the intercept,  $y_n$ ,  $x_n$  and the residual  $r_n$ . As a result, one has

(5A) 
$$p \lim_{N \to \infty} \hat{\beta} = \frac{p \lim_{N \to \infty} \frac{1}{N} \sum_{n=1}^{N} \alpha_n r_n + p \lim_{N \to \infty} \frac{1}{N} \sum_{n=1}^{N} \beta_n r_n y_n + p \lim_{N \to \infty} \frac{1}{N} \sum_{n=1}^{N} \gamma'_n r_n x_n}{p \lim_{N \to \infty} \frac{1}{N} \sum_{n=1}^{N} r_n^2}$$

Making further assumptions that  $\alpha_n$  is uncorrelated with  $r_n$ ,  $\beta_n$  is uncorrelated with  $r_n y_n$ , and  $\gamma_n$  is uncorrelated with  $r_n x_n$ , equation (5A) leads to

(6A) 
$$p_{N \to \infty} \hat{\beta} = \frac{p_{N \to \infty} \frac{1}{N} \sum_{n=1}^{N} \beta_n r_n^2}{p_{N \to \infty} \frac{1}{N} \sum_{n=1}^{N} r_n^2}$$

The probability limit of the OLS estimator is then a weighted average of the economy specific  $\beta_n$ s. It is a consistent estimator of that weighted average.<sup>2</sup>

But what if the assumption that every intercept  $\delta_n$  is completely explained by  $x_n$  and also the assumption that every series  $y_{nt} - a_t$  is a first-order auto-regression, are relaxed? Relaxing these assumptions, and imposing the additional restriction that the  $\lambda s$  and  $\xi s$  and, as a result, the  $\beta s$  and  $\gamma s$  are identical across all economies (for the simplicity of the exposition), (3A) can be re-written as

(7A)  

$$g_{n} = \alpha + \beta y_{n0} + \gamma' x_{n} - \frac{\beta \omega_{n}}{1 - \gamma} + \frac{1}{T} \sum_{i=0}^{T-1} \lambda^{i} \left( \sum_{j=0}^{\min[i,q]} \lambda^{-j} \theta_{nj} \right) \varepsilon_{n,T-i}$$

$$+ \left( \frac{\lambda^{T}}{T} \right) \sum_{i=0}^{q-1} \lambda^{i} \left( \sum_{j=i+1}^{q} \lambda^{-j} \theta_{nj} \right) \varepsilon_{n,-i}$$

where  $\beta = \frac{\lambda^T - 1}{T}$ ,  $\gamma = \frac{-\beta\xi}{1 - \lambda}$ , and  $\alpha = \frac{a_T - a_0}{T - \beta \left(\frac{a_0 + \kappa}{1 - \lambda}\right)}$ . Applying the same steps to equation

(6A) yields

(8A) 
$$p \lim_{N \to \infty} \hat{\beta} = \beta + \frac{(\Phi + \Psi)}{p \lim_{N \to \infty} \frac{1}{N} \sum_{n=1}^{N} r_n^2}$$

<sup>&</sup>lt;sup>2</sup> Strictly speaking, even for this restrictive case, an OLS estimate less than unity does not mean that all the economies in the sample conform to the neoclassical growth model. Rather, it would mean that enough economies conform, so that the weighted average is less than unity. It would mean, therefore, that exogenous growth is the predominant case across the sample.

where 
$$\Phi = \frac{\lambda^T}{T} p \lim_{N \to \infty} \frac{I}{N} \sum_{n=1}^{N} \left[ \sum_{i=0}^{q-1} \lambda^i \left( \sum \lambda^{-j} \theta_{n,j+i+1} \right) r_n \varepsilon_{n,-i} \right]$$
 and  $\Psi = -\frac{\beta}{1-\lambda} p \lim_{N \to \infty} \frac{1}{N} \sum r_n \omega_n$ .

As a result, equation (8A) implies that  $p \lim_{N \to \infty} \hat{\beta}$  differs from  $\beta$  if *either* q > 0 ( $y_{nt} - a_t$  is not a first-order AR process) or the cross-sectional variance of  $\omega_n$  is positive (not all cross-sectional heterogeneity is accounted for). In other words, the OLS estimator is inconsistent unless (a) the log of income per capita has an identical first-order AR representation across economies, and (b) all cross-section heterogeneity is controlled for.

Evans shows that the resulting bias from q > 0 is likely to be negligible in practice but the bias resulting from a positive cross-sectional variance for  $\omega_n$  can be substantial. This is essentially an omitted variable bias. Evans demonstrates that

(9A) 
$$p \lim_{N \to \infty} \hat{\beta} = \left[ \frac{\operatorname{var}(y \mid x, \omega)}{\operatorname{var}(y \mid x)} \right] \beta$$

and

(10A) 
$$p \lim_{N \to \infty} \hat{\gamma} = \left[ \frac{\operatorname{var}(y \mid x, \omega)}{\operatorname{var}(y \mid x)} \right] \gamma.$$

The bracketed portions in equations (9A) and (10A) are the ratio of the cross-sectional variance of  $y_{n0}$  conditional on both  $x_n$  and  $\omega_n$  to the cross-sectional variance of  $y_{n0}$  on  $x_n$ . As such,  $\hat{\beta}$ and  $\hat{\gamma}$  will be biased towards zero unless the *xs* are able to control for a large portion of the cross-economy variation in the *ys*.

The intuition here is that if a large portion of the growth of per capita income is explained by variables left out of the OLS regression, then the estimate of the convergence effect will be biased. In general, omitted variable bias can be either positive or negative. However, in this case, theoretically, the bias is negative. Evans (1997b, Tables on p. 11 and p. 15) estimates  $\beta$  for Mankiw, et al.'s (1992) international data using both the OLS, which yields inconsistent estimates, and the 2SLS approach (as outlined in section 2), which yields consistent estimates of both  $\beta$  and  $\gamma$ . He finds that the 2SLS estimate implies a conditional convergence rate between 4 to 5 times as large as the OLS estimate. The bias produced by the OLS in this case, therefore, is substantial.

### **II. Growth Equation Regression Estimates: Complete Tables**

In Table A1 we report the growth equation estimation results for the entire U.S. In Tables A2–A6 we report the regional growth equation estimation results for the Great Lakes Region, Northeastern Region, Plains Region, Southern Region, and the Western Region, respectively.

The information presented in Tables A1–A8 is organized as follows. In the first column of each table, we report the estimation results for all counties together. In the second and third columns we report the estimation results for the metro and non-metro counties, respectively. In each column, we first report the OLS estimation results, with and without Conley-Rappaport standard errors, and then the 2SLS estimation results.

TABLE A1.— GROWTH EQUATION ESTIMATES - E	NTIRE UNITED STATES
--	---------------------

	AII	Metro	Non-Metro
RHS Variables <sup>1</sup>	OLS CR OLS 3SLS	OLS CR OLS 3SLS	OLS CR OLS 3SLS
Constant	$0.1682  ext{ 0.1682  ext{ 0.3320}} (0.0158)^{a} (0.0140)^{a} (0.0149)^{a}$	$\begin{array}{cccc} 0.0859 & 0.0859 & 0.3280 \\ \left( 0.0358  ight)^{ m b} \left( 0.0339  ight)^{ m b} \left( 0.0344  ight)^{ m a} \end{array}$	$\begin{array}{ccc} 0.1901 & 0.1901 & 0.3275 \\ (0.0181)^{a} & (0.0141)^{a} & (0.0171)^{a} \end{array}$
Log 1970 per capita income <sup>2</sup>	-0.0173 -0.0173 -0.0344	-0.0133 -0.0133 -0.0354	-0.0190 -0.0190 -0.0337
	(0.0007) <sup>a</sup> (0.0013) <sup>a</sup> (0.0014) <sup>a</sup>	(0.0016)ª (0.0019)ª (0.0011)ª	(0.0009) <sup>a</sup> (0.0013) <sup>a</sup> (0.0018) <sup>a</sup>
Land area per capita	$\begin{array}{l} -0.0008 & -0.0008 & -0.0008 \\ \left( 0.0001 \right)^a \left( 0.0003 \right)^a \left( 0.0001 \right)_a \end{array}$	-0.0003 -0.0003 -0.0002 (0.0003) (0.0003) (0.0003)	-0.0010 -0.0010 -0.0009 (0.0001) <sup>a</sup> (0.0003) <sup>a</sup> (0.0001) <sup>a</sup>
Water area per capita	0.0025 $0.0025$ $0.0023(0.0007)^{a} (0.0009) (0.0007)^{a}$	0.0056 0.0056 0.0063 (0.0036) (0.0030) <sup>c</sup> (0.0040)	$\begin{array}{cccc} 0.0020 & 0.0020 & 0.0018 \\ (0.0007)^a & (0.0007)^a & (0.0008)^b \end{array}$
Age: 5-13 years	0.0313 0.0313 0.0337	0.0545 0.0545 0.0428	0.0187 0.0187 0.0200
	(0.0181) <sup>°</sup> (0.0219) (0.0195) <sup>°</sup>	(0.0433) (0.0392) (0.0480)	(0.0204) (0.0207) (0.0216)
Age: 14-17 years	0.0286 0.0286 0.0255	0.0334 0.0334 0.0180	0.0280 0.0280 0.0270
	(0.0162) <sup>c</sup> (0.0172) <sup>c</sup> (0.0175)	(0.0332) (0.0319) (0.0368)	(0.0189) (0.0175) (0.0201)
Age: 18-64 years	0.0136 0.0136 0.0074	0.0340 0.0340 0.0111	0.0094 0.0094 0.0053
	(0.0134) (0.0114) (0.0144)	(0.0309) (0.0296) (0.0342)	(0.0152) (0.0135) (0.0162)
Age: 65+	0.0116 0.0116 -0.0039	0.0305 0.0305 0.0113	0.0074 0.0074 -0.0084
	(0.0123) (0.0109) (0.0132)	(0.0285) (0.0237) (0.0316)	(0.0140) (0.0131) (0.0148)
Blacks	-0.0000 -0.0000 0.0022	0.0033 $0.0033$ $0.0065$	-0.0019 -0.0019 -0.0000
	(0.0013) (0.0016) (0.0013)	( $0.0025$ ) ( $0.0020$ ) <sup>c</sup> ( $0.0028$ ) <sup>a</sup>	(0.0015) (0.0020) (0.0016)
Hispanic	$-0.0057$ $-0.0057$ $-0.0059$ $(0.0014)^{a}$ $(0.0014)^{a}$	-0.0024 -0.0024 -0.0015 (0.0026) (0.0023) (0.0029)	$\begin{array}{l} \text{-0.0081} & \text{-0.0081} & \text{-0.0082} \\ (0.0019)^{\text{a}} & (0.0021)^{\text{a}} & (0.0020)^{\text{a}} \end{array}$
Education: 9-11 years	$-0.0209$ $-0.0209$ $-0.0204$ $(0.0033)^{a}$ $(0.0089)^{b}$ $(0.0035)^{a}$	-0.0250 -0.0250 -0.0173 (0.0070) <sup>a</sup> (0.0081) <sup>a</sup> (0.0078) <sup>a</sup>	$\begin{array}{l} \text{-0.0195} & \text{-0.0195} & \text{-0.0202} \\ \left(0.0038\right)^{a} & \left(0.0099\right)^{b} & \left(0.0041\right)^{a} \end{array}$
Education: H.S. diploma	0.0007 $0.0007$ $0.0091(0.0027) (0.0052) (0.0029)a$	-0.0017 -0.0017 0.0022 (0.0057) (0.0049) (0.0064)	$\begin{array}{cccc} 0.0012 & 0.0012 & 0.0106 \\ (0.0032) & (0.0053)^{\mathrm{b}} & (0.0033)^{\mathrm{a}} \end{array}$
Education: Some college	-0.0106 -0.0106 -0.0014	-0.0029 0.0029 0.0009	-0.0067 -0.0067 0.0031
	(0.0056) <sup>c</sup> (0.0089) (0.0061)	(0.0122) (0.0128) (0.0136)	(0.0064) (0.0097) (0.0068)
Education: Bachelor +	0.0424 0.0424 0.0700 (0.0058) <sup>a</sup> (0.0108) <sup>a</sup> (0.0061) <sup>a</sup>	0.0629 0.0629 0.1151 (0.0111) <sup>a</sup> (0.0175) (0.0116) <sup>a</sup>	$\begin{array}{cccc} 0.0326 & 0.0326 & 0.0544 \\ (0.0072)^{a} & (0.0166)^{b} & (0.0075)^{a} \end{array}$
Education: Public elementary	-0.0000 -0.0000 -0.0000	-0.0000 -0.0000 -0.0000	0.0000 0.0000 0.0000
	(0.0000) (0.0000) (0.0000)	(0.0000) (0.0000) (0.0000)	(0.0000) (0.0000) (0.0000)
Education: Public nursery	-0.0000 -0.0000 -0.0000	0.0000 0.0000 0.0000	-0.0000 -0.0000 -0.0000
	(0.0000) (0.0000) (0.0000)	(0.0000) (0.0000) (0.0000)	(0.0000) (0.0000) (0.0000)
Education: Private elementary	0.0000 0.0000 0.0000	0.0000 0.0000 0.0000	0.0000 0.0000 0.0000
	(0.0000) (0.0000) (0.0000)	(0.0000) (0.0000) (0.0000)	(0.0000) (0.0000) (0.0000)
Education: Private nursery	0.0000 0.0000 0.0000	0.0000 0.0000 0.0000	-0.0000 -0.0000 -0.0000
	(0.0000) (0.0000) (0.0000)	(0.0000) (0.0000) (0.0000)	(0.0000) (0.0000) (0.0000)
Housing	-0.0000 -0.0000 0.0000	-0.0000 -0.0000 0.0000	-0.0000 -0.0000 -0.0000
	(0.0000) (0.0000) (0.0000)	(0.0000) (0.0000) (0.0000)	(0.0000) (0.0000) (0.0000)
Federal government employment	$\begin{array}{l} -0.0145 & -0.0145 & -0.0226 \\ (0.0048)^{a} & (0.0046)^{a} & (0.0051)^{a} \end{array}$	-0.0095 -0.0095 -0.0300 (0.0098) (0.0105) (0.0107)ª	$\begin{array}{ccc} -0.0137 & -0.0137 & -0.0179 \\ (0.0056)^{\text{b}} & (0.0049)^{\text{a}} & (0.0060)^{\text{a}} \end{array}$
State government	$\begin{array}{rrr} -0.0040 & -0.0040 & -0.0177 \\ (0.0037) & (0.0045) & (0.0040)^a \end{array}$	-0.0058 -0.0058 -0.0264	0.0021 0.0021 -0.0081
employment		(0.0071) (0.0049) (0.0076) <sup>a</sup>	(0.0004) (0.0064) (0.0048) <sup>c</sup>
Local government	-0.0211 -0.0211 -0.0198	-0.0141 -0.0141 -0.0214	$\begin{array}{rrr} \text{-0.0165} & \text{-0.0165} & \text{-0.0128} \\ (0.0055)^{a} & (0.0087)^{c} & (0.0059)^{b} \end{array}$
employment	(0.0048) <sup>a</sup> (0.0079) <sup>a</sup> (0.00252) <sup>a</sup>	(0.0108) (0.0117) (0.0120)°	

	AII			Metro		N	lon-Metro_	
RHS Variables <sup>3</sup>	OLS CR OLS	<u>3SLS</u>	<u>OLS</u>	<u>CR OLS</u>	<u>3SLS</u>	<u>OLS</u>	<u>CR OLS</u>	<u>3SLS</u>
Self-employment	0.0076 0.0076 (0.0032) <sup>b</sup> (0.0053)	0.0025 (0.0034)		0.0075 (0.0082)	-0.0029 (0.0082)	0.0099 (0.0037) <sup>8</sup>	0.0099 (0.0051) <sup>c</sup>	0.0066 (0.0039)
Agriculture	-0.0109 -0.0109 (0.0061) <sup>c</sup> (0.0059)	-0.0038 ° (0.0066)		0.0006 (0.0076)	0.0035 (0.0126)		-0.0136 (0.0083)	-0.0057 (0.0079)
Communications	-0.0276 -0.0276 (0.0096) <sup>a</sup> (0.0134)	-0.0205 <sup>b</sup> (0.0103) <sup>b</sup>		0.0033 (0.0213)	-0.0107 (0.0237)	-0.0311 (0.0109) <sup>6</sup>	-0.0311 <sup>a</sup> (0.0132) <sup>b</sup>	-0.0022 (0.0115)
Construction	0.0183 0.0183 (0.0068) <sup>a</sup> (0.0058)	0.0120 <sup>a</sup> (0.0073) <sup>c</sup>	0.0607 (0.0128) <sup>a</sup>	0.0607 (0.0085) <sup>a</sup>	0.0486 (0.0142) <sup>a</sup>	0.0077 (0.0082)	0.0077 (0.0064)	0.0052 (0.0087)
Finance, insurance & real estate	0.0632 0.0632 (0.0117) <sup>a</sup> (0.0233)	0.0731 ª (0.0125)ª	0.0600 (0.0213) <sup>a</sup>	0.0600 (0.0243) <sup>b</sup>	0.0600 (0.0237) <sup>b</sup>	0.0608 (0.0143) <sup>8</sup>	0.0608 (0.0258) <sup>b</sup>	0.0699 (0.0152)
Manufacturing – durables	0.0010 0.0010 (0.0057) (0.0077)	-0.0044 (0.0062)	0.0184 (0.0109) <sup>c</sup>	0.0184 (0.0104) <sup>c</sup>	-0.0015 (0.0120)	-0.0026 (0.0070)	-0.0026 (0.0100)	-0.0046 (0.0074)
Manufacturing – nondurables	-0.0079 -0.0079 (0.0059) (0.0064)	-0.0141 (0.0062) <sup>b</sup>		0.0038 (0.0110)	-0.0144 (0.0121)	-0.0099 (0.0072)	-0.0099 (0.0092)	-0.0123 (0.0076)
Mining	-0.0109 -0.0109 (0.0061) <sup>c</sup> (0.0077)	-0.0148 (0.0066) <sup>b</sup>		0.0061 (0.0133)	-0.0206 (0.0139)		-0.0145 (0.0092)	-0.0147 (0.0078) <sup>0</sup>
Retail	-0.0150 -0.0150 (0.0068) <sup>b</sup> (0.0074)			-0.0076 (0.0113)	-0.0264 (0.0145) <sup>°</sup>	-0.0140 (0.0082) <sup>0</sup>	-0.0140 (0.0098)	-0.0155 (0.0087)
Business & repair services	-0.0059 -0.0059 (0.0117) (0.0157)	-0.0008 (0.0126)		0.0109 (0.0242)	0.0094 (0.0267)	-0.0145 (0.0136)	-0.0145 (0.0187)	-0.0075 (0.0144)
Educational services	-0.0257 -0.0257 (0.0082) <sup>a</sup> (0.0060)	-0.0445 <sup>a</sup> (0.0087) <sup>a</sup>	-0.0246 (0.0121) <sup>b</sup>	-0.0246 (0.0154)	-0.0577 (0.0131) <sup>a</sup>	-0.0232 (0.0124) <sup>6</sup>	-0.0232 (0.0108) <sup>b</sup>	-0.0335 (0.0132) <sup>t</sup>
Professional related services	0.0092 0.0092 (0.0089) (0.0079)	0.0099 (0.0095)		0.0063 (0.0097)	0.0000 (0.0151)	0.0057 (0.0129)	0.0057 (0.0128)	0.0029 (0.0137)
Health services	-0.0070 -0.0070 (0.0085) (0.0087)	-0.0079 (0.0092)		0.0061 (0.0165)	-0.0027 (0.0143)	-0.0082 (0.0125)	-0.0082 (0.0117)	-0.0022 (0.0133)
Personal services	0.0317 0.0317 (0.0078) <sup>a</sup> (0.0105)	0.0341 <sup>a</sup> (0.0084) <sup>a</sup>		0.0172 (0.0102) <sup>c</sup>	0.0144 (0.0169)	0.0399 (0.0092) <sup>6</sup>	0.0399 (0.0157) <sup>b</sup>	0.0445 (0.0098) <sup>6</sup>
Entertainment & recreational services	0.0272 0.0272 (0.0154) <sup>c</sup> (0.0230)	0.0335 (0.0166) <sup>b</sup>	0.0778 (0.0314) <sup>b</sup>	0.0778 (0.0251) <sup>a</sup>	0.0670 (0.0349)°	0.0134 (0.0177)	0.0134 (0.0233)	0.0229 (0.0188)
Transportation	-0.0077 -0.0077 (0.0079) (0.0098)		0.0293 (0.0154) <sup>°</sup>	0.0293 (0.0218)	0.01882 (0.0171)		-0.0163 (0.0118)	-0.0228 (0.0099) <sup>1</sup>
Wholesale trade	0.0402 0.0402 (0.0098) <sup>a</sup> (0.0102)	0.0402 <sup>a</sup> (0.0105) <sup>a</sup>		0.0240 (0.0106)b	0.0085 0 (0.0195)	0.0408 (0.0120) <sup>8</sup>	0.0408 (0.0135) <sup>a</sup>	0.0448 (0.0128)
Poverty	-0.0325 -0.0325 (0.0048) <sup>a</sup> (0.0060)	-0.0603 ª (0.0049)ª		-0.0129 (0.0068) <sup>c</sup>	-0.0484 (0.0111) <sup>a</sup>	-0.0341 (0.0054) <sup>*</sup>	-0.0341 a (0.0058)ª	-0.0565 (0.0056) <sup>8</sup>
College Town	0.0007 0.0007 (0.0005) (0.0005)	0.0009 (0.0006)		0.0009 (0.0006)	0.0013 (0.0007) <sup>c</sup>	0.0003 (0.0009)	0.0003 (0.0004)	0.0001 (0.0010)
Metro area, 1970	0.0011 0.0011 (0.0003) <sup>a</sup> (0.0004)	0.0013 <sup>a</sup> (0.0003) <sup>a</sup>		0.0005 (0.0005	0.0005 (0.0005)	-0.0000 (0.0005)	-0.0000 (0.0004)	0.0001 (0.0006)
R <sup>2</sup> # Observations	0.47 0.92 3,058 3,058	0.72 3,058		0.94 867	0.80 867	0.51 2,191	0.91 2,191	0.66 2,191

TABLE A1.—GROWTH EQUATION ESTIMATES - ENTIRE UNITED STATES (CONTINUED)

<sup>&</sup>lt;sup>3</sup> All RHS variables are initial values from 1970. Also, state dummies were employed in all regressions. 7

	AII	Metro	Non-Metro
RHS Variables <sup>1</sup>	OLS CR OLS 3SLS	OLS CR OLS 3SLS	OLS CR OLS 3SLS
Constant	$\begin{array}{cccc} 0.1143 & 0.1143 & 0.2687 \\ (0.0475)^{\text{b}} & (0.0361)^{\text{a}} & (0.0397)^{\text{a}} \end{array}$	-0.0518 -0.0518 01648 (0.0932) (0.0417) (0.0894)°	$\begin{array}{ccc} 0.1693 & 0.1693 & 0.2634 \\ (0.0608)^{a} & (0.0339)^{a} & (0.0505)^{a} \end{array}$
Log 1970 per capita income <sup>2</sup>	-0.0164 -0.0164 -0.0289 (0.0023) <sup>a</sup> (0.0016) <sup>a</sup> (0.0017) <sup>a</sup>	$\begin{array}{l} \text{-0.0106} & \text{-0.0106} & \text{-0.0328} \\ (0.0047)^{\text{b}} & (0.0014)^{\text{a}} & (0.0025)^{\text{a}} \end{array}$	$\begin{array}{rrr} -0.0195 & -0.0195 & -0.0273 \\ (0.0029)^{a} & (0.0019)^{a} & (0.0019)^{a} \end{array}$
Land area per capita	$\begin{array}{cccc} 0.0111 & 0.0111 & 0.0058 \\ (0.0051)^{\mathrm{b}} & (0.0026)^{\mathrm{a}} & (0.0052) \end{array}$	-0.0689 -0.0689 -0.0793 (0.0438) (0.0229) <sup>a</sup> (0.0483)	0.0068 0.0068 0.0039 (0.0058) (0.0024) <sup>a</sup> (0.0058)
Water area per capita	0.0006 0.0006 0.0010 (0.0008) (0.0005) (0.0008)	0.0328 0.0328 0.0491 (0.0449) (0.0081) <sup>a</sup> (0.0494)	$\begin{array}{cccc} 0.0010 & 0.0010 & 0.0013 \\ (0.0009) & (0.0004)^{\mathrm{b}} & (0.0009) \end{array}$
Age: 5-13 years	0.1074 $0.1074$ $0.0556(0.0534)^{b} (0.0304)^{a} (0.0545)$	0.1761 0.1761 0.1528 (0.1081) (0.0439)ª (0.1193)	$\begin{array}{rrrr} 0.0526 & 0.0526 & 0.0137 \\ (0.0673) & (0.0256)^{\text{b}} & (0.0666) \end{array}$
Age: 14-17 years	$0.1233$ $0.1233$ $0.0744$ $(0.0433)^{a}$ $(0.0263)^{a}$ $(0.0440)^{c}$	0.2577 $0.2577$ $0.2952(0.0959)^{a} (0.0537)^{a} (0.1060)^{b}$	0.0751 0.0751 0.0397 0.0528) (0.0221) <sup>a</sup> (0.0519)
Age: 18-64 years	0.0632 0.0632 0.0212 (0.0392) (0.0181) <sup>a</sup> (0.0400)	0.1130 0.1130 0.1246 (0.0807) (0.0268) <sup>a</sup> (0.0891)	0.0252 0.0252 -0.0059 (0.0492) (0.0191) (0.0484)
Age: 65+	0.0252 0.0252 -0.0161 (0.0369) (0.0235) (0.0374)	0.1315 0.1315 0.1340 (0.0787)° (0.0366)° (0.0870)	-0.0074 -0.0074 -0.0380 (0.0463) (0.0213) (0.0455)
Blacks	$\begin{array}{l} \text{-0.0176} & \text{-0.0176} & \text{-0.0144} \\ (0.0062)^{a} & (0.0036)^{a} & (0.0064)^{b} \end{array}$	$\begin{array}{ccc} \text{-0.0212} & \text{-0.0212} & \text{-0.0238} \\ \text{(0.0111)}^{\text{c}} & \text{(0.0046)}^{\text{a}} & \text{(0.0123)}^{\text{c}} \end{array}$	$\begin{array}{rrr} \text{-0.0086} & \text{-0.0086} & \text{-0.0062} \\ \text{(0.0083)} & \text{(0.0041)}^{\text{b}} & \text{(0.0084)} \end{array}$
Hispanic	-0.0278 -0.0278 -0.0180 (0.0211) (0.0088) <sup>a</sup> (0.0218)	-0.0268 -0.0268 -0.0040 (0.0279) (0.0112) <sup>b</sup> (0.0303)	$\begin{array}{rrrr} \text{-0.0416} & \text{-0.0416} & \text{-0.0345} \\ \text{(0.0349)} & \text{(0.0163)}^{\text{b}} & \text{(0.0352)} \end{array}$
Education: 9-11 years	$\begin{array}{c} \text{-0.0493} & \text{-0.0493} & \text{-0.0462} \\ (0.0091)^{a} & (0.0048)^{a} & (0.0094)^{a} \end{array}$	-0.0651 -0.0651 -0.0704 (0.0207) <sup>a</sup> (0.0097) <sup>a</sup> (0.0229) <sup>a</sup>	$\begin{array}{c} \text{-0.0433} & \text{-0.0433} & \text{-0.0412} \\ \text{(0.0114)}^{\text{a}} & \text{(0.0060)}^{\text{a}} & \text{(0.0115)}^{\text{a}} \end{array}$
Education: H.S. diploma	$-0.0175$ $-0.0175$ $-0.0092$ $(0.0055)^{a}$ $(0.0033)^{a}$ $(0.0055)^{c}$	0.0048 0.0048 0.0126 (0.0119) (0.0056) (0.0131)	$\begin{array}{l} \text{-0.0157} & \text{-0.0157} & \text{-0.0103} \\ (0.0070)^{\text{b}} & (0.0044)^{\text{a}} & (0.0068) \end{array}$
Education: Some college	$0.0328$ 0.0328 0.0423 $(0.0153)^{b}$ $(0.0092)^{a}$ $(0.0158)^{a}$	-0.0237 -0.0237 -0.0261 (0.0334) (0.0137)° (0.0369)	$\begin{array}{llllllllllllllllllllllllllllllllllll$
Education: Bachelor +	0.0398 0.0398 0.0607 (0.0145) <sup>a</sup> (0.0056) <sup>a</sup> (0.0145) <sup>a</sup>	$0.0544$ 0.0544 0.01024 $(0.0303)^{\circ}$ $(0.0142)^{a}$ $(0.0315)^{a}$	$\begin{array}{ccc} 0.0428 & 0.0428 & 0.0513 \\ (0.0195)^{\text{b}} & (0.0074)^{\text{a}} & (0.0194)^{\text{a}} \end{array}$
Education: Public elementary	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	-0.0000 -0.0000 0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)
Education: Public nursery	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)
Education: Private elementary	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)
Education: Private nursery	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)	-0.0000 -0.0000 0.0000 (0.0000) (0.0000) (0.0000)
Housing	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)
Federal government employment	-0.0004 -0.0004 -0.0029 (0.0147) (0.0081) (0.0153)	0.0353 0.0353 0.0114 (0.0315) (0.0136)a (0.0344)	0.0075 0.0075 0.0111 (0.0192) (0.0057) (0.0194)
State government employment	-0.0116 -0.0116 -0.0181 (0.0093) $(0.0059)^{b}$ (0.0096) <sup>c</sup>	0.0016 0.0016 -0.0203 (0.0120) (0.0049) (0.0214)	$\begin{array}{llllllllllllllllllllllllllllllllllll$
Local government employment	-0.0362 -0.0362 -0.0343 (0.0125)° (0.0038)° (0.0130)°	-0.0540 -0.0540 -0.0333 (0.0362) (0.0135)ª (0.0397)	$\begin{array}{r} \text{-0.0345} & \text{-0.0345} & \text{-0.0331} \\ \text{(0.0143)}^{\text{b}} & \text{(0.0061)}^{\text{a}} & \text{(0.0145)}^{\text{b}} \end{array}$

	AII			Metro		N	on-Metro_	
RHS Variables <sup>3</sup>	OLS CR OLS	3SLS	<u>OLS</u>	CR OLS 3S	<u>SLS</u>	<u>OLS</u>	<u>CR OLS</u>	<u>3SLS</u>
Self-employment	-0.0132 -0.0132 (0.0122) (0.0082)	-0.0201 (0.0126)		-0.1041 -0. (0.0243) <sup>a</sup> (0.		-0.0060 (0.0136)		-0.0086 (0.0137)
Agriculture	0.0021 0.0021 (0.0199) (0.0129)	0.0118 (0.0206)		0.1322 0.1 (0.0339) <sup>a</sup> (0.	1720 .0612) <sup>a</sup>	0.0067 (0.0255)	0.0067 (0.0095)	0.0174 (0.0255)
Communications	-0.0054 -0.0054 (0.0266) (0.0154)	-0.0002 (0.0276)		0.0456 0.0 (0.0139) <sup>a</sup> (0.	0647 .0641)	-0.0011 (0.0334)		0.0094 (0.0336)
Construction	0.0061 0.0061 (0.0192) (0.0092)	0.0076 (0.0199)		0.1153 0.1 (0.0189) <sup>a</sup> (0.	1129 .0478) <sup>b</sup>	0.0137 (0.0249)	0.0137 (0.0063) <sup>b</sup>	0.0207 (0.0250)
Finance, insurance & real estate	0.0343 0.0343 (0.0272) (0.0111)	0.0337 ª (0.0282)		0.1013 0.1 (0.0241) <sup>a</sup> (0.		0.0360 (0.0328)	0.0360 (0.0179) <sup>b</sup>	0.0412 (0.0332)
Manufacturing – durables	0.0007 0.0007 (0.0168) (0.0083)	-0.0007 (0.0174)		0.0550 0.0 (0.0140) <sup>a</sup> (0.	0391 .0379)	0.0111 (0.0226)	0.0111 (0.0062) <sup>c</sup>	0.0164 (0.0228)
Manufacturing – nondurables	-0.0016 -0.0016 (0.0168) (0.0076)	-0.0040 (0.0174)		0.0490 0.0 (0.0137) <sup>a</sup> (0.	0333 .0369)	0.0070 (0.0229)	0.0070 (0.0069)	0.0120 (0.0231)
Mining	-0.0167 -0.0167 (0.0179) (0.0091)	-0.0166 ° (0.0186)		0.0426 0.0 (0.0239) <sup>c</sup> (0.	0194 .0514)	-0.0049 (0.0238)		0.0019 (0.0239)
Retail	-0.0162 -0.0162 (0.0193) (0.0058)	-0.0153 ª (0.0201)		0.0167 0.0 (0.0112) (0.	0044 .0460)	0.0049 (0.0254)	0.0049 (0.0063)	0.0123 (0.0256)
Business & repair services	0.0520 0.0520 (0.0404) (0.0219)	0.0576 <sup>b</sup> (0.0419)		0.0564 0.0 (0.0594) (0.	0546 .0912)	0.0420 (0.0510)	0.0420 (0.0198) <sup>b</sup>	0.0540 (0.0515)
Educational services	-0.0896 -0.0896 (0.0316) <sup>a</sup> (0.0141)	-0.1088 ª (0.0326)ª		-0.1091 -0. (0.0275) <sup>a</sup> (0.	.1468 .0834)°		-0.0748 (0.0180) <sup>a</sup>	
Professional related services	0.0679 0.0679 (0.0325) <sup>b</sup> (0.0161)	0.0701 <sup>a</sup> (0.0336) <sup>b</sup>		0.1412 0.1 (0.0307) <sup>a</sup> (0.	1261 .0856)	0.0657 (0.0374) <sup>c</sup>	0.0657 (0.0140) <sup>a</sup>	0.0708 (0.0378) <sup>c</sup>
Health services	-0.0853 -0.0853 (0.0308) <sup>a</sup> (0.0149)			-0.1369 -0. (0.0239) <sup>a</sup> (0.		-0.0627 (0.0362) <sup>c</sup>		-0.6373 (0.0366) <sup>c</sup>
Personal services	0.0223 0.0223 (0.0268) (0.0121)	0.0265 ° (0.0278)		0.1091 0.1 (0.0249) <sup>a</sup> (0.	1057 .0802)	0.0432 (0.0332)	0.0432 (0.0121) <sup>a</sup>	0.0525 (0.0334)
Entertainment & recreational services	0.1002 0.1002 (0.0784) (0.0616)	0.1206 (0.0812)		0.6307 0.5 (0.0858) <sup>a</sup> (0.	5911 .1859) <sup>ª</sup>	0.0115 (0.0944)	0.0115 (0.0628)	0.0358 (0.0952)
Transportation	-0.0137 -0.0137 (0.0210) (0.0107)			0.0109 -0. (0.0104) (0.			-0.0030 (0.0098)	
Wholesale trade	0.0602 0.0602 (0.0240) <sup>b</sup> (0.0101)	0.0614 <sup>a</sup> (0.0248) <sup>b</sup>		0.1147 0.1 (0.0164) <sup>a</sup> (0.	1067 .0499) <sup>b</sup>	0.0671 (0.0317) <sup>b</sup>	0.0671 (0.0110) <sup>a</sup>	0.0708 (0.0320) <sup>b</sup>
Poverty	-0.0228 -0.0228 (0.0082) <sup>a</sup> (0.0059)			0.0072 -0. (0.0240) (0.	.0292 .0274)		-0.0304 (0.0076) <sup>a</sup>	
College Town	-0.0002 -0.0002 (0.0010) (0.0003)			-0.0013 -0. (0.0003) <sup>a</sup> (0.		0.0010 (0.0016)	0.0010 (0.0003) <sup>a</sup>	
Metro area, 1970	0.0010 0.0010 (0.0006) <sup>c</sup> (0.0002)	0.0012 <sup>a</sup> (0.0006) <sup>b</sup>		0.0010 0.0 (0.0004) <sup>b</sup> (0.		0.0002 (0.0009)	0.0002 (0.0003)	0.0002 (0.0009)
R <sup>2</sup> # Observations	0.70 0.96 435 435	0.79 435		0.93 0.8 140 14		0.53 295	0.96 295	0.70 295

 $<sup>^3</sup>$  All RHS variables are initial values from 1970. Also, state dummies were employed in all regressions. \$9\$

TABLE A3.—GROWTH EQUATION ESTIMATES – NORTHEAST REGION
--

	AII	Metro	Non-Metro		
RHS Variables <sup>1</sup>	OLS CR OLS 3SLS	OLS CR OLS 3SLS	OLS CR OLS 3SLS		
Constant	$\begin{array}{rrrr} -0.0626 & -0.0626 & 0.1793 \\ (0.0702) & (0.0559) & (0.0686)^a \end{array}$	-0.1869 -0.1869 0.0934 (0.1292) (0.0730)ª (0.1293)	$\begin{array}{ccc} 0.1697 & 0.1697 & 0.3867 \\ (0.0926)^{\rm c} & (0.0408)^{\rm a} & (0.0890)^{\rm a} \end{array}$		
Log 1970 per capita income <sup>2</sup>	$\begin{array}{cccc} 0.0033 & 0.0033 & -0.0264 \\ (0.0042) & (0.0033) & (0.0024)^a \end{array}$	-0.0000 -0.0000 -0.2684 (0.0079) (0.0025) (0.0041) <sup>a</sup>	$\begin{array}{rrrr} -0.0027 & -0.0027 & -0.0281 \\ (0.0053) & (0.0029) & (0.0028)^a \end{array}$		
Land area per capita	0.0022 0.0022 -0.0010	-0.0621 -0.0621 -0.0470	0.0041 0.0041 0.0017		
	(0.0046) (0.0026) (0.0051)	(0.0419) (0.0107) <sup>a</sup> (0.0466)	(0.0055) (0.0035) (0.0060)		
Water area per capita	0.0305 0.0305 0.0376 (0.0186) (0.0105) <sup>a</sup> (0.0207) <sup>c</sup>	$0.1486  0.1486  0.2111 \\ (0.0850)^{\circ} \ (0.0245)^{a} \ (0.0930)^{b}$	$\begin{array}{rrrr} 0.0349 & 0.0349 & 0.0418 \\ (0.0222) & (0.0161)^{\rm b} & (0.0245)^{\rm c} \end{array}$		
Age: 5-13 years	$0.1577$ $0.1577$ $0.2041$ $(0.0930)^{\circ}$ $(0.0814)^{b}$ $(0.1039)^{\circ}$	0.2684 $0.2684$ $0.3411(0.1676) (0.0782)a (0.1861)c$	$\begin{array}{l} \text{-0.1058} & \text{-0.1058} & \text{-0.0805} \\ \text{(0.1230)} & \text{(0.0520)}^{\text{b}} & \text{(0.1353)} \end{array}$		
Age: 14-17 years	-0.1186 -0.1186 -0.0397	0.1279 $0.1279$ $0.0960$	-0.2200 -0.2200 -0.1177		
	(0.0696) <sup>c</sup> (0.0217) <sup>a</sup> (0.0769)	( $0.1036$ ) ( $0.0656$ ) <sup>b</sup> ( $0.1155$ )	$(0.1010)^{b} (0.0638)^{a} (0.1088)$		
Age: 18-64 years	0.0460 0.0460 0.1058 (0.0626) (0.0488) (0.0695)	0.1663 0.1663 0.2113 (0.1043) (0.0419) <sup>a</sup> (0.1159) <sup>c</sup>	-0.1407 $-0.1407$ $-0.0920$ (0.0843) <sup>c</sup> (0.0339) <sup>a</sup> (0.0922)		
Age: 65+	0.0210 0.0210 0.0410	0.0882 0.0882 0.1086	-0.1139 -0.1139 -0.0998		
	(0.0590) (0.0379) (0.0659)	(0.1194) (0.0416) <sup>b</sup> (0.1336)	(0.0759) (0.0389)ª (0.0835)		
Blacks	-0.0102 -0.0102 0.0049	-0.0660 -0.0660 -0.0560	-0.0059 -0.0059 0.0042		
	(0.0073) (0.0026) <sup>a</sup> (0.0078)	(0.0124) <sup>a</sup> (0.0057) <sup>a</sup> (0.0135) <sup>a</sup>	(0.0092) (0.0033)° (0.0099)		
Hispanic	$\begin{array}{l} -0.0665 & -0.0665 & -0.0547 \\ (0.0208)^a & (0.0057)^a & (0.0232)^b \end{array}$	-0.0950 -0.0950 -0.0811 (0.0202)ª (0.0078)ª (0.0221)ª	$\begin{array}{llllllllllllllllllllllllllllllllllll$		
Education: 9-11 years	0.0120 0.0120 0.0078	0.0770 0.0770 0.0548	-0.0026 -0.0026 -0.0023		
	(0.0137) (0.0067) <sup>°</sup> (0.0153)	(0.0256) <sup>a</sup> (0.0101) <sup>a</sup> (0.0377) <sup>c</sup>	(0.0167) (0.0064) (0.0184)		
Education: H.S. diploma	-0.0090 -0.0090 -0.0167 (0.0104) (0.0047) <sup>c</sup> (0.0116)	$-0.0357$ $-0.0357$ $-0.0414$ $(0.0182)^{\circ}$ $(0.0068)$ $(0.0203)^{ m b}$	-0.0063 -0.0063 -0.0119 (0.0129) (0.0047) (0.0141)		
Education: Some college	-0.0189 -0.0189 -0.0209	0.0406 0.0406 0.0480	0.0037 0.0037 0.0095		
	(0.0258) (0.0137) (0.0288)	(0.0418) (0.0261) (0.0468)	(0.0313) (0.0094) (0.0345)		
Education: Bachelor +	0.0589 0.0589 0.1050 (0.0219)ª (0.0101)ª (0.0234)ª	$0.1037$ $0.1037$ $0.0900$ $(0.0421)^{b}$ $(0.0178)^{a}$ $(0.0469)^{c}$	$\begin{array}{cccc} 0.0746 & 0.0746 & 0.1241 \\ (0.0271)^{a} & (0.0114)^{a} & (0.0276)^{c} \end{array}$		
Education: Public elementary	-0.0000 -0.0000 -0.0000	0.0000 0.0000 -0.0000	0.0000 0.0000 0.0000		
	(0.0000) (0.0000) (0.0000) <sup>a</sup>	(0.0000) (0.0000) (0.0000)	(0.0000) (0.0000) (0.0000)		
Education: Public nursery	0.0000 0.0000 0.0000	0.0000 $0.0000$ $0.0000$	-0.0000 -0.0000 -0.0000		
	(0.0000) (0.0000) (0.0000)	$(0.0000)^{a}$ $(0.0000)$ $(0.0000)^{a}$	(0.0000) (0.0000) (0.0000)		
Education: Private elementary	$\begin{array}{l} -0.0000 & -0.0000 & -0.0000 \\ (0.0000)^{b} & (0.0000) & (0.0000)^{a} \end{array}$	$\begin{array}{llllllllllllllllllllllllllllllllllll$	$\begin{array}{cccc} 0.0000 & 0.0000 & 0.0000 \\ (0.0000)^{\text{b}} & (0.0000) & (0.0000)^{\text{c}} \end{array}$		
Education: Private nursery	$\begin{array}{llllllllllllllllllllllllllllllllllll$	-0.0000 -0.0000 0.0000 (0.0000) (0.0000) (0.0000)	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)		
Housing	0.0000 $0.0000$ $0.0000$	-0.0000 -0.0000 -0.0000	-0.0000 -0.0000 -0.0000		
	$(0.0000)^{c}$ $(0.0000)$ $(0.0000)^{a}$	(0.0000) (0.0000) (0.0000)	(0.0000) (0.0000) (0.0000)		
Federal government	-0.0118 -0.0118 -0.0252	0.0304 0.0304 0.0163	-0.0112 -0.0112 -0.0267		
employment	(0.0184) (0.0094) (0.0205)	(0.0322) (0.0171)° (0.0357)	(0.0231) (0.0061)° (0.0252)		
State government	0.0010 0.0010 -0.0095	-0.0107 -0.0107 -0.0172	0.0124 0.0124 0.0049		
employment	(0.0122) (0.0060) (0.0135)	(0.0268) (0.0110) (0.0299)	(0.0140) (0.0033) <sup>a</sup> (0.0153)		
Local government	-0.0261 -0.0261 -0.0431	-0.0511 -0.0511 -0.1152 (0.0564) (0.0189) $^{\rm a}$ (0.0596) $^{\rm c}$	-0.0300 -0.0300 -0.0477		
employment	(0.0235) (0.0169) (0.0262)		(0.0304) (0.0146) <sup>b</sup> (0.0334)		

### TABLE A3.—GROWTH EQUATION ESTIMATES – NORTHEAST REGION (CONTINUED)

	AII		N	Metro	N	Ion-Metro
RHS Variables <sup>3</sup>	OLS CR OLS	3SLS	<u>OLS</u> C	ROLS 3SLS	<u>OLS</u>	CR OLS 3SLS
Self-employment		-0.0180 <sup>b</sup> (0.0305)		0.0666 -0.0041 0.0369) (0.0688)	-0.0477 (0.0329)	$\begin{array}{r} \text{-0.0477} & \text{-0.0452} \\ (0.0140)^{a} & (0.0362) \end{array}$
Agriculture	0.0292 0.0292 (0.0297) (0.0147)	0.0173 <sup>b</sup> (0.0332)		0.0610 0.0655 0.0395) (0.0829)	0.0365 (0.0374)	$\begin{array}{ccc} 0.0365 & 0.0184 \\ \left( 0.0195  ight)^{ m b} & \left( 0.0410  ight) \end{array}$
Communications	0.0721 0.0721 (0.0457) (0.0158)	0.1049 <sup>a</sup> (0.0509) <sup>b</sup>		0.1286 0.1943 0.0290) <sup>a</sup> (0.0959) <sup>b</sup>	0.0446 (0.0556)	$\begin{array}{ccc} 0.0446 & 0.0527 \\ (0.0211)^{b} & (0.0612) \end{array}$
Construction	$\begin{array}{ccc} 0.0533 & 0.0533 \\ (0.0257)^{ m b} & (0.0069) \end{array}$	0.0600 (0.0287) <sup>b</sup>		0.0401 0.0470 0.0318) (0.0860)	0.0348 (0.0312)	$\begin{array}{ccc} 0.0348 & 0.0289 \\ \left( 0.0085  ight)^{a} & \left( 0.0344  ight) \end{array}$
Finance, insurance & real estate	-0.0201 -0.0201 (0.0313) (0.0106)	-0.0177 (0.0350)		0.0149 0.0358 0.0266) (0.0562)	-0.0352 (0.0514)	-0.0352 -0.0420 (0.0251) (0.0566)
Manufacturing – durables	-0.0102 -0.0102 (0.0203) (0.0094)	-0.0163 (0.0227)		0.0103 0.0381 0.0178) (0.0509)	-0.0147 (0.0254)	-0.0147 -0.0309 (0.0074) <sup>b</sup> (0.0277)
Manufacturing – nondurables	-0.0054 -0.0054 (0.0207) (0.0107)			0.0217 0.0530 0.0188) (0.0538)	-0.0091 (0.0255)	-0.0091 -0.0302 (0.0099) (0.0277)
Mining	0.0039 0.0039 (0.0263) (0.0129)	-0.0185 (0.0292)		0.0304 0.0539 0.0277) (0.0751)	-0.0062 (0.0332)	-0.0062 -0.0394 (0.0083) (0.0358)
Retail	-0.0325 -0.0325 (0.0261) (0.0081)			0.0367 0.0002 0.0233) (0.0668)	-0.0173 (0.0331)	-0.0173 -0.0249 (0.0094) <sup>c</sup> (0.0364)
Business & repair services	-0.0054 -0.0054 (0.0577) (0.0248)	0.0273 (0.0643)		0.3629 0.3966 0.0457) <sup>a</sup> (0.1248) <sup>a</sup>	-0.0014 (0.0702)	-0.0014 0.0445 (0.0207) (0.0766)
Educational services	-0.0536 -0.0536 (0.0468) (0.0232)	-0.0870 <sup>b</sup> (0.0522) <sup>c</sup>		0.0616 -0.1225 0.0349) <sup>°</sup> (0.1500)	-0.0873 (0.0528)	-0.0873 -0.1009 (0.0152) (0.0580) <sup>c</sup>
Professional related services	0.0379 0.0379 (0.0498) (0.0223)	0 .0159 ° (0.0556)	-0.0821 -0 (0.1543) (0	0.0821 0.1167 0.0374) <sup>♭</sup> (0.1598)	0.0567 (0.0561)	0.0567 0.0032 (0.0118) (0.0605)
Health services	0.0008 0.0008 (0.0465) (0.0226)	0 .0128 (0.0520)		0.2066 0.0351 0.0339)ª (0.1586)	-0.0286 (0.0519)	-0.0286 0.0019 (0.0145) <sup>b</sup> (0.0567)
Personal services	-0.0289 -0.0289 (0.0293) (0.0153)	-0.0341 (0.0328)		0.1337 0.1335 0.0815) (0.1278)	-0.0078 (0.0329)	-0.0078 -0.0153 (0.0118) (0.0362)
Entertainment & recreational services	0.2028 0.2028 (0.1087) <sup>c</sup> (0.0602)	0.1964 )ª (0.1217)		).1034 0.2496 ).0666) (0.2186)	-0.0157 (0.1355)	-0.0157 -0.1018 (0.0530) (0.1480)
Transportation	0.0038 0.0038 (0.0293) (0.0089)		0.1494 0 (0.0678) <sup>b</sup> (0	0.1494 0.1878 0.0249) <sup>a</sup> (0.0749) <sup>b</sup>	-0.0239 (0.0361)	-0.0239 -0.0540 (0.0089) <sup>a</sup> (0.0391)
Wholesale trade	0.0257 0.0257 (0.0394) (0.0242)	0.0131 (0.0441)	-0.0389 -0 (0.0783) (0	0.0389 0.0082 0.0372) (0.0863)	0.0386 (0.0518)	$\begin{array}{ccc} 0.0386 & 0.0007 \\ \left( 0.0177  ight)^{ m b} & \left( 0.0564  ight) \end{array}$
Poverty	0.0219 0.0219 (0.0196) (0.0136)	-0.0526 (0.0186) <sup>c</sup>		0.0992 -0.0145 0.0231)ª (0.0431)	-0.0011 (0.225)	-0.0011 -0.0549 (0.0241) (0.0215) <sup>c</sup>
College Town	-0.0005 -0.0005 (0.0010) (0.0037)	-0.0005 (0.0012)		0.0011 -0.0026 0.0057) (0.0013) <sup>b</sup>	-0.0016 (0.0019)	-0.0016 -0.0009 (0.0007) <sup>b</sup> (0.0021)
Metro area, 1970	0.0003 0.0003 (0.0007) (0.0001)	0.0006 <sup>a</sup> (0.0008)		0.0001 -0.0005 0.0005) (0.0012)	0.0003 (0.0010)	0.0003 0.0004 (0.0006) (0.0011)
R <sup>2</sup> # Observations	0.65 0.97 244 244	0.85 244		.98 0.96 0 90	0.67 154	0.97 0.85 154 154

<sup>&</sup>lt;sup>3</sup> All RHS variables are initial values from 1970. Also, state dummies were employed in all regressions.

	AII	Metro	Non-Metro
RHS Variables <sup>1</sup>	OLS CR OLS 3SLS	OLS <u>CR OLS</u> <u>3SLS</u>	OLS CR OLS 3SLS
Constant	$0.1491  0.1491  0.3661 \\ (0.0281)^{a} (0.0126)^{a} (0.0262)^{a}$	$\begin{array}{cccc} 0.2023 & 0.2023 & 0.4218 \\ (0.0749)^{a} & (0.0308)^{a} & (0.0664)^{a} \end{array}$	$\begin{array}{ccc} 0.1526 & 0.1526 & 0.3492 \\ (0.0300)^{a} & (0.0189)^{a} & (0.0288)^{a} \end{array}$
Log 1970 per capita income <sup>2</sup>	-0.0170 -0.0170 -0.0381 (0.0015) <sup>a</sup> (0.0014) <sup>a</sup> (0.0011) <sup>a</sup>	-0.0163 -0.0163 -0.0344 (0.0034) <sup>a</sup> (0.0027) <sup>a</sup> (0.0014) <sup>a</sup>	$\begin{array}{l} \text{-0.0169} & \text{-0.0169} & \text{-0.0383} \\ (0.0016)^{a} & (0.0014)^{a} & (0.0013)^{a} \end{array}$
Land area per capita	$\begin{array}{rrrr} -0.0083 & -0.0083 & -0.0161 \\ (0.0033) & (0.0042) & (0.0036)^a \end{array}$	$\begin{array}{rrr} -0.0144 & -0.0144 & -0.0176 \\ (0.0093) & (0.0059)^{\rm b} & (0.0099)^{\rm c} \end{array}$	$\begin{array}{r} \text{-0.0067} & \text{-0.0067} & \text{-0.0166} \\ (0.0036)^{\text{c}} & (0.0040)^{\text{c}} & (0.0039)^{\text{a}} \end{array}$
Water area per capita	0.0010 0.0010 0.0036 (0.0069) (0.0052) (0.0076)	$0.0377  0.0377  0.0554 \\ (0.0099)^{a} \ (0.0163)^{b} \ (0.0316)^{c}$	0.0007 0.0007 0.0064 (0.0069) (0.0048) (0.0076)
Age: 5-13 years	$0.0635$ $0.0635$ $0.0671$ $(0.0326)^{\circ}$ $(0.0196)^{a}$ $(0.0360)^{\circ}$	-0.0528 -0.0528 -0.0870 (0.0907) (0.0464) (0.0964)	$\begin{array}{ccc} 0.0658 & 0.0658 & 0.0829 \\ (0.0337)^{c} & (0.0235)^{a} & (0.0375)^{b} \end{array}$
Age: 14-17 years	$\begin{array}{rrrr} \text{-0.0198} & \text{-0.0198} & \text{-0.0592} \\ \text{(0.0284)} & \text{(0.0187)} & \text{(0.0312)}^{\text{c}} \end{array}$	$\begin{array}{rrrr} \text{-0.1036} & \text{-0.1036} & \text{-0.1046} \\ (0.0709) & (0.0428)^{\text{b}} & (0.0756) \end{array}$	$\begin{array}{rrrr} \text{-0.0117} & \text{-0.0117} & \text{-0.0559} \\ (0.0306) & (0.0164) & (0.0339)^{\text{c}} \end{array}$
Age: 18-64 years	0.0167 0.0167 -0.0044 (0.0230) (0.0110) (0.0254)	-0.0400 -0.0400 -0.0862 (0.0623) (0.0269) (0.0657)	0.0092 0.0092 0.0077 (0.0245) (0.0149) (0.0272)
Age: 65+	$\begin{array}{llllllllllllllllllllllllllllllllllll$	-0.0689 -0.0689 -0.0996 (0.0593) (0.0255) <sup>a</sup> (0.0628)	$0.0397$ $0.0397$ $0.0295$ $(0.0227)^{\circ}$ $(0.0135)^{a}$ $(0.0253)$
Blacks	$\begin{array}{r} \text{-0.0034} & \text{-0.0034} & \text{-0.0013} \\ \text{(0.0017)}^{\text{b}} & \text{(0.0011)}^{\text{a}} & \text{(0.0018)} \end{array}$	0.0028 0.0028 -0.0004 (0.0050) (0.0025) (0.0053)	$\begin{array}{c} \text{-0.0069} & \text{-0.0069} & \text{-0.0034} \\ (0.0019)^a & (0.0010)^a & (0.0021)^c \end{array}$
Hispanic	0.0126 0.0126 0.0187 (0.0177) (0.0147) (0.0196)	-0.0504 -0.0504 -0.0424 (0.0275)° (0.0159) <sup>a</sup> (0.0293)	-0.0173 -0.0173 -0.0376 (0.0472) (0.0451) (0.0525)
Education: 9-11 years	$\begin{array}{rrrr} 0.0074 & 0.0074 & 0.0175 \\ (0.0054) & (0.0043) & (0.0059)^a \end{array}$	0.0001 0.0001 0.0081 (0.0151) (0.0058) (0.0160)	$\begin{array}{rrrr} 0.0119 & 0.0119 & 0.0211 \\ (0.0057)^{\mathrm{b}} & (0.0035)^{\mathrm{a}} & (0.0063)^{\mathrm{a}} \end{array}$
Education: H.S. diploma	$\begin{array}{llllllllllllllllllllllllllllllllllll$	$\begin{array}{llllllllllllllllllllllllllllllllllll$	$\begin{array}{cccc} 0.0062 & 0.0062 & 0.0154 \\ (0.0055) & (0.0047) & (0.0061)^{\rm b} \end{array}$
Education: Some college	$\begin{array}{llllllllllllllllllllllllllllllllllll$	-0.0059 -0.0059 -0.0098 (0.0359) (0.0128) (0.0382)	$\begin{array}{ccc} 0.0437 & 0.0437 & 0.0437 \\ (0.0144)^{a} & (0.0097)^{a} & (0.0160)^{a} \end{array}$
Education: Bachelor +	$\begin{array}{llllllllllllllllllllllllllllllllllll$	$\begin{array}{llllllllllllllllllllllllllllllllllll$	$0.0501  0.0501  0.0911 \\ (0.0150)^{a} \ (0.0061)^{a} \ (0.0163)^{a}$
Education: Public elementary	$\begin{array}{rrrr} -0.0000 & -0.0000 & -0.0000 \\ (0.0000) & (0.0000) & (0.0000)^{\mathrm{b}} \end{array}$	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)
Education: Public nursery	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)
Education: Private elementary	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	$\begin{array}{cccc} 0.0000 & 0.0000 & 0.0000 \\ (0.0000)^{\circ} & (0.0000) & (0.0000) \end{array}$
Education: Private nursery	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)
Housing	$\begin{array}{llllllllllllllllllllllllllllllllllll$	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	-0.0000 -0.0000 0.0000 (0.0000) (0.0000) (0.0000)
Federal government employment	$\begin{array}{rrr} -0.0133 & -0.0133 & -0.0208 \\ (0.0082) & (0.0054)^{\text{b}} & (0.0090)^{\text{b}} \end{array}$	-0.0213 -0.0213 -0.0345 (0.0200) (0.0097) <sup>b</sup> (0.0211)	-0.0128 -0.0128 -0.0132 (0.0093) (0.0073) <sup>°</sup> (0.0103)
State government employment	0.0062 0.0062 -0.0048 (0.0069) (0.0059) (0.0075)	0.0028 0.0028 -0.0153 (0.0172) (0.0087) (0.0179)	0.0044 0.0044 -0.0043 (0.0077) (0.0055) (0.0085)
Local government employment	0.0033 0.0033 0.0009 (0.0101) (0.0051) (0.0111)	-0.0001 -0.0001 0.0041 (0.0328) (0.0155) (0.0349)	0.0067 0.0067 0.0061 (0.0104) (0.0046) (0.0116)

	All			Metro		N	on-Metro_	
RHS Variables <sup>3</sup>	OLS CR OLS	<u>3SLS</u>	<u>OLS</u>	CR OLS	<u>3SLS</u>	<u>OLS</u>	CR OLS	<u>3SLS</u>
Self-employment	-0.0011 -0.0011	-0.0060	0.0260	0.0260	0.0161	-0.0027	-0.0027	-0.0091
	(0.0063) (0.0058)	(0.0070)	(0.0207)	(0.0163)	(0.0220)	(0.0066)	(0.0024)	(0.0073)
Agriculture	-0.0039 -0.0039	-0.0021	0.0039	0.0039	-0.0034	-0.0086	-0.0086	0.0039
	(0.0108) (0.0099)	(0.0119)	(0.0248)	(0.0119)	(0.0263)	(0.0125)	(0.0108)	(0.0139)
Communications	-0.0477 -0.0477	-0.0390	-0.0183	-0.0183	-0.0478	-0.0492	-0.0492	-0.0299
	(0.0142) <sup>a</sup> (0.0071) <sup>a</sup>	' (0.0157) <sup>b</sup>	(0.0410)	(0.0191)	(0.0433)	(0.0150) <sup>6</sup>	' (0.0089) <sup>a</sup>	' (0.0166) <sup>c</sup>
Construction	0.0244 0.0244	0.0123	0.0261	0.0261	0.0098	0.0156	0.0156	0.0117
	(0.0111) <sup>b</sup> (0.0075) <sup>s</sup>	(0.0122)	(0.0245)	(0.0114) <sup>t</sup>	(0.0259)	(0.0131)	(0.0101)	(0.0145)
Finance, insurance	0.0246 0.0246	0.0429	0.0404	0.0404	0.0652	-0.0125	-0.0125	-0.0040
& real estate	(0.0224) (0.0111) <sup>t</sup>	' (0.0246) <sup>c</sup>	(0.0536)	(0.0310)	(0.0569)	(0.0245)	(0.0108)	(0.0273)
Manufacturing –	-0.0025 -0.0025	-0.0125	0.0092	0.0092	-0.0111	-0.0082	-0.0082	-0.0091
durables	(0.0101) (0.0103)	(0.0111)	(0.0223)	(0.0093)	(0.0234)	(0.0116)	(0.0122)	(0.0130)
Manufacturing –	-0.0112 -0.0112	-0.0223	-0.0024	-0.0024	-0.0208	-0.0157	-0.0157	-0.0188
nondurables	(0.0100) (0.0094)	(0.0110) <sup>b</sup>	(0.0225)	(0.0086)	(0.0237)	(0.0116)	(0.0121)	(0.0129)
Mining	-0.0128 -0.0128	-0.0159	-0.0222	-0.0222	-0.0592	-0.0120	-0.0120	-0.0118
	(0.0107) (0.0097)	(0.0119)	(0.0303)	(0.0140)	(0.0314) <sup>c</sup>	(0.0122)	(0.0124)	(0.0136)
Retail	-0.0298 -0.0298	-0.0424	-0.0139	-0.0139	-0.0377	-0.0280	-0.0280	-0.0303
	(0.0122) <sup>b</sup> (0.0104) <sup>s</sup>	a (0.0135) <sup>a</sup>	(0.0278)	(0.0186)	(0.0292)	(0.0142) <sup>t</sup>	' (0.0100) <sup>a</sup>	(0.0158) <sup>c</sup>
Business & repair services	0.0295 0.0295	0.0252	0.0458	0.0458	0.0387	0.0237	0.0237	0.0270
	(0.0213) (0.0168) <sup>6</sup>	(0.0235)	(0.0604)	(0.0461)	(0.0643)	(0.0223)	(0.0190)	(0.0249)
Educational services	-0.0018 -0.0018 (0.0221) (0.0171)	-0.0231 (0.0243)		-0.1198 (0.0287) <sup>a</sup>	-0.0830 (0.0599)	0.0461 (0.0239) <sup>0</sup>	0.0461 (0.0171) <sup>ª</sup>	0.0063 (0.0264)
Professional related services	-0.0217 -0.0217	-0.0276	0.0655	0.0655	0.0073	-0.0631	-0.0631	-0.0486
	(0.0223) (0.0137)	(0.0247)	(0.0570)	(0.0265) <sup>a</sup>	(0.0596)	(0.0245) <sup>8</sup>	' (0.0136)ª	(0.0273) <sup>c</sup>
Health services	0.0118 0.0118 (0.0225) (0.0153)	0.0043 (0.0248)	-0.0630 (0.0591)	-0.0630 (0.0294) <sup>b</sup>		0.0486 (0.0241) <sup>t</sup>	0.0486 (0.0163) <sup>*</sup>	0.0266 (0.0268)
Personal services	0.0232 0.0232	0.0256	0.0621	0.0621	0.0794	0.0118	0.0118	0.0184
	(0.0130)° (0.0091) <sup>t</sup>	° (0.0144) <sup>°</sup>	(0.0360) <sup>6</sup>	(0.0114) <sup>a</sup>	(0.0382) <sup>b</sup>	(0.0143)	(0.0089)	(0.0160)
Entertainment & recreational services	0.1040 0.1040	0.1121	0.1554	0.1554	0.0936	0.0760	0.0760	0.1027
	(0.0409) <sup>b</sup> (0.0244) <sup>s</sup>	(0.0452) <sup>b</sup>	(0.1082)	(0.0559) <sup>*</sup>	(0.1146)	(0.0436) <sup>6</sup>	(0.0236) <sup>*</sup>	(0.0485) <sup>b</sup>
Transportation	-0.0159 -0.0159 (0.0132) (0.0104)			0.0019 (0.0112)			-0.0275 (0.0146) <sup>°</sup>	
Wholesale trade	0.0484 0.0484	0.0390	0.0802	0.0802	0.0656	0.0364	0.0364	0.0288
	(0.0169) <sup>a</sup> (0.0113) <sup>a</sup>	a (0.0186) <sup>b</sup>	(0.0436) <sup>6</sup>	(0.0187) <sup>a</sup>	(0.0464)	(0.0183) <sup>t</sup>	(0.0139)ª	(0.0204)
Poverty	-0.0157 -0.0157 (0.0042) <sup>a</sup> (0.0057) <sup>a</sup>		-0.0183 (0.0120)	-0.0183 (0.0095) <sup>°</sup>	-0.0459 (0.1155) <sup>a</sup>		-0.0137 ' (0.0053) <sup>*</sup>	-0.0416 (0.0042) <sup>a</sup>
Metro area, 1970	-0.0006 -0.0006 (0.0006) (0.0005)	-0.0007 (0.0007)	-0.0013 (0.0011)	-0.0013 (0.0005) <sup>*</sup>			-0.0017 (0.0004) <sup>a</sup>	
R <sup>2</sup>	0.33 0.95	0.77	0.37	0.95	0.75	0.39	0.95	0.75
# Observations	1009 1009	1009	252	252	252	757	757	757

TABLE A4.—	-GROWTH EQUATION	ESTIMATES -	SOUTH	(CONTINUED)
------------	------------------	-------------	-------	-------------

 $<sup>^3</sup>$  All RHS variables are initial values from 1970. Also, state dummies were employed in all regressions. 13

TABLE A5GI	ROWTH EQUATION	ESTIMATES - F	PLAINS REGION
------------	----------------	---------------	---------------

	AII	Metro	Non-Metro
RHS Variables <sup>1</sup>	OLS CR OLS 3SLS	OLS CR OLS 3SLS	OLS CR OLS 3SLS
Constant	$0.1685  0.1685  0.2475 \\ (0.0288)^{a} \ (0.0229)^{a} \ (0.0262)^{a}$	$\begin{array}{cccc} 0.0473 & 0.0473 & 0.1472 \\ (0.0681) & (0.0330) & (0.0730)^{\mathrm{b}} \end{array}$	$0.1736$ $0.1736$ $0.2918$ $(0.0318)^{a}$ $(0.0294)^{a}$ $(0.0295)^{a}$
Log 1970 per capita income <sup>2</sup>	$\begin{array}{rrr} -0.0179 & -0.0179 & -0.0266 \\ (0.0015)^{a} & (0.0018)^{a} & (0.0046)^{a} \end{array}$	$\begin{array}{r} \text{-0.0084} & \text{-0.0084} & \text{-0.0271} \\ (0.0037)^{\text{b}} & (0.0023)^{\text{a}} & (0.0030)^{\text{a}} \end{array}$	$\begin{array}{rrr} -0.0182 & -0.0182 & -0.0308 \\ (0.0016)^{a} & (0.0019)^{a} & (0.0048)^{a} \end{array}$
Land area per capita	-0.0013 -0.0013 -0.0013 (0.0002) <sup>a</sup> (0.0003) <sup>a</sup> (0.0002)	-0.0014 -0.0014 -0.0014 (0.0010) (0.0006) <sup>b</sup> (0.0012)	$\begin{array}{ccc} -0.0013 & -0.0013 & -0.0012 \\ (0.0002)^{a} & (0.0003)^{a} & (0.0003)^{a} \end{array}$
Water area per capita	0.0032 0.0032 0.0021 (0.0037) (0.0021) (0.0038)	$\begin{array}{cccc} 0.1273 & 0.1273 & 0.0918 \\ (0.0755)^{\circ} & (0.0215)^{a} & (0.0843) \end{array}$	$\begin{array}{rrrr} 0.0049 & 0.0049 & 0.0031 \\ (0.0038) & (0.0024)^{\text{b}} & (0.0040) \end{array}$
Age: 5-13 years	0.0174 0.0174 0.0252 (0.0317) (0.0378) (0.0323)	$\begin{array}{rrrr} 0.0733 & 0.0733 & 0.1695 \\ (0.0858) & (0.0416)^\circ & (0.0938)^\circ \end{array}$	-0.0027 -0.0027 0.0006 (0.0342) (0.0391) (0.0358)
Age: 14-17 years	$\begin{array}{llllllllllllllllllllllllllllllllllll$	0.0314 0.0314 0.0424 (0.0731) (0.0474) (0.0819)	0.0347 0.0347 0.0324 (0.0338) (0.0254) (0.0354)
Age: 18-64 years	0.0147 0.0147 0.0177 (0.0243) (0.0262) (0.0248)	$\begin{array}{cccc} 0.0508 & 0.0508 & 0.1197 \\ (0.0668) & (0.0210)^{\mathrm{b}} & (0.0734) \end{array}$	-0.0023 -0.0023 -0.0045 (0.0266) (0.0275) (0.0278)
Age: 65+	-0.0073 -0.0073 -0.0148 (0.0225) (0.0257) (0.0230)	$\begin{array}{cccc} 0.0168 & 0.0168 & 0.0852 \\ (0.0581) & (0.0275)^{a} & (0.0634) \end{array}$	-0.0171 -0.0171 -0.0354 (0.0249) (0.0276) (0.0259)
Blacks	-0.0009 -0.0009 0.0007 (0.0103) (0.0112) (0.0105)	-0.0193 -0.0193 -0.0284 (0.0160) (0.0064) <sup>a</sup> (0.0179)	0.0068 0.0068 0.0112 (0.0128) (0.0112) (0.0134)
Hispanic	$\begin{array}{rrrr} -0.0060 & -0.0060 & -0.0080 \\ (0.0042) & (0.0027) & (0.0043)^{\rm c} \end{array}$	-0.0000 -0.0000 -0.0008 (0.0087) (0.0034) (0.0098)	$\begin{array}{rrrr} \text{-0.0042} & \text{-0.0042} & \text{-0.0086} \\ (0.0050) & (0.0033) & (0.0052)^c \end{array}$
Education: 9-11 years	$\begin{array}{rrr} \text{-0.0561} & \text{-0.0561} & \text{-0.0542} \\ \text{(0.0084)}^{\text{a}} & \text{(0.0068)} & \text{(0.0087)}^{\text{a}} \end{array}$	$\begin{array}{l} \text{-0.0459}  \text{-0.0459}  \text{-0.0490} \\ \left( 0.0186 \right)^{\text{b}} \left( 0.0077 \right)^{\text{a}} \left( 0.0209 \right)^{\text{b}} \end{array}$	$\begin{array}{rrr} \text{-0.0573} & \text{-0.0573} & \text{-0.0553} \\ (0.0096)^{a} & (0.0068)^{a} & (0.0100)^{a} \end{array}$
Education: H.S. diploma	$\begin{array}{llllllllllllllllllllllllllllllllllll$	$\begin{array}{cccc} 0.0297 & 0.0297 & 0.0035 \\ (0.0124)^{b} & (0.0065)^{a} & (0.0139)^{b} \end{array}$	$\begin{array}{cccc} 0.0060 & 0.0060 & 0.0166 \\ (0.0060) & (0.0057) & (0.0061)^a \end{array}$
Education: Some college	$\begin{array}{rrr} -0.0347 & -0.0347 & -0.0316 \\ (0.0091)^{a} & (0.0076)^{a} & (0.0093)^{a} \end{array}$	0.0001 0.0001 -0.0095 (0.0212) (0.0226) (0.0237)	$\begin{array}{r} \text{-0.0307} & \text{-0.0307} & \text{-0.0246} \\ (0.0101)^{\text{a}} & (0.0080)^{\text{a}} & (0.0106)^{\text{b}} \end{array}$
Education: Bachelor +	0.0090 0.0090 0.0194 (0.0119) (0.0103) (0.0120)	$\begin{array}{ccc} 0.0468 & 0.0468 & 0.0841 \\ (0.0238)^{\circ} & (0.0140)^{a} & (0.0254)^{a} \end{array}$	-0.0096 -0.0096 0.0062 (0.0138) (0.0115) (0.0143)
Education: Public elementary	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)	$\begin{array}{llllllllllllllllllllllllllllllllllll$	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)
Education: Public nursery	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)
Education: Private elementary	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)
Education: Private nursery	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)
Housing	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	$\begin{array}{cccc} 0.0000 & 0.0000 & 0.0000 \\ (0.0000)^{b} & (0.0000) & (0.0000)^{b} \end{array}$	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)
Federal government employment	$\begin{array}{rrr} \text{-0.0201} & \text{-0.0201} & \text{-0.0261} \\ \text{(0.0102)}^{\text{b}} & \text{(0.0101)}^{\text{b}} & \text{(0.0104)}^{\text{b}} \end{array}$	-0.0201 -0.0201 -0.0139 (0.0253) (0.0077) <sup>a</sup> (0.0284)	$\begin{array}{rrrr} \text{-0.0182} & \text{-0.0182} & \text{-0.0264} \\ \text{(0.0114)} & \text{(0.0120)} & \text{(0.0119)}^{\text{b}} \end{array}$
State government employment	$\begin{array}{cccc} 0.0132 & 0.0132 & 0.0064 \\ (0.0077)^{\circ} & (0.0075)^{\circ} & (0.0077) \end{array}$	$\begin{array}{rrrr} \text{-0.0179} & \text{-0.0179} & \text{-0.0357} \\ \text{(0.0141)} & \text{(0.0145)} & \text{(0.0153)}^{\text{b}} \end{array}$	$\begin{array}{cccc} 0.0243 & 0.0243 & 0.0161 \\ (0.0088)^{a} & (0.0083)^{a} & (0.0092)^{c} \end{array}$
Local government employment	0.0001 0.0001 0.0048 (0.0085) (0.0115) (0.0086)	-0.0157 -0.0157 -0.0067 (0.0200) (0.0162) (0.0224)	$\begin{array}{cccc} 0.0101 & 0.0101 & 0.0188 \\ (0.0098) & (0.0125) & (0.0102)^{\circ} \end{array}$

	All			Metro	Non-Metro	_
RHS Variables <sup>3</sup>	OLS CR OLS	<u>3SLS</u>	<u>OLS</u>	CR OLS 3SLS	OLS CR OLS 3	<u>ISLS</u>
Self-employment	0.0115 0.0115 (0.0060) <sup>c</sup> (0.0085	-0.0078 ) (0.0061)	-0.0060 (0.0150)	-0.0060 -0.0201 (0.0079) (0.0165)	$\begin{array}{cccc} 0.0148 & 0.0148 & 0.0148 & 0.0148 & 0.0065 \end{array}$	0.0102 0.0068)
Agriculture	-0.0030 -0.0030 (0.0125) (0.0075		0.0076 (0.0283)	0.0076 0.0388 (0.0108) (0.0309)		0.0148 0.0155)
Communications	-0.0045 -0.0045 (0.0198) (0.0117		0.0122 (0.0543)	0.0122 0.0424 (0.0334) (0.0605)		0.0224 0.0234)
Construction	0.0002 0.0002 (0.0148) (0.0106	-0.0051 ) (0.0151)	0.0303 (0.0313)	0.0303 0.0348 (0.0288) (0.0351)		) .0042 0.0185)
Finance, insurance & real estate	0.0919 0.0919 (0.0228) <sup>a</sup> (0.0271	0.0944 ) <sup>a</sup> (0.0233) <sup>a</sup>	0.0022 (0.0460)	0.0022 0.0279 (0.0207) (0.0512)	0.1126 $0.1126$ (0.0264) <sup>a</sup> (0.0298) <sup>a</sup>	0.1174 0.0276)ª
Manufacturing – durables	0.0156 0.0156 (0.0118) (0.0097	0.0108 ) (0.0121)	0.0138 (0.0268)	0.0138 0.0267 (0.0088) (0.0299)	$\begin{array}{cccc} 0.0224 & 0.0224 & 0.0224 & 0.0224 & 0.0224 & 0.00264 & 0.00086)^a \end{array}$	0.0185 0.0146)
Manufacturing – nondurables	0.0010 0.0010 (0.0129) (0.0088	-0.0028 ) 0.0132)	-0.0142 (0.0277)	-0.0142 -0.0006 (0.0101) (0.0309)	••••••	0.0113 0.0162)
Mining	0.0030 0.0030 (0.0124) (0.0087	-0.0036 ) (0.0127)	0.0256 (0.0331)	0.0256 0.0278 (0.0159) (0.0371)		0.0073 0.0155)
Retail	0.0001 0.0001 (0.0132) (0.0068	-0.0031 ) (0.0135)	-0.0334 (0.0269)	-0.0334 -0.0405 (0.0093) <sup>a</sup> (0.0301)	$\begin{array}{cccc} 0.0164 & 0.0164 & (0.0156) \\ (0.0156) & (0.0086)^{\circ} & (0.0086)^{\circ} \end{array}$	0.0151 0.0164)
Business & repair services	-0.0341 -0.0341 (0.0214) (0.0183	-0.0353 ) (0.0219)	0.1676 (0.0575)	0.1676 0.1212 a (0.0330) <sup>a</sup> (0.0637) <sup>c</sup>	-0.0274 -0.0274 -( (0.0236) (0.0184) (0	0.0260 0.0247)
Educational services	-0.0353 -0.0353 (0.0187) <sup>b</sup> (0.0086		0.0085 (0.0516)	0.0085 -0.0676 (0.0241) (0.0553)	-0.0363 -0.0363 -( (0.0209)° (0.0111)ª ((	0.0394 0.0219)°
Professional related services	0.0276 0.0276 (0.0200) (0.0131	0.0216 ) <sup>b</sup> (0.0204)	-0.0117 (0.0521)	-0.0117 0.0550 (0.0215) (0.0565)	$\begin{array}{cccc} 0.0377 & 0.0377 & (0.0233) & (0.0164)^{b} \end{array}$	0.0261 0.0243)
Health services	-0.0225 -0.0225 (0.0188) (0.0104		0.0409 (0.0512)	0.0409 -0.0223 (0.0248) (0.0556)	-0.0256 -0.0256 -( (0.0210) (0.0131) <sup>°</sup> (	0.0092 0.0219)
Personal services	0.0770 0.0770 (0.0172) <sup>a</sup> (0.0143	0.0750 ) <sup>a</sup> (0.0176) <sup>a</sup>	0.0274 (0.0440)	$\begin{array}{ccc} 0.0274 & 0.0769 \\ (0.0130)^{b} & (0.0481) \end{array}$	0.1012 0.1012 ( (0.0199) <sup>a</sup> (0.0179) <sup>a</sup> (	0.0994 0.0208)ª
Entertainment & recreational services	0.1159 0.1159 (0.0381) <sup>a</sup> (0.0229	0.1151 ) <sup>a</sup> (0.0390) <sup>a</sup>	0.1574 (0.0801)	0.1574 0.2110 ° (0.0583) <sup>ª</sup> (0.0890) <sup>b</sup>	0.1330 0.1330 ( (0.0440) <sup>a</sup> (0.0276) <sup>a</sup> (	0.1293 0.0460)ª
Transportation	0.0118 0.0118 (0.0159) (0.0116		0.0241 (0.0318)	$\begin{array}{ccc} 0.0241 & 0.0343 \\ (0.0097)^{b} & (0.0356) \end{array}$	0.0152 0.0152 ( (0.0189) (0.0115) (0	
Wholesale trade	0.0528 0.0528 (0.0197)ª (0.0163	0.0541 ) <sup>a</sup> (0.0202) <sup>a</sup>	-0.0323 (0.0449)	-0.0323 0.0158 (0.0175) <sup>c</sup> (0.0492)	$\begin{array}{cccc} 0.0538 & 0.0538 & 0 \\ (0.0226)^{ ext{b}} & (0.0198)^{ ext{a}} \end{array}$	0.0572 0.0237) <sup>⊳</sup>
Poverty	-0.0276 -0.0276 (0.0055) <sup>a</sup> (0.0124	-0.0412 <sup>b</sup> (0.0051) <sup>a</sup>	0.0154 (0.0144)	0.0154 -0.0243 (0.0163) (0.0135) <sup>°</sup>	-0.0302 -0.0302 -( (0.0061) <sup>a</sup> (0.0141) ((	
Metro area, 1970	0.0017 0.0017 (0.0012) (0.0005		-0.0004 (0.0015)	-0.0004 -0.0007 (0.0008) (0.0017)	-0.0017 -0.0017 -( (0.0023) (0.0008) <sup>b</sup> ((	
R <sup>2</sup> # Observations	0.54 0.90 832 832	0.59 832	0.73 143	0.97 0.85 143 143		).59 689

#### TABLE A5.—GROWTH EQUATION ESTIMATES - PLAINS REGION (CONTINUED)

 $<sup>\</sup>frac{1}{3}$  All RHS variables are initial values from 1970. Also, state dummies were employed in all regressions.

|--|

	AII	Metro	Non-Metro
RHS Variables <sup>1</sup>	OLS CR OLS 3SLS	OLS CR OLS 3SLS	OLS CR OLS 3SLS
Constant	$\begin{array}{cccc} 0.0207 & 0.0207 & 0.3553 \\ (0.0398)^{a} & (0.0276) & (0.0380)^{a} \end{array}$	$\begin{array}{cccc} 0.0698 & 0.0698 & 0.4243 \\ (0.0777) & (0.0617) & (0.0776)^{a} \end{array}$	$0.1938  0.1938  0.3057 \\ (0.0523)^{a} \ (0.0292)^{a} \ (0.0512)^{a}$
Log 1970 per capita income <sup>2</sup>	$\begin{array}{rrr} \text{-0.0210} & \text{-0.0210} & \text{-0.0359} \\ (0.0018)^{a} & (0.0019) & (0.0020)^{a} \end{array}$	$-0.0108$ $-0.0108$ $-0.0391$ $(0.0033)^{a}$ $(0.0026)^{a}$ $(0.0019)^{a}$	$\begin{array}{r} -0.0224 & -0.0224 & -0.0356 \\ (0.0024)^{a} & (0.0018)^{a} & (0.0022)^{a} \end{array}$
Land area per capita	$\begin{array}{c} -0.0008 & -0.0008 & -0.0008 \\ (0.0002)^a & (0.0003) & (0.0003)^a \end{array}$	$\begin{array}{rrr} -0.0006 & -0.0006 & -0.0003 \\ (0.0004) & (0.0003)^{\rm b} & (0.0005) \end{array}$	$\begin{array}{ccc} -0.0010 & -0.0010 & -0.0009 \\ (0.0003)^{a} & (0.0004)^{b} & (0.0004)^{a} \end{array}$
Water area per capita	$\begin{array}{llllllllllllllllllllllllllllllllllll$	$\begin{array}{cccc} 0.0069 & 0.0069 & 0.0097 \\ (0.0052) & (0.0032)^{\text{b}} & (0.0061) \end{array}$	$\begin{array}{cccc} 0.0095 & 0.0095 & 0.0092 \\ (0.0031)^{a} & (0.0036)^{b} & (0.0033)^{a} \end{array}$
Age: 5-13 years	-0.0079 -0.0079 -0.0077 (0.0462) (0.0343) (0.0494)	0.0410 0.0410 -0.0477 (0.0932) (0.0532) (0.1089)	0.0025 0.0025 0.0248 (0.0601) (0.0471) (0.0636)
Age: 14-17 years	0.0069 0.0069 0.0113 (0.0412) (0.0315) (0.0440)	0.0815 0.0815 -0.0109 (0.0711) (0.0555) (0.0826)	-0.0237 -0.0237 0.0089 (0.0597) (0.0435) (0.0629)
Age: 18-64 years	0.0084 0.0084 0.0015 (0.0352) (0.0260) (0.0376)	0.0142 0.0142 -0.0706 (0.0693) (0.0497) (0.0807)	0.0383 0.0383 0.0497 (0.0478) (0.0310) (0.0506)
Age: 65+	-0.0030 -0.0030 -0.0202 (0.0311) (0.0268) (0.0331)	0.0324 0.0324 -0.0641 (0.0611) (0.0419) (0.0707)	0.0070 0.0070 0.0043 (0.0411) (0.0359) (0.0435)
Blacks	0.0053 0.0053 0.0064 (0.0044) (0.0028) (0.0047)	$\begin{array}{cccc} 0.0070 & 0.0070 & 0.0121 \\ (0.0063) & (0.0031)^{\rm b} & (0.0074) \end{array}$	0.0052 0.0052 0.0053 (0.0065) (0.0035) (0.0069)
Hispanic	-0.0038 -0.0038 -0.0052 (0.0028) (0.0030) (0.0030) <sup>c</sup>	-0.0006 -0.0006 0.0019 (0.0051) (0.0032) (0.0060)	-0.0044 -0.0044 -0.0056 (0.0038) (0.0036) (0.0040)
Education: 9-11 years	-0.0233 -0.0233 -0.0347 (0.0087) <sup>a</sup> (0.0118) (0.0092) <sup>a</sup>	$-0.0365$ $-0.0365$ $-0.0339$ $(0.0156)^{b}$ $(0.0162)^{b}$ $(0.0184)^{c}$	$\begin{array}{rrr} \text{-0.0177} & \text{-0.0177} & \text{-0.0285} \\ \text{(0.0110)} & \text{(0.0090)}^{\text{b}} & \text{(0.0114)}^{\text{b}} \end{array}$
Education: H.S. diploma	0.0116 0.0116 0.0228 (0.0086) (0.0143) (0.0091) <sup>b</sup>	-0.0043 -0.0043 0.0065 (0.0138) (0.0177) (0.0162)	$\begin{array}{ccc} 0.0268 & 0.0268 & 0.0381 \\ (0.0117)^{\mathrm{b}} & (0.0127)^{\mathrm{b}} & (0.0122)^{\mathrm{a}} \end{array}$
Education: Some college	$\begin{array}{cccc} 0.0267 & 0.0267 & 0.0345 \\ (0.0145)^{\circ} & (0.0105) & (0.0155)^{\mathrm{b}} \end{array}$	-0.0238 -0.0238 0.0081 (0.0249) (0.0201) (0.0290)	$0.0449  ext{ 0.0449  ext{ 0.0527}} (0.0191)^{ ext{b}} (0.0147)^{ ext{a}} (0.0201)^{ ext{a}}$
Education: Bachelor +	0.0224 0.0224 0.0419 (0.0140) (0.0268) (0.0148) <sup>a</sup>	0.0385 0.0385 0.0855 (0.0237) (0.0253) (0.0271)ª	0.0021 0.0021 0.0284 (0.0203) (0.0262) (0.0210)
Education: Public elementary	$\begin{array}{rrrr} -0.0000 & -0.0000 & -0.0000 \\ (0.0000) & (0.0000) & (0.0000)^{b} \end{array}$	$\begin{array}{rrrr} -0.0000 & -0.0000 & -0.0000 \\ (0.0000) & (0.0000) & (0.0000)^{\circ} \end{array}$	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)
Education: Public nursery	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)
Education: Private elementary	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)	-0.0000 -0.0000 -0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)
Education: Private nursery	$\begin{array}{cccc} 0.0000 & 0.0000 & 0.0000 \\ (0.0000)^{\rm b} & (0.0000) & (0.0000)^{\rm a} \end{array}$	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 -0.0000 (0.0000) (0.0000) 0.0000
Housing	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)	0.0000 0.0000 0.0000 (0.0000) (0.0000) (0.0000)
Federal government employment	-0.0134 -0.0134 -0.0167 (0.0106) (0.0081) (0.0113)	0.0003 0.0003 -0.0108 (0.0191) (0.0161) (0.0224)	-0.0082 -0.0082 -0.0104 (0.0141) (0.0083) (0.0150)
State government employment	-0.0011 -0.0011 -0.0149 (0.0097) (0.0087) (0.0102)	-0.0041 -0.0041 -0.0207 (0.0138) (0.0113) (0.0160)	0.0124 0.0124 -0.0038 (0.0148) (0.0107) (0.0154)
Local government employment	$\begin{array}{rrr} \text{-0.0243} & \text{-0.0243} & \text{-0.0294} \\ \text{(0.0116)}^{\text{b}} & \text{(0.0081)} & \text{(0.0124)}^{\text{b}} \end{array}$	$\begin{array}{rrrr} \text{-0.0115} & \text{-0.0115} & \text{-0.0398} \\ (0.0194) & (0.0200) & (0.0225)^{\circ} \end{array}$	-0.0264 -0.0264 -0.0271 (0.0164) (0.0095) <sup>a</sup> (0.0174)

	A	.II		Metro		N	on-Metro_	
RHS Variables <sup>3</sup>	OLS CR	OLS 3SLS	OLS	CR OLS	<u>3SLS</u>	<u>OLS</u>	CR OLS	<u>3SLS</u>
Self-employment		070 0.0100 059) (0.0089)	0.0103 (0.0152)	0.0103 (0.0113)	0.0135 (0.0179)	0.0007 (0.0115)	0.0007 (0.0099)	0.0039 (0.0122)
Agriculture		0090 -0.0034 110) (0.0146)	0.0100 (0.0219)	0.0100 (0.0152)	0.0242 (0.0257)	0.0029 (0.0185)	0.0029 (0.0128)	0.0063 (0.0195)
Communications		9434 -0.0537 9263)° (0.0262) <sup>b</sup>	0.0312 (0.0406)	0.0312 (0.0305)	0.0118 (0.0476)	-0.0583 (0.0322) <sup>c</sup>	-0.0583 (0.0190) <sup>a</sup>	-0.0727 (0.0340) <sup>b</sup>
Construction		139 0.0153 138) (0.0171)	0.0693 (0.0259)	0.0693 a (0.0156)ª	0.0766 <sup>a</sup> (0.0305) <sup>b</sup>	0.0050 (0.0212)	0.0050 (0.0140)	0.0065 (0.0224)
Finance, insurance & real estate		996 0.1230 258) <sup>a</sup> (0.0327) <sup>a</sup>	0.1633 (0.0481)	0.1633 (0.0275)	0.1678 <sup>a</sup> (0.0566) <sup>a</sup>	0.0454 (0.0436)	0.0454 (0.0383)	0.0802 (0.0458)
Manufacturing – durables		092 0.0039 087) (0.0142)	0.0384 (0.0212)	0.0384 ° (0.0147) <sup>°</sup>	0.0192 a (0.0248)	0.0187 (0.0183)	0.0187 (0.0112) <sup>°</sup>	0.0165 (0.0194)
Manufacturing – nondurables		039 -0.0081 123) (0.0154)	0.0083 (0.0226)	0.0083 (0.0199)	-0.0015 (0.0266)	0.0038 (0.0204)	0.0038 (0.0149)	0.0016 (0.0216)
Mining		077 -0.0088 102) (0.0147)	0.0227 (0.0230)	0.0277 (0.0167) <sup>°</sup>	0.0029 (0.0269)	-0.0000 (0.0186)	-0.0000 (0.0108)	-0.0003 (0.0197)
Retail		294 -0.0393 108) <sup>a</sup> (0.0169) <sup>b</sup>	-0.0200 (0.0277)	-0.0200 (0.0276)	-0.0344 (0.0325)	-0.0167 (0.0214)	-0.0167 (0.0125)	-0.0241 (0.0226)
Business & repair services		045 -0.0032 193) (0.0267)	-0.0212 (0.0414)	-0.0212 (0.0283)	-0.0321 (0.0487)	0.0371 (0.0376)	0.0371 (0.0247)	0.0494 (0.0398)
Educational services		1361 -0.0542 1061) <sup>a</sup> (0.0149) <sup>a</sup>	0.0053 (0.0189)	0.0053 (0.0102)	-0.0563 (0.0206) <sup>a</sup>	-0.0242 (0.0348)	-0.0242 (0.0193)	-0.0404 (0.0368)
Professional related services		101 0.0175 088) (0.0168)	0.0190 (0.0213)	0.0190 (0.0124)	0.0253 (0.0251)	-0.0066 (0.0344)	-0.0066 (0.0235)	0.0066 (0.0364)
Health services		033 0.0007 181) (0.0188)	0.0244 (0.0234)	0.0244 (0.0249)	0.0265 (0.0275)	0.0232 (0.0358)	0.0232 (0.0279)	0.0127 (0.0379)
Personal services		056 0.0114 122) (0.0178)	0.0264 (0.0253)	0.0264 (0.0150)°	0.0249 (0.0298)	0.0297 (0.0252)	0.0297 (0.0163) <sup>c</sup>	0.0386 (0.0266)
Entertainment & recreational services	-0.0155 -0.0 (0.0243) (0.0	155 -0.0126 151) (0.0260)	0.0318 (0.0454)	0.0318 (0.0190) <sup>°</sup>	0.0301 (0.0533)	-0.0079 (0.0308)	-0.0079 (0.0157)	-0.0054 (0.0326)
Transportation		063 -0.0041 173) (0.0206)	0.0466 (0.0319)	0.0466 (0.0401)	0.0545 (0.0375)	-0.0082 (0.0254)	-0.0082 (0.0184)	-0.0074 (0.0269)
Wholesale trade	0.0058 0.0 (0.0217) (0.0	058 0.0071 126) (0.0232)	0.0278 (0.0294)	0.0278 (0.0164) <sup>c</sup>	0.0049 (0.0344)	-0.0064 (0.0359)	-0.0064 (0.0237)	0.0141 (0.0378)
Poverty	-0.0151 -0.0 (0.0070) <sup>b</sup> (0.0	151 -0.0375 120) (0.0067) <sup>a</sup>	-0.0261 (0.0113)	-0.0261 <sup>b</sup> (0.0181)	-0.0758 (0.0115) <sup>ª</sup>	-0.0051 (0.0094)	-0.0051 (0.0139)	-0.0218 (0.0094) <sup>b</sup>
College Town		024 0.0028 0007) <sup>a</sup> (0.0014) <sup>c</sup>	0.0021 (0.0014)	0.0021 (0.0009) <sup>t</sup>	0.0037 ' (0.0017) <sup>b</sup>	-0.0008 (0.0040)	-0.0008 (0.0200)	-0.0006 (0.0042)
Metro area, 1970	0.0025 0.0 (0.0010) <sup>a</sup> (0.0	025 0.0024 0009) <sup>a</sup> (0.0010) <sup>b</sup>	0.0007 (0.0011)	0.0007 (0.0012)	0.0005 (0.0013)	0.1938 (0.0523) <sup>8</sup>	0.1938 ' (0.0009) <sup>ª</sup>	0.0033 (0.0023)
R <sup>2</sup> # Observations	0.52 0.8 538 538		0.60 242	0.91 242	0.84 242	0.60 296	0.87 296	0.67 296

### TABLE A6.—GROWTH EQUATION ESTIMATES – WESTERN REGION (CONTINUED)

 $<sup>^3</sup>$  All RHS variables are initial values from 1970. Also, state dummies were employed in all regressions. \$17

## Bar-Ilan University Department of Economics WORKING PAPERS

1-01	<b>The Optimal Size for a Minority</b> Hillel Rapoport and Avi Weiss, January 2001.
2-01	An Application of a Switching Regimes Regression to the Study of Urban Structure Gershon Alperovich and Joseph Deutsch, January 2001.
3-01	The Kuznets Curve and the Impact of Various Income Sources on the Link Between Inequality and Development Joseph Deutsch and Jacques Silber, February 2001.
4-01	<b>International Asset Allocation: A New Perspective</b> Abraham Lioui and Patrice Poncet, February 2001.
5-01	מודל המועדון והקהילה החרדית יעקב רוזנברג, פברואר 2001.
6-01	<b>Multi-Generation Model of Immigrant Earnings: Theory and Application</b> Gil S. Epstein and Tikva Lecker, February 2001.
7-01	Shattered Rails, Ruined Credit: Financial Fragility and Railroad Operations in the Great Depression
8-01	<ul><li>Daniel A.Schiffman, February 2001.</li><li>Cooperation and Competition in a Duopoly R&amp;D Market</li><li>Damiano Bruno Silipo and Avi Weiss, March 2001.</li></ul>
9-01	A Theory of Immigration Amnesties Gil S. Epstein and Avi Weiss, April 2001.
10-01	<b>Dynamic Asset Pricing With Non-Redundant Forwards</b> Abraham Lioui and Patrice Poncet, May 2001.
	Electronic versions of the papers are available at http://www.biu.ac.il/soc/ec/wp/working_papers.html

- 11-01 Macroeconomic and Labor Market Impact of Russian Immigration in Israel
   Sarit Cohen and Chang-Tai Hsieh, May 2001.
- 12-01 Network Topology and the Efficiency of Equilibrium Igal Milchtaich, June 2001.
- 13-01 General Equilibrium Pricing of Trading Strategy Risk Abraham Lioui and Patrice Poncet, July 2001.
- 14-01 **Social Conformity and Child Labor** Shirit Katav-Herz, July 2001.
- 15-01 **Determinants of Railroad Capital Structure, 1830–1885** Daniel A. Schiffman, July 2001.
- 16-01 Political-Legal Institutions and the Railroad Financing Mix, 1885–1929Daniel A. Schiffman, September 2001.
- 17-01 Macroeconomic Instability, Migration, and the Option Value of Education

Eliakim Katz and Hillel Rapoport, October 2001.

- 18-01 Property Rights, Theft, and Efficiency: The Biblical Waiver of Fines in the Case of Confessed Theft Eliakim Katz and Jacob Rosenberg, November 2001.
- 19-01 **Ethnic Discrimination and the Migration of Skilled Labor** Frédéric Docquier and Hillel Rapoport, December 2001.
- 1-02 Can Vocational Education Improve the Wages of Minorities and Disadvantaged Groups? The Case of Israel
   Shoshana Neuman and Adrian Ziderman, February 2002.
- 2-02 What Can the Price Gap between Branded and Private Label Products Tell Us about Markups?
   Robert Barsky, Mark Bergen, Shantanu Dutta, and Daniel Levy, March 2002.
- 3-02 Holiday Price Rigidity and Cost of Price Adjustment Daniel Levy, Georg Müller, Shantanu Dutta, and Mark Bergen, March 2002.
- 4-02 **Computation of Completely Mixed Equilibrium Payoffs** Igal Milchtaich, March 2002.

- 5-02 Coordination and Critical Mass in a Network Market An Experimental Evaluation
   Amir Etziony and Avi Weiss, March 2002.
- 6-02 **Inviting Competition to Achieve Critical Mass** Amir Etziony and Avi Weiss, April 2002.
- 7-02 Credibility, Pre-Production and Inviting Competition in a Network Market Amir Etziony and Avi Weiss, April 2002.
- 8-02 **Brain Drain and LDCs' Growth: Winners and Losers** Michel Beine, Fréderic Docquier, and Hillel Rapoport, April 2002.
- 9-02 Heterogeneity in Price Rigidity: Evidence from a Case Study Using Micro-Level Data

Daniel Levy, Shantanu Dutta, and Mark Bergen, April 2002.

- 10-02 **Price Flexibility in Channels of Distribution: Evidence from Scanner Data** Shantanu Dutta, Mark Bergen, and Daniel Levy, April 2002.
- 11-02 Acquired Cooperation in Finite-Horizon Dynamic Games Igal Milchtaich and Avi Weiss, April 2002.
- 12-02 **Cointegration in Frequency Domain** Daniel Levy, May 2002.
- 13-02 Which Voting Rules Elicit Informative Voting? Ruth Ben-Yashar and Igal Milchtaich, May 2002.
- 14-02 Fertility, Non-Altruism and Economic Growth: Industrialization in the Nineteenth Century Elise S. Brezis, October 2002.
- 15-02 Changes in the Recruitment and Education of the Power Elites in Twentieth Century Western Democracies Elise S. Brezis and François Crouzet, November 2002.
- 16-02 **On the Typical Spectral Shape of an Economic Variable** Daniel Levy and Hashem Dezhbakhsh, December 2002.
- 17-02 International Evidence on Output Fluctuation and Shock Persistence Daniel Levy and Hashem Dezhbakhsh, December 2002.

- 1-03 **Topological Conditions for Uniqueness of Equilibrium in Networks** Igal Milchtaich, March 2003.
- 2-03 Is the Feldstein-Horioka Puzzle Really a Puzzle? Daniel Levy, June 2003.
- 3-03 Growth and Convergence across the US: Evidence from County-Level Data

Matthew Higgins, Daniel Levy, and Andrew Young, June 2003.

- 4-03 **Economic Growth and Endogenous Intergenerational Altruism** Hillel Rapoport and Jean-Pierre Vidal, June 2003.
- 5-03 **Remittances and Inequality: A Dynamic Migration Model** Frédéric Docquier and Hillel Rapoport, June 2003.
- 6-03 Sigma Convergence Versus Beta Convergence: Evidence from U.S. County-Level Data

Andrew T. Young, Matthew J. Higgins, and Daniel Levy, September 2003.

7-03 Managerial and Customer Costs of Price Adjustment: Direct Evidence from Industrial Markets

Mark J. Zbaracki, Mark Ritson, Daniel Levy, Shantanu Dutta, and Mark Bergen, September 2003.

### 8-03 First and Second Best Voting Rules in Committees

Ruth Ben-Yashar and Igal Milchtaich, October 2003.

9-03 Shattering the Myth of Costless Price Changes: Emerging Perspectives on Dynamic Pricing

Mark Bergen, Shantanu Dutta, Daniel Levy, Mark Ritson, and Mark J. Zbaracki, November 2003.

1-04 Heterogeneity in Convergence Rates and Income Determination across U.S. States: Evidence from County-Level Data

Andrew T. Young, Matthew J. Higgins, and Daniel Levy, January 2004.

- 2-04 "The Real Thing:" Nominal Price Rigidity of the Nickel Coke, 1886-1959 Daniel Levy and Andrew T. Young, February 2004.
- 3-04 Network Effects and the Dynamics of Migration and Inequality: Theory and Evidence from Mexico

David Mckenzie and Hillel Rapoport, March 2004.

- 4-04 **Migration Selectivity and the Evolution of Spatial Inequality** Ravi Kanbur and Hillel Rapoport, March 2004.
- 5-04 Many Types of Human Capital and Many Roles in U.S. Growth: Evidence from County-Level Educational Attainment Data

Andrew T. Young, Daniel Levy and Matthew J. Higgins, March 2004.

6-04 When Little Things Mean a Lot: On the Inefficiency of Item Pricing Laws Mark Bergen, Daniel Levy, Sourav Ray, Paul H. Rubin and Benjamin Zeliger, May 2004.

### 7-04 **Comparative Statics of Altruism and Spite** Igal Milchtaich, June 2004.

8-04 Asymmetric Price Adjustment in the Small: An Implication of Rational Inattention

Daniel Levy, Haipeng (Allan) Chen, Sourav Ray and Mark Bergen, July 2004.

1-05 Private Label Price Rigidity during Holiday Periods

Georg Müller, Mark Bergen, Shantanu Dutta and Daniel Levy, March 2005.

### 2-05 Asymmetric Wholesale Pricing: Theory and Evidence

Sourav Ray, Haipeng (Allan) Chen, Mark Bergen and Daniel Levy, March 2005.

3-05 Beyond the Cost of Price Adjustment: Investments in Pricing Capital

Mark Zbaracki, Mark Bergen, Shantanu Dutta, Daniel Levy and Mark Ritson, May 2005.

### 4-05 Explicit Evidence on an Implicit Contract

Andrew T. Young and Daniel Levy, June 2005.

5-05 Abracadabra! Social Norms and Public Perceptions through Harry Potter's Looking Glasses

Avichai Snir and Daniel Levy, September 2005.

# 6-05 Growth and Convergence across the US: Evidence from County-Level Data (revised version)

Matthew J. Higgins, Daniel Levy, and Andrew T. Young,, September 2005.