# Price Points and Price Rigidity 

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#### Abstract

We study the link between price points and price rigidity, using two datasets containing over 100 million observations. We find that (i) 9 is the most frequently used price-ending for the penny, dime, dollar and ten-dollar digits, (ii) 9-ending prices are between $24 \%-73 \%$ less likely to change in comparison to non-9-ending prices, (iii) the average size of the price change is higher if it ends with 9 in comparison to non-9-ending prices, and (iv) the most common price changes are multiples of dimes, dollars, and ten-dollars. We conclude that price points might constitute a substantial source of retail price rigidity.


## I. Introduction

With the increased popularity of new Keynesian models, the understanding of the sources of nominal price rigidity has become even more important. ${ }^{1}$ One of the recent theories of price rigidity is price point theory, which Blinder, et al. (1998) list among the leading 12 theories of price rigidity. According to Blinder, et al. (p. 26), practitioners' "...belief in pricing points is part of the folklore of pricing..." Consistent with this observation, they offer evidence from interviews on the importance of price points. For example, in their study of 200 large U.S. firms, they find that $88 \%$ of the retail firms assign substantial importance to price points in their pricing decisions. ${ }^{2}$ Kashyap (1995), who was the first to explore the link between price points and price rigidity, observes that catalog prices tend to be "stuck" at certain ending prices. He concludes that the existing theories cannot explain his findings, and therefore offers price points theory.

As Blinder, et al. (1998) note, however, a major difficulty with price point theory is that not much is known about the actual importance of price points and their relationship to price rigidity. Price points, although of interest by themselves (e.g., Landsburg, 1995), will be particularly important for macroeconomics if they can be shown to contribute to price rigidity, across a wide range of products and retailers. The literature offers growing evidence on the use of price points, but there is a lack of direct evidence on the link between price points and price rigidity in the U.S. Indeed, the literature documenting a link between price points and price rigidity using the U.S. data is limited to Blinder, et al.'s and Kashyap's (1995) studies. Kashyap emphasizes the need for more direct evidence, stating that—"A study focusing on more goods ... would have much more power to determine the significance of price points".

We fill this gap in the literature by offering new evidence on the link between price points and price rigidity using two datasets. One is a large weekly scanner price dataset, covering 29 product categories over an eight-year period from a major Midwestern U.S. retailer. The second dataset comes from the Internet and includes daily prices over a two-year period for 474 consumer electronic goods with a wide range of prices, such as music CDs, digital cameras, notebook PCs, etc., from 293 different e-retailers. Taken together, the two datasets cover a diverse set of products, a wide range of prices, different retail formats, retailers and time periods.

We find that across the two datasets, 9 is the most popular price point for the penny,

[^1]dime, dollar and the ten-dollar digits. We also find that the most common price changes are in multiples of dimes, dollars, and ten-dollar increments-an outcome that is consistent with efforts to keep the terminal digits at 9 . When we estimate the probability of a price change, we find that 9-ending prices are at least $24 \%$ (and as much as $73 \%$ ) less likely to change in comparison to non 9-ending prices. We also find that the average size of the 9-ending price changes are larger in comparison to non-9-ending prices, which underscores the extent of the 9-ending price rigidity.

We draw from the emerging macroeconomic literature on consumer inattention to explain these findings and argue how 9-ending prices can be the outcome of firms' reaction to consumers' inattention. ${ }^{3}$ Because consumers face large amounts of information, they may choose to be inattentive to the rightmost digits of retail prices. It is well-known that people process numeric information from left-to-right, and the processing of the rightmost digits in the price offers the least net benefit. Consequently time-pressed consumers may choose not to pay attention to them. In response, firms will set those digits to the highest possible number, 9. We conclude that price points may be a substantial source of retail price rigidity, and that consumer inattention may offer a plausible explanation for their use.

The paper is organized as follows. We describe the data in section II. In section III, we study the distribution of price-endings. In section IV, we study the distribution of the size of price changes. In section V, we estimate the effect of 9-endings on price rigidity. In section VI, we study the link between price points and the size of price changes. In section VII, we offer an explanation for the 9 -ending pricing practice. Section VIII concludes.

## II. Two Datasets

The most obvious prediction of Kashyap’s (1995) price point theory is that price points should be most important to retail firms (Blinder, et al 1998, Stahl 2009). We examine retail prices from two large datasets. One is Dominick's weekly price data for 29 product categories over an eight-year period. The other contains daily prices from the Internet on 474 products varying from music CDs, to DVDs, to hard disks, and to notebook PCs. The two datasets cover a wide variety of products, a wide price range, and different retail formats. In addition, although Dominick's is a grocery chain where prices are set on a chain-wide basis, our Internet data come from 293 different retailers presumably employing different pricing-decision models. Therefore, the conclusions we draw are not specific to a retail format, retailer, product, or price range.

Dominick's, a large supermarket chain in the Chicago metro area, operates about 100 stores with a market share of about $25 \%$. The data consist of up to 400 weekly observations of

[^2]retail prices in 29 different product categories, covering the period from September 14, 1989 to May 8, 1997. The prices are the actual transaction prices as recorded by the chain's scanners. If an item was on sale, then the price data reflect the sale prices. We use all the data from all stores, a total of over 98 million weekly price observations. The data contains a binary variable indicating whether a product was on sale. We use this variable in estimating the model. See Chevalier, et al. (2003) for more details about the data. Table 1 presents descriptive statistics.

The Internet data were obtained through the use of a price information gathering agent. It was programmed to download price data from www.BizRate.com, a popular price comparison site, from 3:00 a.m. to 5:00 a.m. From a list of products available at BizRate, we generated a large sample of product IDs using stratified proportionate random sampling (Wooldridge, 2002). The software agent then automatically built a panel of selling prices given the product IDs. ${ }^{4}$ The resulting dataset consists of 743 daily price observations for 474 personal electronic products in 10 product categories, from 293 different Internet-based retailers, over a period of more than two years from March 26, 2003 to April 15, 2005. The categories include music CDs, movie DVDs, video games, notebook PCs, personal digital assistants (PDAs), computer software, digital cameras and camcorders, DVD players, PC monitors, and hard drives. ${ }^{5}$ In total, the Internet data contains over 2.5 million daily price observations. Table 2 presents descriptive statistics.

## III. Evidence on the Popularity of 9-Ending Prices

"I asked the best economist I know, at least for such things-my wife, if she recalled a price not ending in a 9 at our local grocery store. Not really, she said. Maybe sometimes there are prices ending in a 5, but not really."

Jurek Konieczny (2003), Discussant Comments at the CEU Conference

We begin by presenting results on the frequency distribution of price-endings in the two datasets. In the analysis of Dominick's data, our focus is on $9 \mathbb{\$}$ and $99 \$$ price-endings because

[^3]the overwhelming majority of the prices in retail grocery stores are well below $\$ 10.00 .{ }^{6}$ In the Internet data, the prices range from $\$ 5.49$ to $\$ 6,478.00$, with the average prices in different categories spanning $\$ 12.90$ to $\$ 1,694.58$. In the Internet data, therefore, given the wider price range, we study not only $9 \$$ and $99 \$$ price-endings, but also other 9-ending prices in both the cents and the dollars digits, including \$9, \$9.99, \$99, and \$99.99.

In Figure 1, we report the frequency distribution of the last digit in Dominick's data. If a digit's appearance as a price-ending was random, then we should see $10 \%$ of the prices ending with each digit. As the figure indicates, however, about $65 \%$ of the prices end with 9 . The next most popular price-ending is 5 , accounting for about $11 \%$ of all price endings. Only a small proportion of the prices ends with the other digits. The pattern is very similar at the category level, with 9 as the most popular price-ending for all categories except cigarettes., 8

Next, we consider the frequency distribution of the last two digits. With two digits, there are 100 possible endings, $00 \Phi, 01 \Phi, \ldots, 98 \Phi$, and 99థ. Thus, with a random distribution, the probability of each ending is only $1 \%$. According to Figure 2, however, most prices end with either $09 \Phi, 19 \Phi, \ldots$, or 99 . This is not surprising since 9 is the dominant single-digit ending. But of these, more than $15 \%$ of the prices end with 994 . In contrast, only $4 \%$ to $6 \%$ of the prices end with $09 ¢, 19 \Phi, \ldots$, and $89 \Phi$, each. ${ }^{9}$ We found a similar pattern for individual categories. ${ }^{10}$

Figure 3 displays the frequency distribution of the last digit in the Internet data. 9 is the most popular terminal digit (33.4\%), followed by 0 (24.1\%), and 5 (17.4\%). The frequency distribution of the last two digits exhibits a similar pattern, with 99\$ as the most popular priceending (26.7\%), followed by 00\$ (20.3\%), 95\$ (13.8\%), and 98\$ (4.8\%). See Figure 4.

[^4]As mentioned above, the Internet dataset also includes some high-price product categories, which allows us to examine price-endings in dollar digits as well. In Figure 5, therefore, we present the frequency distribution of the last dollar digit in the Internet data. According to the figure, 9 is the most popular ending for the dollar digit, with $\$ 9$ price-endings over-represented with $36.1 \%$, followed by $\$ 4$ price-endings with $9.9 \%$, and $\$ 5$ price-endings with $9.2 \%$. A similar pattern emerges for the last two dollar digits as indicated by Figure 6. Not surprisingly, the last two dollar digits of most prices contain 9 , such as $\$ 99, \$ 89$, and $\$ 09$. But more prices end with $\$ 99$ than any other 9 price-endings. Moreover, almost $10 \%$ end with $\$ 99$ among the 100 possible dollar endings (i.e., $\$ 0$ through $\$ 99$ ).

We also examined the frequency distribution of the last three digits of prices in the Internet data. ${ }^{11}$ According to Table 3, \$9.99 is the most popular ending for the last three digits (13.2\%), followed by $\$ 9.00$ (10.0\%), and $\$ 9.95$ (4.9\%). ${ }^{12}$ When we examine the last four digits of the prices (last column of Table 4), $\$ 99.99$ is the most popular ending for the last four digits (3.47\%), followed by $\$ 99.00$ (3.46\%), and $\$ 19.99$ (2.16\%). ${ }^{13}$

In the Internet data, three individual product categories with low average prices exhibit some variation in price endings. ${ }^{14}$ For example, for the dollar-digit, $\$ 3, \$ 4$ and $\$ 5$ price-endings are the most common for CDs and DVDs because prices of CDs and DVDs are often between $\$ 13$ and $\$ 16$. Also, the $\$ 99$ and $\$ 99.99$ endings are not common in those two categories and the category of video games (see Table 4), because the average prices in these categories are far less than $\$ 100$ (i.e., $\$ 13.46$ for CDs, $\$ 27.43$ for DVDs, and $\$ 30.83$ for video games). It isn’t surprising, therefore, that we do not see a lot of 9-endings for the dollar and ten-dollar digits in those product categories.

To summarize, in both datasets, 9 is the most popular terminal digit overall. But the popularity of 9 is not limited to the penny digit. Rather, it is present in the dime, dollar, and tendollar digits too. The fact that our data include a variety of products with wide-ranging prices and different retail formats, further underscores the use of 9 as a terminal digit in our datasets.

## IV. Frequency Distribution of Price Changes by Size

Having documented the dominance of 9 as the terminal digit in both datasets, we next assess the extent to which the specific price point, 9 , that we have identified may be contributing

[^5]to the retail price rigidity. Figure 7 displays the frequency distribution of price changes in Dominick's data. Although the actual price changes occasionally go over \$1, these are few. We thus limit the analysis to price changes of up to $\$ 1$. According to the figure, the most common price changes, in fact, over $35 \%$ of the price changes are multiples of 10 cents. ${ }^{15}$ Consequently, the terminal digits are kept at 9 even after a price change. This indicates that terminal prices are "stuck" at 9.

In the Internet data, the observed price range is much wider and thus we observe a wider range for price changes. The price changes vary in magnitude from 14 to $\$ 1,568$, but the most common changes are in multiples of dollars and in multiples of dimes. As shown in Table 5, among the top ten most common changes, eight are multiples of dollars, and nine are multiples of dimes. The only exception is $1 \$$ which ranks tenth. Thus, similar to Dominick's dataset, the sizes of Internet price changes are such that they preserve the 9-endings.

Because of the wider range of price changes found in the Internet data, the ten most common price changes account for less than $30 \%$ of all price changes. As an alternative way to identify the prevalence of price changes in multiples of dimes, dollars, and tens of dollars, we categorize price changes based on how many digits in a price are affected by a price change (i.e., whether it affects the penny digit only, the penny and dime digits, or the penny, dime and dollar digits, etc.). For example, if we focus on price changes affecting the penny digit only, we can group all possible price changes into ten categories: those that change a price by $0 \Phi, 1 \Phi, \ldots, 9 \Phi$. In the first group will be price changes in multiples of dimes (excluding $0 \mathbb{\Phi}$ where a price does not change); in the second group, $1 母, 11$, ..., $91 \$, \$ 1.01, \ldots$. , etc. Similarly, we can group price changes into 100 groups based on how they affect the penny and the dime digits, one of which will be the category into which all price changes in multiples of dollars fall (again, excluding $0 \$$ where price does not change). Finally, we can group price changes into 1,000 groups based on how they affect the penny, dime and dollar-digits, one of which will be the category into which all price changes in tens of dollars fall (again excluding 0¢, where price does not change).

When we categorize price changes in this manner, we find that price changes in multiples of dimes are the most frequent among the ten possible changes to the penny digit, accounting for 55.12\% of all price changes. In addition, we find that among the 100 possible changes to the penny and dime digits, the most popular ones are multiples of dollars, which account for more than $42.86 \%$ of all changes. Finally, among the 1,000 possible changes to the last three digits, multiples of ten dollars are the most common, accounting for $9.60 \%$ of all changes. Similar

[^6]results are obtained for individual product categories. Changes in multiples of dimes and in multiples of dollars are the most common for all ten product categories in our dataset. Changes in multiples of ten dollars are the most common for seven product categories (video games, software, PDAs, DVD players, PC monitors, digital cameras, and notebook PCs). ${ }^{16}$ Based on the above results, we conclude that when prices change, they most often change in multiples of dimes, multiples of dollars, or in multiples of tens of dollars. Consequently, the terminal digits are kept at 9 even after a price change. This indicates that terminal prices are "stuck" at 9.

## V. The Effect of Price Points on Price Rigidity

To more directly study the link between 9-ending prices and price rigidity, we use a binomial logit model to estimate price change probabilities (Agresti, 2002; Hosmer and Lemeshow, 2000; Greene, 2003). Using the method of maximum likelihood, we estimate

$$
\begin{equation*}
\ln (q /(1-q))=a+b D_{9-\text { Ending }}+c D_{\text {Sale }}+e_{t}, \tag{1}
\end{equation*}
$$

where $q$ is the probability of a price change, $D_{9-E n d i n g}$ is a 9 -ending dummy variable which equals 1 if the price ends with 9 (i.e., $9 \$$-ending or $99 \Phi$-ending) and 0 otherwise, and $D_{\text {Sale }}$ is a sale dummy variable which equals 1 if the product is on sale and 0 otherwise. The regression equation includes the sale dummy because, according to Schindler (2006) and Anderson and Simester (2003), prices ending with 9 may be related to sales, and sale prices are more likely to change than regular prices. Indeed, if we consider a sample series of Frozen Concentrate Orange Juice, Heritage House, 12 oz. (UPC $=3828190029$ from Store No. 78), which is plotted in Figure 8 , it is clear that sale prices are always reversed, unless there is a change in the list price, which is rare. For example, in the sample of 400 observations shown in this figure, there are only about $14-16$ changes in the list price. By including the sale dummy, we account for any potential effect of sales when estimating price change probabilities.

The estimation results for Dominick's data are reported in Table 6. In the table, we report the estimated coefficients of each dummy along with the odds ratio that the coefficients imply. For all 27 product categories, the coefficient estimate on the 9-ending dummy is negative, and the coefficient estimate on the sale dummy is positive as expected (all p-values < 0.0001 ). The odds ratios, which equal $e^{\text {Coefficient }}$, are all smaller than 1 for the 9 -ending dummy, indicating that prices that end with $9 \mathbb{4}$ are less likely to change than prices that do not end with $9 \mathbb{4}$. On average, prices that end with $9 \mathbb{\$}$ are more than $40 \%$ less likely to change than prices that do not end with $9 \mathbb{4}$. Sale prices are about 65 times more likely to change than regular prices.

[^7]We obtain similar results for the 99¢-ending prices. The coefficient estimate on the 99\$ending dummy is negative and significant for all 27 categories, as shown in Table 6. The odds ratios indicate that prices that end with $99 \$$ are $24 \%$ less likely to change than prices that do not end with 994. Also, all product categories showed positive and significant coefficients on the sale dummy, and sale prices are about 67 times more likely to change than regular prices.

Next, we estimate the same logit regression model for the Internet data, but now we use $9 \$, 99 \Phi, \$ 9, \$ 9.99$, $\$ 99$, and $\$ 99.99$, in turn, as the independent variable. We did not include a sale dummy in these regressions as such information was not available in our data. ${ }^{17}$ The results of the logit regression for each independent variable are reported in Table 7. Similar to what we found with Dominick's data, 9-ending prices are less likely to change than other prices. Overall, $9 \$$-ending prices are $31.90 \%$, $99 \$$-ending prices $44.59 \%$, \$9-ending prices $45.89 \%$, \$99-ending prices $59.74 \%$, $\$ 9.99$-ending prices $58.90 \%$, and $\$ 99.99$-ending prices are $72.87 \%$, less likely to change than other prices. We obtained similar results for each product category. Although music CDs and video games showed some unexpected results, in 95\% of all possible cases in the category-level analyses, the effect of 9 price-endings on the probability of price changes is negative and significant. Thus, we find that prices tend to be "stuck" at 9-endings, making them more rigid: 9-ending prices are $24 \%$ to $73 \%$ less likely to change than non-9-ending prices.

## VI. The Effect of Price Points on the Size of Price Change

".. if pricing points inhibit price changes, then they might also be expected to affect the sizes of price increases. Specifically if prices that are at price points are fixed longer than other prices, then any subsequent price adjustments might be expected to be larger than average."

Anil Kashyap (1995, p. 267)

If 9-ending prices are less likely to change in comparison to non-9-ending prices, then the average size of change of 9-ending prices should be larger when they do change, in comparison to non-9-ending prices. This assumes that the cost of a price change is the same regardless of the price ending, which is indeed the case according to the menu cost estimates of Levy, et al. (1997, 1998 , 2008) and Dutta, et al. (1999) for large U.S. supermarket and drugstore chains.

In Tables 8 and 9, we report the average size of price changes for $9 \Phi$-ending and non- $9 ¢-$ ending prices, and for $99 ¢$ and non- $99 \$$-ending prices, respectively, in the Dominick's data. According to Table 8, in 23 of the 27 categories, the average change is indeed higher for $9 \$$ ending prices. The exceptions are the categories of frozen dinners, frozen entrees, and frozen juices (perhaps because they have short expiration periods), and front-end candies. Across all

[^8]product categories, the average price change is $47 \$$ if the price ends with $9 \Phi$, in contrast to $37 \$$ change when it does not end with $9 \mathbb{}$, a $27 \%$ difference.

The findings obtained for the 99\$-ending prices are even stronger. According to Table 9, in 26 of the 27 categories (frozen entrees being the only exception), the average change is higher for $99 \$$-ending prices. The differences for individual categories are also bigger here in comparison to Table 8. Across all product categories, the average price change is 57 f if the price ends with $99 \Phi$, in contrast to $42 \$$ change when it does not end with $99 \Phi$, a $35 \%$ difference.

In Tables 10-15, we report the findings for the Internet data. Here we consider prices ending with $9 屯, 99$, $\$ 9$, $\$ 9.99$, $\$ 99$, and $\$ 99.99$. The results are as follows. For $9 \$$-ending prices (Table 10): in 8 out of the 10 categories (the exceptions being PDAs and Notebooks PCs), the average price change is higher by about $12 \%$ if the price ends with $9 \Phi$ in comparison to non$9 \Phi$ ending prices. For $99 \$$-ending prices (Table 11): in 9 out of the 10 categories (the exception being PDAs), the average price change is higher by about $29 \%$ if the price ends with $99 \$$ in comparison to non-99¢-ending prices. For \$9-ending prices (Table 12): in 9 out of the 10 categories (the exception being Music CDs), the average price change is higher by about $97 \%$ if the price ends with $\$ 9$ in comparison to non-\$9-ending prices. For \$9.99-ending prices (Table 13): in all 10 categories, the average price change is higher by about $53 \%$ if the price ends with $\$ 9.99$ in comparison to non-\$9.99-ending prices. For \$99-ending prices (Table 14): in all 8 categories (Music CDs and Video Games contain no prices with \$99-ending), the average price change is higher by about $165 \%$ if the price ends with $\$ 99$ in comparison to non-\$99-ending prices. For \$99.99-ending prices (Table 15): in all 8 categories (Music CDs and Video Games contain no prices with $\$ 99.99$-ending), the average price change is higher by about $150 \%$ if the price ends with $\$ 99.99$ in comparison to non-\$99.99-ending prices.

Thus, the results are very robust in the Internet data as well: in 52 of the 56 cases, the average size of the price change is higher if the price ends with a 9-ending price point. Moreover, in many individual cases the differences in the size of price changes are quite substantial. For example, for $9 \nsubseteq$ price-endings, the average price changes of the $9 \Phi$-ending and non-9\$-ending prices are $\$ 1.30$ and $\$ 1.01$, respectively, a difference of about $30 \%$. In some cases, the differences are even larger. These findings, all significant at the $p<0.0001$ level, are consistent with our predictions: as 9-ending prices are less likely to change, the average size of the change of 9-ending prices are systematically larger when they do change, in comparison to non-9-ending prices.

## VII. Making Sense of Ignoring Cents

"Why are so many items sold for $\$ 2.99$ and so few for $\$ 3.00$ ? There is an enormous temptation to attribute this phenomenon-to a mild form of irrationality in which consumers notice only the first digit of the price and are lulled into thinking that $\$ 2.99$ is 'about $\$ 2.00$ ' instead of 'about $\$ 3.00$.' In fact, this explanation seems so self-evident that even many economists believe it. For all I know, they could be right. Perhaps someday a careful analysis of such behavior will form the basis for a modified economics in which people are assumed to depart from rationality in certain systematic ways."

Steven Landsburg (1995, p. 15)

Having documented overwhelming popularity of 9-ending prices in our data, and having demonstrated that they lead to a substantial degree of price rigidity, we explore what can explain these findings. As Kashyap (1995) notes, the existing economics literature does not offer a "tight" theoretical explanation for the popularity of price points in retail pricing and for their link to price rigidity. ${ }^{18}$

Drawing from the emerging macroeconomic literature on consumer inattention we hypothesize that 9-ending prices may be an outcome of firms' reaction to consumers' inattention. Consumers with limited time often need to assess and compare the prices of dozens and sometimes hundreds of products, and therefore, they are likely to use time-saving devices. One natural action, for example, might be to ignore some price information. ${ }^{19}$ Specifically, we argue that the benefit of paying attention to each additional digit of a price declines as we move from left to right in the price digits. ${ }^{20}$ On the other hand, since people process multi-digit numeric information, including prices, from left to right (Schindler and Kirby, 1997; Hinrichs, et al. 1982; Poltrock and Schwartz, 1984; and Lambert, 1975), the effort they need to recognize, process, and recall numeric information increases as the number of digits increases. Thus, the marginal cost of processing each additional digit increases. The marginal benefit of the rightmost digit is the lowest but its marginal cost is the highest, making it the least valuable among all digits. The last digit, thus, offers the time-constrained consumer the lowest net marginal value giving him an incentive to ignore it. ${ }^{21,22}$ A price-setter that knows that her customers ignore the last digit will

[^9]make it as high as possible, setting it to 9 (Basu, 1997). ${ }^{23}$
This is illustrated in Figure 9. Under consumer inattention, there will be a range of inattention along the demand curve. In this price range, say $\pm 10 ¢$, consumers are inattentive and thus they do not respond to price changes. The optimal pricing strategy in this case will be to set the price at the highest point in the vertical segment of the demand curve, which will be $9 .{ }^{24}$

According to the above argument, consumers' incentive to be attentive increases and therefore, the optimality of the use of the 9 digit decreases as we move from the rightmost digits to the left in the price. This implies that we should still see more $99 \Phi$ endings than $89 థ, 79 థ, \ldots$, $9 ¢$ endings among the rightmost two digits, but that the dominance of 99 © over $89 థ, 79$, etc. should be weaker than the dominance of $9 \Phi$ over $8 \Phi, 7 \Phi$, and so on. This process will continue towards the dollar-digit as well as the ten-dollar digit. Indeed, this is what we observe in both Dominick’s data ( $65 \%$ for $9 \nmid$ vs. $15 \%$ for 99 ¢) and our Internet data ( $31.9 \%$ for $9 \mathbb{4}, 26.3 \%$ for 99థ, 13.5\% for \$9.99, and 3.9\% for \$99.99).

Now consider the implications of consumer inattention for price rigidity. Consumer inattention suggests that there will be a discontinuity in price adjustment within the range of inattention. When changes in market conditions are not large enough to warrant a price change larger than the range of inattention based on the ignored digit, firms might choose not to respond. For example, when the price-setter is facing a price change decision that requires a price increase from $\$ 1.79$ to $\$ 1.80$, the increase will not be optimal if the customers ignore the last digit and perceive the change to be bigger (i.e., as a $10 \$$ increase) than it actually is. Similarly, a price decrease from $\$ 1.79$ to $\$ 1.78$ will have no effect on the quantity demanded if consumers ignore the last digit. Thus, $9 \Phi$-ending prices will lead to price rigidity.

However, when a price change is justified, then the price-setter will have incentive to make price changes in multiples of $10 ¢$. For example, a firm that faces a series of $1 \Phi$ cost increases may not change its price for many periods, but when the firm does react, it may increase the price by 10¢, even though the cost increase in that particular period was only $1 \mathbb{4}$. The implication is that the store could change the price from $\$ 1.79$ to $\$ 1.89$, instead of to $\$ 1.80$, without any additional cost, but with much higher benefit. That would be true even in a world with costs of price adjustment (Mankiw 1985) because of the largely fixed nature of such costs in the retail supermarket industry (Levy, et al. 1997, 1998; Dutta, et al. 1999). This explanation, we believe, offers a possible resolution of the puzzle posed by Landsburg (1995). The empirical

[^10]findings we reported in Section IV are consistent with these predictions. ${ }^{25}$

## VIII. Conclusion

To our knowledge, this is the first study that directly examines the effect of price points on price rigidity across a broad range of product categories, price levels, and retailers, in traditional retailing and Internet-based selling formats, using date from the U.S. We find that 9ending prices are at least $24 \%$ (and as much as $73 \%$ ) less likely to change compared to non-9ending prices. Further, most common price changes are such that they preserve the terminal digits at 9 , and the size of the price changes is larger for 9-ending than non-9-ending prices. Thus, 9-ending prices form a substantial barrier to price changes. ${ }^{26}$ These findings are robust, occurring in both datasets, with a wide range of prices, products, retail formats, and retailers, and lend strong support to the price point theory.

There are a variety of macroeconomic settings where these insights on price points, price rigidity and consumer inattention might be relevant. For example, dropping the smallest currency unit has been a recent topic of debate in the U.S., Canada and Europe. ${ }^{27}$ The smallest currency unit might define the price ranges of customer inattention. This appears to be true in the case of products that are sold through automated devices, such as soda and candy bar vending machines, parking meters, coin-operated laundry machines, etc. ${ }^{28}$ As another example, the common use of price points has recently received a considerable attention in many European

[^11]Union countries in the context of the conversion of prices from local currencies to the Euro. The concern was about the possibility that retailers may have acted opportunistically by rounding their prices upward after conversion to the Euro in their attempt to preserve the price points. ${ }^{29}$

Our findings also may have other potential macroeconomic implications. Typically nominal magnitudes are not important for optimal decision-making. Yet, our results imply decision rules by customers and firms that may affect price points and price adjustments. In such situations, the nominal magnitude of numeric information attached to economic quantities may matter.

In our data, 9 is the most popular terminal digit overall, consistent with the findings reported by Friedman (1967). There may, however, be a variety of other dimensions of price points still to explore. Price points may vary across countries. ${ }^{30}$ For example, Konieczny and Skrzypacz $(2003,2004)$ and Konieczny and Rumler (2007) note that 9-ending prices are particularly popular in the U.S., Canada, Germany, and Belgium, but they are rare in Spain, Italy, Poland, and Hungary. ${ }^{31}$ In Asian countries (Malaysia, Hong Kong, Singapore, Japan, and China), Heeler and Nguyen (2001) find an unusual popularity of 8-endings. ${ }^{32}$ Knotek (2004, 2006a) focuses on other types of pricing practices: the common use of round prices, which he terms "convenient prices" because their use reduces the amount of the change used in a transaction. Levy and Young (2004) report that the nominal price of Coca Cola was fixed for almost 70 years at $5 \Phi$, also a "convenient price." ${ }^{, 33}$ Future work might study this phenomenon across other products, industries, retailers, and countries to assess the generalizability of our results, and to uncover the boundaries of our reasoning.

We conclude by suggesting that the Internet provides a unique context for micro-level

[^12]studies of price setting behavior (Bergen, et al. 2005). The ability to access transaction price data using software agents allows us to explore pricing and price adjustment patterns at low costs at a previously unimaginable level of microeconomic detail. It allows empirical research methods (e.g., massive quasi-experimental data mining methods), to take advantage of natural experiments in the real world (e.g., Kauffman and Lee, 2007; Kauffman and Wood, 2007, 2008). With the expanding retail activities on the Internet, and new techniques and tools that have become available, we expect such opportunities to increase further in the future.

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Table 1. Descriptive Statistics for Weekly Retail Price Observations in Dominick's Data

| Category | Number of Observations | Number of Products | Number of Stores | Mean Price | Std. <br> Dev. | Min. Price | Max. Price |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Analgesics | 3,040,159 | 638 | 93 | \$5.18 | \$2.36 | \$0.47 | \$23.69 |
| Bath Soap | 418,087 | 579 | 93 | \$3.16 | \$1.60 | \$0.47 | \$18.99 |
| Bathroom Tissue | 1,149,953 | 127 | 93 | \$2.10 | \$1.68 | \$0.25 | \$11.99 |
| Beer | 1,966,139 | 787 | 89 | \$5.69 | \$2.70 | \$0.99 | \$26.99 |
| Bottled Juice | 4,294,956 | 506 | 93 | \$2.24 | \$0.97 | \$0.32 | \$8.00 |
| Canned Soup | 5,504,477 | 445 | 93 | \$1.13 | \$0.49 | \$0.23 | \$5.00 |
| Canned Tuna | 2,382,969 | 278 | 93 | \$1.80 | \$1.07 | \$0.22 | \$12.89 |
| Cereals | 4,707,750 | 489 | 93 | \$3.12 | \$0.76 | \$0.25 | \$7.49 |
| Cheeses | 6,752,297 | 657 | 93 | \$2.42 | \$1.12 | \$0.10 | \$16.19 |
| Cigarettes | 1,801,440 | 793 | 93 | \$7.69 | \$7.90 | \$0.59 | \$25.65 |
| Cookies | 7,568,399 | 1,124 | 93 | \$2.10 | \$0.63 | \$0.25 | \$8.79 |
| Crackers | 2,228,265 | 330 | 93 | \$2.01 | \$0.57 | \$0.25 | \$6.85 |
| Dish Detergent | 2,164,726 | 287 | 93 | \$2.34 | \$0.90 | \$0.39 | \$7.00 |
| Fabric Softeners | 2,278,536 | 318 | 93 | \$2.82 | \$1.45 | \$0.10 | \$9.99 |
| Front-End-Candies | 4,437,054 | 503 | 93 | \$0.61 | \$0.24 | \$0.01 | \$6.99 |
| Frozen Dinners | 1,654,049 | 266 | 93 | \$2.37 | \$0.89 | \$0.25 | \$9.99 |
| Frozen Entrees | 7,172,065 | 898 | 93 | \$2.33 | \$1.06 | \$0.25 | \$15.99 |
| Frozen Juices | 2,368,129 | 175 | 93 | \$1.39 | \$0.45 | \$0.22 | \$6.57 |
| Grooming Products | 4,065,657 | 1,381 | 93 | \$2.94 | \$1.37 | \$0.49 | \$11.29 |
| Laundry Detergents | 3,277,439 | 581 | 93 | \$5.61 | \$3.22 | \$0.25 | \$24.49 |
| Oatmeal | 981,034 | 96 | 93 | \$2.65 | \$0.66 | \$0.49 | \$5.00 |
| Paper Towels | 940,740 | 163 | 93 | \$1.50 | \$1.41 | \$0.31 | \$13.99 |
| Refrigerated Juices | 2,166,726 | 225 | 93 | \$2.24 | \$0.91 | \$0.39 | \$7.05 |
| Shampoos | 4,676,362 | 2,930 | 93 | \$2.95 | \$1.79 | \$0.27 | \$29.99 |
| Snack Crackers | 3,487,548 | 420 | 93 | \$2.18 | \$0.57 | \$0.10 | \$8.00 |
| Soaps | 1,835,196 | 334 | 93 | \$2.51 | \$1.48 | \$0.10 | \$10.99 |
| Soft Drinks | 10,741,661 | 1,608 | 93 | \$2.34 | \$1.89 | \$0.10 | \$26.02 |
| Toothbrushes | 1,839,530 | 491 | 93 | \$2.18 | \$0.85 | \$0.39 | \$9.99 |
| Toothpastes | 2,981,513 | 608 | 93 | \$2.43 | \$0.89 | \$0.31 | \$10.99 |
| Total | 98,691,750 | 18,037 | 93 | \$2.59 | \$2.16 | \$0.01 | \$29.99 |

Note: The table covers the entire weekly price data from the Dominick's in its 93 stores for a period of 400 weeks from September 14, 1989 to May 8, 1997. The data are available at: gsbwww.uchicago.edu/kilts/research/db/dominicks/.

Table 2. Descriptive Statistics for the Daily Price Observations in the Internet Data

| Category | Number of <br> Observations | Number of <br> Products | Number of <br> Retailers | Mean <br> Price | Std. <br> Dev. | Min. <br> Price | Max. <br> Price |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Music CDs | 302,914 | 46 | 15 | $\$ 13.46$ | $\$ 3.50$ | $\$ 3.99$ | $\$ 26.98$ |
| Movie DVDs | 447,519 | 49 | 22 | $\$ 27.42$ | $\$ 26.70$ | $\$ 4.95$ | $\$ 144.99$ |
| Video Games | 244,625 | 49 | 38 | $\$ 30.83$ | $\$ 12.57$ | $\$ 4.90$ | $\$ 57.99$ |
| Software | 382,297 | 48 | 83 | $\$ 294.07$ | $\$ 417.60$ | $\$ 4.95$ | $\$ 5,695.00$ |
| Hard Drives | 263,244 | 46 | 73 | $\$ 330.67$ | $\$ 556.29$ | $\$ 39.00$ | $\$ 3,670.98$ |
| PDAs | 148,731 | 45 | 92 | $\$ 346.60$ | $\$ 193.24$ | $\$ 32.99$ | $\$ 956.95$ |
| DVD Players | 220,236 | 49 | 104 | $\$ 369.51$ | $\$ 247.75$ | $\$ 57.99$ | $\$ 1,489.00$ |
| PC Monitors | 319,369 | 51 | 87 | $\$ 682.89$ | $\$ 659.13$ | $\$ 85.78$ | $\$ 3,010.41$ |
| Digital Cameras | 247,917 | 46 | 143 | $\$ 760.12$ | $\$ 688.76$ | $\$ 175.95$ | $\$ 6,000.00$ |
| Notebook PCs | 79,386 | 45 | 45 | $\$ 1,666.68$ | $\$ 475.80$ | $\$ 699.00$ | $\$ 3,199.00$ |
| Total | $\mathbf{2 , 6 5 6 , 2 3 8}$ | $\mathbf{4 7 4}$ | $\mathbf{2 9 3}$ | $\$ 337.06$ | $\$ 536.13$ | $\$ 3.99$ | $\$ 6,000.00$ |

Note: The table covers 743 daily price observations from March 26, 2003 to April 15, 2005, from the Internet retailers. The retailers have many different product categories (e.g., Amazon.com sells books, CDs, DVDs, computer products and electronics, etc.). Consequently, the sum of the number of retailers in each product category will not necessarily be consistent with the total number of stores in all product categories. In addition, some retailers do not have all products (e.g., in our sample, Amazon has 15 music CDs while Barnes \& Noble has 20). Also, the length of individual product's price time series varies due to different life cycle of products. Thus, the number of observations in the Music CDs category, for example, 302,914 , is less than total available combinations (i.e., $46 \times 15 \times 743=$ 512,670.)

Table 3. Top 10 Highest Frequencies of Last Three Digits of Prices in the Internet Data

| Rank | CDs | DVDs | Video Games | SW | PDAs | Hard Drives | DVD Players | PC <br> Monitors | Digital Cameras | Notebook PCs | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 4.99 | \$9.99 | \$9.99 | \$9.00 | \$9.00 | \$9.99 | \$9.99 | \$9.00 | \$9.99 | \$9.00 | \$9.99 |
|  | 7.76\% | 5.13\% | 37.78\% | 11.55\% | 23.43\% | 8.97\% | 23.13\% | 16.60\% | 23.58\% | 48.43\% | 13.17\% |
| 2 | \$2.99 | \$4.99 | \$9.82 | \$9.95 | \$9.99 | \$9.00 | \$9.00 | \$9.99 | \$9.00 | \$9.99 | \$9.00 |
|  | 5.20\% | 4.89\% | 4.51\% | 11.49\% | 15.21\% | 6.18\% | 10.74\% | 8.99\% | 21.60\% | 16.62\% | 9.98\% |
| 3 | \$3.99 | \$3.99 | \$8.95 | \$9.99 | \$9.95 | \$9.95 | \$9.95 | \$9.95 | \$9.95 | \$9.95 | \$9.95 |
|  | 4.35\% | 2.78\% | 3.62\% | 7.72\% | 5.26\% | 4.41\% | 5.77\% | 4.03\% | 8.76\% | 5.01\% | 4.86\% |
| 4 | \$1.99 | \$0.99 | \$7.99 | \$5.00 | \$8.00 | \$5.00 | \$9.97 | \$5.00 | \$5.00 | \$9.98 | \$4.99 |
|  | 4.22\% | 2.72\% | 3.35\% | 4.53\% | 3.09\% | 3.44\% | 5.39\% | 3.42\% | 5.19\% | 3.28\% | 3.24\% |
| 5 | \$3.98 | \$5.99 | \$4.99 | \$0.00 | \$5.00 | \$4.99 | \$9.90 | \$0.00 | \$8.00 | \$5.00 | \$5.00 |
|  | 3.26\% | 2.65\% | 3.20\% | 3.40\% | 2.74\% | 2.57\% | 4.85\% | 2.80\% | 2.80\% | 2.49\% | 2.48\% |
| 6 | \$5.99 | \$2.99 | \$9.95 | \$8.00 | \$4.99 | \$2.00 | \$5.00 | \$5.95 | \$4.99 | \$7.00 | \$2.99 |
|  | 2.96\% | 2.57\% | 2.85\% | 2.84\% | 2.48\% | 2.26\% | 4.13\% | 2.41\% | 2.37\% | 1.73\% | 1.46\% |
| 7 | \$9.99 | \$6.99 | \$9.88 | \$4.95 | \$0.00 | \$7.00 | \$4.99 | \$0.95 | \$7.00 | \$4.00 | \$8.95 |
|  | 2.43\% | 2.37\% | 2.76\% | 2.73\% | 1.85\% | 2.16\% | 3.24\% | 2.33\% | 2.26\% | 1.64\% | 1.45\% |
| 8 | \$4.98 | \$5.98 | \$8.99 | \$8.95 | \$4.95 | \$6.00 | \$8.00 | \$2.95 | \$0.00 | \$4.95 | \$8.00 |
|  | 2.40\% | 2.34\% | 2.72\% | 2.53\% | 1.69\% | 2.14\% | 2.26\% | 2.26\% | 1.85\% | 1.00\% | 1.44\% |
| 9 | \$7.99 | \$1.98 | \$6.99 | \$2.00 | \$8.95 | \$8.99 | \$9.96 | \$8.95 | \$9.98 | \$7.99 | \$7.99 |
|  | 2.26\% | 2.08\% | 2.04\% | 2.21\% | 1.68\% | 2.10\% | 2.21\% | 2.05\% | 1.56\% | 0.97\% | 1.43\% |
| 10 | \$8.99 | \$7.99 | \$6.95 | \$7.00 | \$5.99 | \$3.00 | \$9.94 | \$6.95 | \$9.90 | \$5.99 | \$4.95 |
|  | 2.11\% | 2.07\% | 1.79\% | 2.15\% | 1.47\% | 2.02\% | 1.51\% | 1.98\% | 1.44\% | 0.95\% | 1.42\% |

Note: Each cell contains the last three digits of prices and their proportions in the product category. Boldmarked prices in the first three rows indicate that they are in the top three frequent price endings in each category. The rightmost column includes all categories. The figures in each column are ordered from the most frequent ending to the least frequent ending.

Table 4. Top 10 Highest Frequencies of Last Four Digits of Prices in the Internet Data

| Rank | CDs | DVDs | Games | SW | PDAs | Hard | DVD <br> Drives | Players <br> Ponitors | Digital <br> Cameras | Notebook <br> PCs | Total |
| :---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $\mathbf{1}$ | $\mathbf{\$ 1 4 . 9 9}$ | $\mathbf{\$ 0 9 . 9 9}$ | $\mathbf{\$ 1 9 . 9 9}$ | $\mathbf{\$ 9 9 . 0 0}$ | $\mathbf{\$ 4 9 . 0 0}$ | $\mathbf{\$ 2 9 . 9 9}$ | $\mathbf{\$ 9 9 . 9 9}$ | $\mathbf{\$ 9 9 . 0 0}$ | $\mathbf{\$ 9 9 . 9 9}$ | $\mathbf{\$ 9 9 . 0 0}$ | $\mathbf{\$ 9 9 . 9 9}$ |
|  | $7.48 \%$ | $2.66 \%$ | $\mathbf{1 4 . 3 4 \%}$ | $3.54 \%$ | $5.77 \%$ | $1.30 \%$ | $7.87 \%$ | $5.98 \%$ | $13.51 \%$ | $27.47 \%$ | $3.47 \%$ |
| $\mathbf{2}$ | $\mathbf{\$ 1 2 . 9 9}$ | $\mathbf{\$ 1 3 . 9 9}$ | $\mathbf{\$ 2 9 . 9 9}$ | $\mathbf{\$ 9 9 . 9 5}$ | $\mathbf{\$ 9 9 . 0 0}$ | $\mathbf{\$ 5 9 . 9 9}$ | $\mathbf{\$ 4 9 . 9 9}$ | $\mathbf{\$ 9 9 . 9 9}$ | $\mathbf{\$ 9 9 . 0 0}$ | $\mathbf{\$ 4 9 . 0 0}$ | $\mathbf{\$ 9 9 . 0 0}$ |
|  | $4.90 \%$ | $2.56 \%$ | $10.47 \%$ | $3.46 \%$ | $5.76 \%$ | $1.27 \%$ | $3.72 \%$ | $3.78 \%$ | $9.02 \%$ | $9.29 \%$ | $3.46 \%$ |
| $\mathbf{3}$ | $\mathbf{\$ 1 1 . 9 9}$ | $\mathbf{\$ 1 4 . 9 9}$ | $\mathbf{\$ 4 9 . 9 9}$ | $\mathbf{\$ 9 9 . 9 9}$ | $\mathbf{\$ 9 9 . 9 9}$ | $\mathbf{\$ 0 9 . 9 9}$ | $\mathbf{\$ 1 9 . 9 9}$ | $\mathbf{\$ 4 9 . 0 0}$ | $\mathbf{\$ 9 9 . 9 5}$ | $\mathbf{\$ 9 9 . 9 9}$ | $\mathbf{\$ 1 9 . 9 9}$ |
|  | $4.00 \%$ | $2.31 \%$ | $9.05 \%$ | $3.33 \%$ | $4.82 \%$ | $1.09 \%$ | $2.90 \%$ | $1.89 \%$ | $3.26 \%$ | $8.00 \%$ | $2.16 \%$ |
| $\mathbf{4}$ | $\$ 13.99$ | $\$ 15.99$ | $\$ 39.99$ | $\$ 89.95$ | $\$ 59.00$ | $\$ 49.99$ | $\$ 99.00$ | $\$ 49.99$ | $\$ 49.99$ | $\$ 79.00$ | $\$ 49.99$ |
|  | $3.57 \%$ | $2.14 \%$ | $3.21 \%$ | $1.71 \%$ | $2.44 \%$ | $1.01 \%$ | $2.35 \%$ | $1.72 \%$ | $3.18 \%$ | $3.04 \%$ | $2.00 \%$ |
| $\mathbf{5}$ | $\$ 13.98$ | $\$ 15.98$ | $\$ 19.82$ | $\$ 49.95$ | $\$ 79.00$ | $\$ 59.00$ | $\$ 69.99$ | $\$ 29.00$ | $\$ 49.00$ | $\$ 99.98$ | $\$ 29.99$ |
|  | $3.26 \%$ | $2.03 \%$ | $2.74 \%$ | $1.50 \%$ | $2.44 \%$ | $0.91 \%$ | $2.30 \%$ | $1.62 \%$ | $3.15 \%$ | $2.84 \%$ | $1.55 \%$ |
| $\mathbf{6}$ | $\$ 15.99$ | $\$ 10.99$ | $\$ 18.95$ | $\$ 79.95$ | $\$ 49.99$ | $\$ 99.99$ | $\$ 49.00$ | $\$ 39.00$ | $\$ 29.00$ | $\$ 29.00$ | $\$ 49.00$ |
|  | $2.43 \%$ | $1.83 \%$ | $2.11 \%$ | $1.37 \%$ | $2.41 \%$ | $0.86 \%$ | $1.87 \%$ | $1.35 \%$ | $1.99 \%$ | $2.84 \%$ | $1.43 \%$ |
| $\mathbf{7}$ | $\$ 14.98$ | $\$ 11.98$ | $\$ 19.88$ | $\$ 19.00$ | $\$ 19.00$ | $\$ 79.99$ | $\$ 79.99$ | $\$ 59.00$ | $\$ 79.99$ | $\$ 29.99$ | $\$ 14.99$ |
|  | $2.40 \%$ | $1.44 \%$ | $1.99 \%$ | $1.35 \%$ | $2.00 \%$ | $0.86 \%$ | $1.83 \%$ | $1.27 \%$ | $1.81 \%$ | $2.17 \%$ | $1.40 \%$ |
| $\mathbf{8}$ | $\$ 10.99$ | $\$ 10.95$ | $\$ 17.99$ | $\$ 79.00$ | $\$ 19.99$ | $\$ 39.99$ | $\$ 39.99$ | $\$ 19.00$ | $\$ 79.00$ | $\$ 30.00$ | $\$ 99.95$ |
|  | $1.89 \%$ | $1.40 \%$ | $1.33 \%$ | $1.14 \%$ | $1.59 \%$ | $0.83 \%$ | $1.65 \%$ | $1.07 \%$ | $1.62 \%$ | $1.89 \%$ | $1.09 \%$ |
| $\mathbf{9}$ | $\$ 15.18$ | $\$ 16.99$ | $\$ 48.95$ | $\$ 89.00$ | $\$ 29.99$ | $\$ 79.00$ | $\$ 29.00$ | $\$ 69.00$ | $\$ 39.00$ | $\$ 19.99$ | $\$ 09.99$ |
|  | $1.89 \%$ | $1.39 \%$ | $1.28 \%$ | $1.1 \%$ | $1.41 \%$ | $0.73 \%$ | $1.64 \%$ | $1.04 \%$ | $1.34 \%$ | $1.55 \%$ | $0.97 \%$ |
| $\mathbf{1 0}$ | $\$ 7.99$ | $\$ 17.99$ | $\$ 49.95$ | $\$ 19.95$ | $\$ 39.00$ | $\$ 39.00$ | $\$ 79.00$ | $\$ 79.00$ | $\$ 69.00$ | $\$ 49.99$ | $\$ 79.00$ |
|  | $1.85 \%$ | $1.34 \%$ | $1.24 \%$ | $1.05 \%$ | $1.34 \%$ | $0.71 \%$ | $1.62 \%$ | $1.00 \%$ | $1.32 \%$ | $1.53 \%$ | $0.87 \%$ |

Note: Each cell contains the last four digits of prices and their proportions in the product category. Boldmarked prices in the first three rows indicate that they are in the top three frequent price endings in each category. The rightmost column includes all categories.

Table 5. Top 10 Highest Frequencies of Price Changes in the Internet Data

| Rank | CD | DVD | Video <br> Game | SW | PDA | Hard <br> Drive | DVD <br> Player | PC <br> Monitor | Digital <br> Camera | Notebook PC | Total | $\begin{aligned} & \text { W/O } \\ & 3 \text { Cat } \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | \$1.00 | \$1.00 | \$10.00 | \$1.00 | \$10.00 | \$1.00 | \$10.00 | \$1.00 | \$10.00 | \$50.00 | \$1.00 | 00 |
|  | 10.26\% | 7.73\% | 11.44\% | 6.78\% | 7.54\% | 10.03\% | 4.46\% | 3.29\% | 8.09\% | 11.30\% | 6.74\% | 5.63\% |
| 2 | \$0.10 | \$0.20 | \$1.00 | \$2.00 | \$5.00 | \$2.00 | \$20.00 | \$2.00 | \$20.00 | \$100.00 | \$2.00 | \$2.00 |
|  | 6.77\% | 3.42\% | 9.82\% | 5.15\% | 4.41\% | 7.54\% | 3.95\% | 3.29\% | 5.89\% | 7.63\% | 4.49\% | 4.66\% |
| 3 | \$2.00 | \$2.00 | \$5.00 | \$5.00 | \$2.00 | \$3.00 | \$30.00 | \$10.00 | \$4.00 | \$200.00 | \$10.00 | \$10.00 |
|  | 5.22\% | 3.40\% | 7.40\% | 4.11\% | 4.02\% | 5.55\% | 2.70\% | 3.27\% | 3.46\% | 3.97\% | 3.24\% | 4.31\% |
| 4 | \$0.20 | \$0.01 | \$2.00 | \$10.00 | \$1.00 | \$4.00 | \$5.00 | \$3.00 | \$5.00 | \$20.00 | \$3.00 | \$3.00 |
|  | 3.59\% | 2.34\% | 5.57\% | 3.76\% | 3.41\% | 4.00\% | 2.51\% | 3.02\% | 3.44\% | 3.05\% | 3.09\% | 3.60\% |
| 5 | \$0.01 | \$0.09 | \$20.00 | \$3.00 | \$20.00 | \$5.00 | \$1.00 | \$5.00 | \$2.00 | \$10.00 | \$5.00 | \$5.00 |
|  | 3.46\% | 2.30\% | 4.69\% | 3.56\% | 3.24\% | 3.98\% | 2.13\% | 2.23\% | 3.28\% | 2.44\% | 2.72\% | 3.38\% |
| 6 | \$0.50 | \$0.10 | \$3.00 | \$4.00 | \$30.00 | \$10.00 | \$3.00 | \$4.00 | \$6.00 | \$60.00 | \$4.00 | \$4.00 |
|  | 2.45\% | 2.29\% | 4.25\% | 3.04\% | 2.57\% | 2.83\% | 2.13\% | 1.91\% | 3.28\% | 2.14\% | 2.30\% | 2.90\% |
| 7 | \$0.06 | \$3.00 | \$3.06 | \$20.00 | \$3.00 | \$6.00 | \$4.00 | \$6.00 | \$50.00 | \$30.00 | \$20.00 | \$20.00 |
|  | 2.06\% | 2.21\% | 2.64\% | 2.44\% | 2.35\% | 2.10\% | 1.95\% | 1.83\% | 2.97\% | 1.98\% | 1.80\% | 2.56\% |
| 8 | \$0.14 | \$0.30 | \$0.11 | \$6.00 | \$6.00 | \$7.00 | \$2.00 | \$20.00 | \$1.00 | \$40.00 | \$6.00 | \$6.00 |
|  | 1.88\% | 1.79\% | 2.05\% | 2.19\% | 2.29\% | 1.84\% | 1.57\% | 1.61\% | 2.87\% | 1.83\% | 1.55\% | 2.18\% |
| 9 | \$0.02 | \$0.08 | \$18.00 | \$0.01 | \$4.00 | \$8.00 | \$6.00 | \$30.00 | \$30.00 | \$150.00 | \$0.10 | \$30.00 |
|  | 1.69\% | 1.31\% | 1.61\% | 2.03\% | 1.90\% | 1.08\% | 1.51\% | 1.54\% | 2.87\% | 1.83\% | 1.38\% | 1.50\% |
| 10 | \$0.30 | \$0.50 | \$7.00 | \$8.00 | \$15.00 | \$20.00 | \$50.00 | \$7.00 | \$3.00 | \$70.00 | \$0.01 | \$7.00 |
|  | 1.69\% | 1.26\% | 1.39\% | 1.54\% | 1.79\% | 1.06\% | 1.51\% | 1.44\% | 2.64\% | 1.68\% | 1.38\% | 1.47\% |

Note: The rightmost column shows the results after three product categories (CDs, DVDs, and video games) are left out. Bold-marked prices in the first three rows indicate that they are in the top three frequent price changes in each category.

Table 6. Results of the Logit Regression (Equation 1) Estimation for Dominick's Data

| Category | 9¢-Ending |  |  |  | 99¢-Ending |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $D_{9}$ (9-Ending $=1$ ) |  | $D_{\text {Sale }}($ Sale = 1) |  | $D_{99}$ (9-Ending $=1$ ) |  | $D_{\text {Sale }}($ Sale = 1) |  |
|  | Coeff. | O/R | Coeff. | O/R | Coeff. | O/R | Coeff. | O/R |
| Analgesics | -0.6781 | 0.51 | 3.9829 | 52.63 | -0.1847 | 0.83 | 3.9805 | 52.63 |
| Bath Soap | -0.8155 | 0.44 | 4.6464 | 100.00 | -0.2273 | 0.80 | 4.7925 | 125.00 |
| Bathroom Tissues | -0.5036 | 0.60 | 3.6723 | 40.00 | -0.3426 | 0.71 | 3.6795 | 40.00 |
| Bottled Juices | -0.2891 | 0.75 | 4.1268 | 62.50 | -0.2042 | 0.81 | 4.1422 | 62.50 |
| Canned Soup | -0.1112 | 0.89 | 4.6189 | 100.00 | -0.1629 | 0.85 | 4.6238 | 100.00 |
| Canned Tuna | -0.5331 | 0.59 | 4.5788 | 100.00 | -0.4714 | 0.62 | 4.5281 | 90.91 |
| Cereals | -0.2558 | 0.77 | 4.7368 | 111.11 | -0.1603 | 0.85 | 4.7239 | 111.11 |
| Cheeses | -0.9142 | 0.40 | 3.8187 | 45.45 | -0.6098 | 0.54 | 3.8378 | 45.45 |
| Cookies | -0.8173 | 0.44 | 4.1490 | 62.50 | -0.1876 | 0.83 | 4.2162 | 66.67 |
| Crackers | -0.4412 | 0.64 | 4.0389 | 55.56 | -0.0441 | 0.96 | 4.1185 | 62.50 |
| Dish Detergent | -0.6283 | 0.53 | 4.7074 | 111.11 | -0.6024 | 0.55 | 4.7350 | 111.11 |
| Fabric Softeners | -0.3779 | 0.69 | 4.6161 | 100.00 | -0.1980 | 0.82 | 4.5797 | 100.00 |
| Front-end-candies | -0.4477 | 0.64 | 4.8119 | 125.00 | -1.3781 | 0.25 | 4.8630 | 125.00 |
| Frozen Dinners | -0.5808 | 0.56 | 3.5407 | 34.48 | -0.4377 | 0.65 | 3.7235 | 41.67 |
| Frozen Entrees | -0.5642 | 0.57 | 3.2641 | 26.32 | -0.1291 | 0.88 | 3.4461 | 31.25 |
| Frozen Juices | -0.2451 | 0.78 | 3.9482 | 52.63 | -0.1008 | 0.90 | 3.9182 | 50.00 |
| Grooming Products | -0.9030 | 0.41 | 3.3588 | 28.57 | -0.2406 | 0.79 | 3.6612 | 38.46 |
| Laundry Detergents | -0.5783 | 0.56 | 4.1731 | 66.67 | -0.1446 | 0.87 | 4.1543 | 62.50 |
| Oatmeal | -0.5805 | 0.56 | 4.1839 | 66.67 | -0.2548 | 0.78 | 4.1707 | 66.67 |
| Paper Towels | -0.5186 | 0.60 | 4.3241 | 76.92 | -0.1546 | 0.86 | 4.2669 | 71.43 |
| Refrigerated Juices | -0.5042 | 0.60 | 3.6385 | 38.46 | -0.2908 | 0.75 | 3.6428 | 38.46 |
| Shampoos | -0.7868 | 0.46 | 3.1548 | 23.26 | -0.2957 | 0.74 | 3.3005 | 27.03 |
| Snack Crackers | -0.8517 | 0.43 | 3.8756 | 47.62 | -0.3930 | 0.68 | 4.1214 | 62.50 |
| Soaps | -0.6709 | 0.51 | 4.2641 | 71.43 | -0.3583 | 0.70 | 4.2807 | 71.43 |
| Soft Drinks | -0.6709 | 0.51 | 4.2641 | 71.43 | -0.3583 | 0.70 | 4.2807 | 71.43 |
| Tooth Brushes | -0.3154 | 0.73 | 3.6447 | 38.46 | -0.0709 | 0.93 | 3.6285 | 37.04 |
| Tooth Pastes | -0.2343 | 0.79 | 3.7560 | 43.48 | -0.2760 | 0.76 | 3.7405 | 41.67 |
| Average |  | 0.59 |  | 64.90 |  | 0.76 |  | 66.83 |

Note: $D_{9}$ (or $D_{99}$ ) is 9-ending dummy variable, which equals 1 if the price ends with 9 (or 99 ) and 0 otherwise. $D_{\text {Sale }}$ is a sale dummy, which equals 1 if the product is on sale in the given week and 0 otherwise. All $p$-values are less than 0.0001 . The average odds ratios ( $\mathrm{O} / \mathrm{R}$ ) reported in the last row of the table are the simple averages of the odds ratios for each product category.

Table 7. Results of Logit Regression (Equation 1) Estimation for the Internet Data

| Category | 9థ- <br> Endings | 99¢- <br> Endings | \$9- <br> Endings | \$99- <br> Endings | \$9.99- <br> Endings | \$99.99- <br> Endings |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Music CDs | $\begin{array}{r} \hline-0.0743^{* *} \\ (0.9284) \\ \hline \end{array}$ | $\begin{array}{r} \hline-0.5085^{* * *} \\ (0.6014) \\ \hline \end{array}$ | $\begin{gathered} \hline \hline-0.0174 \\ (0.9827) \\ \hline \end{gathered}$ |  | $\begin{array}{r} \hline-0.4283^{* *} \\ (0.6516) \\ \hline \end{array}$ |  |
| Movie DVDs | $\begin{array}{r} -0.5036^{* * *} \\ (0.6043) \\ \hline \end{array}$ | $\begin{array}{r} -0.6154^{* * *} \\ (0.5404) \end{array}$ | $\begin{array}{r} -0.2596^{* * *} \\ (0.7714) \end{array}$ |  | $\begin{array}{r} -0.7835^{* * *} \\ (0.4568) \end{array}$ |  |
| Video Games | $\begin{aligned} & 0.1087^{*} \\ & (1.1148) \end{aligned}$ | $\begin{aligned} & 0.0558 \\ & (1.0573) \\ & \hline \end{aligned}$ | $\begin{array}{r} -0.2951^{* * *} \\ (0.7445) \\ \hline \end{array}$ |  | $\begin{array}{r} -0.2779^{* * *} \\ (0.7574) \\ \hline \end{array}$ |  |
| Software | $\begin{gathered} -0.3011^{* * *} \\ (0.7400) \\ \hline \end{gathered}$ | $\begin{array}{r} -0.48899^{* * *} \\ (0.6133) \\ \hline \end{array}$ | $\begin{array}{r} -0.6148^{* * *} \\ (0.5407) \\ \hline \end{array}$ | $\begin{array}{r} -1.0667^{* * *} \\ (0.3441) \\ \hline \end{array}$ | $\begin{array}{r} -0.9012^{* * *} \\ (0.4061) \\ \hline \end{array}$ | $\begin{array}{r} -1.3153^{* * *} \\ (0.2684) \\ \hline \end{array}$ |
| PDAs | $\begin{array}{r} -0.2553^{* *} \\ (0.7747) \\ \hline \end{array}$ | $\begin{array}{r} -0.3586^{* * *} \\ (0.6986) \\ \hline \end{array}$ | $\begin{array}{r} -0.4654^{* * *} \\ (0.6279) \\ \hline \end{array}$ | $\begin{array}{r} -0.7911^{* * *} \\ (0.4533) \\ \hline \end{array}$ | $\begin{array}{r} -0.5418^{* * *} \\ (0.5817) \\ \hline \end{array}$ | $\begin{array}{r} -1.1819^{* * *} \\ (0.3067) \\ \hline \end{array}$ |
| Hard Drives | $\begin{array}{r} -0.2806^{* * *} \\ (0.7553) \\ \hline \end{array}$ | $\begin{array}{r} -0.3698^{* * *} \\ (0.6909) \\ \hline \end{array}$ | $\begin{array}{r} -0.4711^{* * *} \\ (0.6243) \\ \hline \end{array}$ | $\begin{array}{r} -0.6199^{* * *} \\ (0.5380) \\ \hline \end{array}$ | $\begin{array}{r} -0.6796^{* * *} \\ (0.5068) \\ \hline \end{array}$ | $\begin{array}{r} -0.5254^{* * *} \\ (0.5913) \end{array}$ |
| DVD Players | $\begin{array}{r} -0.4939^{* * *} \\ (0.6102) \\ \hline \end{array}$ | $\begin{gathered} -0.5763^{* * *} \\ (0.5620) \end{gathered}$ | $\begin{array}{r} -0.6695^{* * *} \\ (0.5120) \end{array}$ | $\begin{gathered} -0.6790^{* * *} \\ (0.5071) \end{gathered}$ | $\begin{gathered} -0.6389^{* * *} \\ (0.5279) \\ \hline \end{gathered}$ | $\begin{array}{r} -0.8103^{* * *} \\ (0.4447) \\ \hline \end{array}$ |
| PC Monitors | $\begin{array}{r} -0.2729^{* * *} \\ (0.7612) \\ \hline \end{array}$ | $\begin{array}{r} -0.4617^{* * *} \\ (0.6302) \end{array}$ | $\begin{array}{r} -0.5507^{* * *} \\ (0.5766) \\ \hline \end{array}$ | $\begin{array}{r} -0.8433^{* * *} \\ (0.4303) \\ \hline \end{array}$ | $\begin{array}{r} -0.8375^{* * *} \\ (0.4328) \\ \hline \end{array}$ | $\begin{array}{r} -1.2927^{* * *} \\ (0.2745) \\ \hline \end{array}$ |
| Digital Cameras | $\begin{array}{r} -0.4389^{* * *} \\ (0.6447) \\ \hline \end{array}$ | $\begin{array}{r} -0.4933^{* * *} \\ (0.6106) \end{array}$ | $\begin{array}{r} -0.5297^{* * *} \\ (0.5888) \end{array}$ | $\begin{array}{r} -1.1229^{* * *} \\ (0.3253) \end{array}$ | $\begin{array}{r} -0.5879^{* * *} \\ (0.5500) \end{array}$ | $\begin{array}{r} -1.4480^{* * *} \\ (0.2350) \end{array}$ |
| Notebook PCs | $\begin{array}{r} -0.5566^{* * *} \\ (0.5731) \\ \hline \end{array}$ | $\begin{array}{r} -0.8885^{* * *} \\ (0.4113) \\ \hline \end{array}$ | $\begin{gathered} -1.0680^{* * *} \\ (0.3437) \\ \hline \end{gathered}$ | $\begin{gathered} -0.7654^{* * *} \\ (0.4652) \\ \hline \end{gathered}$ | $\begin{array}{r} -0.9528^{* * *} \\ (0.3856) \\ \hline \end{array}$ | $\begin{array}{r} -1.1891^{* * *} \\ (0.3045) \\ \hline \end{array}$ |
| Total | $\begin{array}{r} -0.3690^{* * *} \\ (0.6914) \\ \hline \end{array}$ | $\begin{array}{r} -0.5703^{* * *} \\ (0.5653) \end{array}$ | $\begin{array}{r} -0.6472^{* * *} \\ (0.5235) \end{array}$ | $\begin{array}{r} -1.0179^{* * *} \\ (0.3613) \end{array}$ | $\begin{array}{r} -0.8761^{* * *} \\ (0.4164) \\ \hline \end{array}$ | $\begin{array}{r} -1.3528^{* * *} \\ (0.2585) \\ \hline \end{array}$ |
| Note: Each cell contains a coefficient and odds ratio in parenthesis; significance levels: ${ }^{* * *}<0.01$, ${ }^{* *}$ $<0.05,{ }^{*}<0.10$. The estimated coefficients in italics indicate unsupportive results. |  |  |  |  |  |  |

Table 8. Average Size of Price Change in Dominick's Data: 9¢- vs. Non-9¢-Ending Prices

| Category | 9¢-Ending |  | Non-9¢-Ending |  | Corr. | t-Stat | $p$-Value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean Price Change | Sample Size | Mean Price Change | Sample Size |  |  |  |
| Analgesics | \$0.7625 | 367,969 | \$0.4672 | 102,550 | 0.173 | 76.47 | 0.000 |
| Bath Soap | \$0.5786 | 58,735 | \$0.5473 | 18,298 | 0.019 | 64.41 | 0.000 |
| Bathroom Tissues | \$0.2499 | 156,863 | \$0.2260 | 184,414 | 0.031 | 210.19 | 0.000 |
| Bottled Juices | \$0.3121 | 457,490 | \$0.2650 | 583,025 | 0.060 | 255.92 | 0.000 |
| Canned Soup | \$0.2196 | 304,439 | \$0.1948 | 741,357 | 0.033 | 162.99 | 0.000 |
| Canned Tuna | \$0.1946 | 170,023 | \$0.1421 | 281,703 | 0.091 | 268.59 | 0.000 |
| Cereals | \$0.5010 | 271,757 | \$0.4701 | 494,597 | 0.027 | -153.45 | 0.000 |
| Cheeses | \$0.2943 | 872,489 | \$0.2128 | 1,039,738 | 0.122 | 505.32 | 0.000 |
| Cookies | \$0.4947 | 1,135,112 | \$0.3656 | 709,697 | 0.129 | 359.98 | 0.000 |
| Crackers | \$0.2964 | 283,278 | \$0.2366 | 279,353 | 0.098 | 317.30 | 0.000 |
| Dish Detergent | \$0.2798 | 240,532 | \$0.2119 | 183,222 | 0.133 | 392.69 | 0.000 |
| Fabric Softeners | \$0.3955 | 212,288 | \$0.2597 | 191,319 | 0.168 | 210.90 | 0.000 |
| Front-end-candies* | \$0.1454 | 137,453 | \$0.2164 | 385,234 | -0.113 | 86.40 | 0.000 |
| Frozen Dinners* | \$0.5008 | 230,423 | \$0.5452 | 336,201 | -0.033 | -109.08 | 0.000 |
| Frozen Entrees* | \$0.7031 | 883,284 | \$0.7551 | 1,183,557 | -0.029 | -432.43 | 0.000 |
| Frozen Juices* | \$0.2567 | 301,114 | \$0.2816 | 395,344 | -0.029 | 203.22 | 0.000 |
| Grooming Products | \$0.6285 | 1,017,513 | \$0.4849 | 287,969 | 0.085 | 266.89 | 0.000 |
| Laundry Detergents | \$0.9036 | 446,767 | \$0.5548 | 210,342 | 0.194 | -103.55 | 0.000 |
| Oatmeal | \$0.4239 | 72,753 | \$0.4115 | 107,971 | 0.012 | -8.37 | 0.000 |
| Paper Towels | \$0.1913 | 109,596 | \$0.1702 | 152,846 | 0.030 | 205.91 | 0.000 |
| Refrigerated Juices | \$0.3780 | 405,144 | \$0.2987 | 418,402 | 0.115 | 243.81 | 0.000 |
| Shampoos | \$1.4476 | 1,916,061 | \$1.0888 | 238,976 | 0.065 | -440.40 | 0.000 |
| Snack Crackers | \$0.3251 | 488,341 | \$0.2903 | 405,005 | 0.047 | 371.01 | 0.000 |
| Soaps | \$0.3147 | 180,935 | \$0.1700 | 190,632 | 0.218 | 280.21 | 0.000 |
| Soft Drinks | \$1.0409 | 4,614,455 | \$0.6155 | 1,219,151 | 0.140 | -311.91 | 0.000 |
| Tooth Brushes | \$0.5063 | 350,705 | \$0.3653 | 123,840 | 0.191 | 376.47 | 0.000 |
| Tooth Pastes | \$0.4255 | 468,688 | \$0.3497 | 291,045 | 0.108 | 340.88 | 0.000 |
| Total | \$0.7452 | 16,154,207 | \$0.4033 | 10,755,788 | 0.181 | -44.00 | 0.000 |
| Average | \$0.4730 |  | \$0.3777 |  |  |  |  |
| Median | \$0.3955 |  | \$0.2987 |  |  |  |  |
| Note: Categories with unsupportive results are indicated by $*$ and italic. Corr. is the correlation between 9 -ending prices and the size of price change. The $\boldsymbol{p}$-value is a significance level derived from a paired-sample t-test. Cross-category paired t-tests showed that the price changes are of a larger magnitude when prices end with $9\left(t_{26}=3.911, p=.001\right)$. |  |  |  |  |  |  |  |

Table 9. Average Size of Price Change in Dominick’s Data: 99¢- vs. Non-99¢-Ending Prices

| Category | 99¢-Ending |  | Non-99¢-Ending |  | Corr. | t-Stat | $p$-Value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean Price Change | Sample Size | Mean Price Change | Sample Size |  |  |  |
| Analgesics | \$0.8931 | 106,038 | \$0.6415 | 364,481 | 0.149 | -424.68 | 0.000 |
| Bath Soap | \$0.7149 | 15,608 | \$0.5346 | 61,425 | 0.102 | -130.63 | 0.000 |
| Bathroom Tissues | \$0.3302 | 36,944 | \$0.2257 | 304,333 | 0.085 | -159.40 | 0.000 |
| Bottled Juices | \$0.3760 | 104,451 | \$0.2756 | 936,064 | 0.077 | -397.80 | 0.000 |
| Canned Soup | \$0.2703 | 56,527 | \$0.1981 | 989,269 | 0.048 | -378.08 | 0.000 |
| Canned Tuna | \$0.3303 | 19,566 | \$0.1543 | 432,160 | 0.128 | -246.09 | 0.000 |
| Cereals | \$0.6374 | 56,437 | \$0.4686 | 709,917 | 0.080 | -602.23 | 0.000 |
| Cheeses | \$0.3563 | 160,237 | \$0.2403 | 1,751,990 | 0.097 | -557.70 | 0.000 |
| Cookies | \$0.5612 | 270,448 | \$0.4251 | 1,574,361 | 0.099 | -707.64 | 0.000 |
| Crackers | \$0.4902 | 62,297 | \$0.2489 | 500,334 | 0.165 | -292.33 | 0.000 |
| Dish Detergent | \$0.3273 | 52,117 | \$0.2397 | 371,637 | 0.113 | -211.97 | 0.000 |
| Fabric Softeners | \$0.5585 | 62,370 | \$0.2896 | 341,237 | 0.241 | -237.44 | 0.000 |
| Front-end-candies | \$0.2326 | 11,923 | \$0.1969 | 510,764 | 0.019 | -405.47 | 0.000 |
| Frozen Dinners | \$0.5585 | 56,617 | \$0.5237 | 510,007 | 0.016 | -449.31 | 0.000 |
| Frozen Entrees* | \$0.7229 | 188,496 | \$0.7339 | 1,878,345 | -0.004 | -1002.78 | 0.000 |
| Frozen Juices | \$0.2794 | 67,862 | \$0.2699 | 628,596 | 0.007 | -279.32 | 0.000 |
| Grooming Products | \$0.6756 | 247,298 | \$0.5785 | 1,058,184 | 0.054 | -595.05 | 0.000 |
| Laundry Detergents | \$1.1475 | 158,974 | \$0.6785 | 498,135 | 0.239 | -527.15 | 0.000 |
| Oatmeal | \$0.5420 | 12,921 | \$0.4068 | 167,803 | 0.068 | -261.98 | 0.000 |
| Paper Towels | \$0.3555 | 15,137 | \$0.1682 | 247,305 | 0.126 | -158.70 | 0.000 |
| Refrigerated Juices | \$0.4874 | 101,063 | \$0.3168 | 722,483 | 0.162 | -447.38 | 0.000 |
| Shampoos | \$1.6000 | 503,157 | \$1.3492 | 1,651,880 | 0.062 | -987.01 | 0.000 |
| Snack Crackers | \$0.3673 | 97,690 | \$0.3022 | 795,656 | 0.055 | -403.24 | 0.000 |
| Soaps | \$0.3907 | 43,874 | \$0.2203 | 327,693 | 0.166 | -176.49 | 0.000 |
| Soft Drinks | \$1.2138 | 1,385,935 | \$0.8704 | 4,447,671 | 0.118 | -1370.09 | 0.000 |
| Tooth Brushes | \$0.5972 | 108,407 | \$0.4317 | 366,138 | 0.215 | -351.86 | 0.000 |
| Tooth Pastes | \$0.5097 | 117,086 | \$0.3758 | 642,647 | 0.141 | -457.98 | 0.000 |
| Total | \$0.9156 | 4,119,480 | \$0.5532 | 22,790,515 | 0.143 | -2494.28 | 0.000 |
| Average | \$0.5750 |  | \$0.4209 |  |  |  |  |
| Median | \$0.5097 |  | \$0.3168 |  |  |  |  |

Note: Categories with unsupportive results are indicated by * and italic. Corr. is the correlation between 9ending prices and the size of price change. The $\boldsymbol{p}$-value is a significance level derived from a paired-sample t -test. Cross-category paired t -tests showed that the price changes are of a larger magnitude when prices end with $9\left(t_{26}=7.657, p=.000\right)$.

Table 10. Average Size of Price Change in Internet Data: 9¢- vs. Non-9¢-Ending Prices

|  | 9世-Ending |  | Non-9¢-Ending |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Category | Mean Price <br> Change | Sample <br> Size | Mean Price <br> Change | Sample <br> Size | Corr. | t-Stat | p-Value |
| Music CDs | $\$ 1.30$ | 2,268 | $\$ 1.01$ | 2,352 | 0.097 | 29.45 | 0.000 |
| Movie DVDs | $\$ 2.71$ | 2,813 | $\$ 1.68$ | 5,888 | 0.122 | 40.16 | 0.000 |
| Video Games | $\$ 8.12$ | 832 | $\$ 6.95$ | 532 | 0.075 | 34.55 | 0.000 |
| Software | $\$ 14.94$ | 778 | $\$ 13.51$ | 4,751 | 0.014 | 27.60 | 0.000 |
| PDAs* | $\$ 22.30$ | 355 | $\$ 25.86$ | 1,436 | -0.039 | 28.88 | 0.000 |
| Hard Drives | $\$ 27.65$ | 1,435 | $\$ 14.29$ | 5,517 | 0.097 | 25.10 | 0.000 |
| DVD Players | $\$ 36.02$ | 383 | $\$ 28.43$ | 1,210 | 0.065 | 24.07 | 0.000 |
| PC Monitors | $\$ 41.35$ | 809 | $\$ 28.45$ | 5,150 | 0.072 | 37.83 | 0.000 |
| Digital Cameras | $\$ 45.76$ | 852 | $\$ 36.97$ | 3,018 | 0.046 | 30.60 | 0.000 |
| Notebook PCs* | $\$ 86.42$ | 92 | $\$ 97.58$ | 563 | -0.031 | 19.57 | 0.000 |
| Total* | $\$ 16.08$ | 10,617 | 17.87 | 30.417 | -0.016 | 69.30 | 0.000 |
| Average | $\$ 28.66$ |  | $\$ 25.47$ |  |  |  |  |
| Median | $\$ 25.00$ |  | $\$ 20.00$ |  |  |  |  |

Note: Categories with unsupportive results are indicated by * and italic. Corr. is the correlation between 9 -ending prices and the size of price change. The $\boldsymbol{p}$-value is a significance level derived from a paired-sample t-test. Cross-category paired t-tests showed that the price changes are of a larger magnitude when prices end with $9\left(t_{9}=1.324, p=.10\right)$.

Table 11. Average Size of Price Change in Internet Data: 99¢- vs. Non-99¢-Ending Prices

|  | 99థ-Ending |  | Non-99థ-Ending |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Category | Mean Price <br> Change | Sample <br> Size | Mean Price <br> Change | Sample <br> Size | Corr. | t-Stat | $\boldsymbol{p}$-Value |
| Music CDs | $\$ 1.95$ | 1,142 | $\$ 0.89$ | 3,478 |  | 43.25 | 0.000 |
| Movie DVDs | $\$ 3.39$ | 1,532 | $\$ 1.72$ | 7,169 | 0.160 | 43.81 | 0.000 |
| Video Games | $\$ 8.45$ | 744 | $\$ 6.72$ | 620 | 0.113 | 34.96 | 0.000 |
| Software | $\$ 16.58$ | 553 | $\$ 13.39$ | 4,976 | 0.026 | 27.69 | 0.000 |
| PDAs* | $\$ 23.74$ | 300 | $\$ 25.44$ | 1,491 | -0.017 | 28.92 | 0.000 |
| Hard Drives | $\$ 30.28$ | 1,083 | $\$ 14.60$ | 5,869 | 0.102 | 25.17 | 0.000 |
| DVD Players | $\$ 39.32$ | 329 | $\$ 27.90$ | 1,264 | 0.093 | 24.10 | 0.000 |
| PC Monitors | $\$ 48.86$ | 544 | $\$ 28.33$ | 5,415 | 0.096 | 37.89 | 0.000 |
| Digital Cameras | $\$ 47.53$ | 852 | $\$ 36.78$ | 3,018 | 0.054 | 30.62 | 0.000 |
| Notebook PCs | $\$ 103.15$ | 64 | $\$ 95.24$ | 591 | 0.019 | 19.58 | 0.000 |
| Total | $\$ 20.59$ | 7,056 | $\$ 16.75$ | 33,978 | 0.029 | 69.68 | 0.000 |
| Average | $\$ 32.33$ |  | $\$ 25.10$ |  |  |  |  |
| Median | $\$ 27.00$ |  | $\$ 20.00$ |  |  |  |  |

Note: Categories with unsupportive results are indicated by * and italic. Corr. is the correlation between 9 -ending prices and the size of price change. The $\boldsymbol{p}$-value is a significance level derived from a paired-sample t-test. Cross-category paired t-tests showed that the price changes are of a larger magnitude when prices end with $9\left(t_{9}=3.148, p=.006\right)$.

Table 12. Average Size of Price Change in Internet Data: \$9- vs. Non-\$9-Endings

| Category | \$9-Ending |  | Non-\$9-Ending |  | Corr. | t-Stat | $p$-Value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean Price Change | Sample Size | Mean Price Change | Sample Size |  |  |  |
| Music CDs* | \$1.04 | 587 | \$1.17 | 4,033 | -0.030 | 45.25 | 0.000 |
| Movie DVDs | \$3.20 | 926 | \$1.87 | 7,775 | 0.104 | 45.20 | 0.000 |
| Video Games | \$9.01 | 659 | \$6.41 | 705 | 0.172 | 35.40 | 0.000 |
| Software | \$20.38 | 1,347 | \$11.56 | 4,182 | 0.104 | 27.42 | 0.000 |
| PDAs | \$31.66 | 710 | \$20.88 | 1,081 | 0.144 | 28.72 | 0.000 |
| Hard Drives | \$19.88 | 1,169 | \$16.47 | 5,783 | 0.023 | 25.14 | 0.000 |
| DVD Players | \$42.22 | 641 | \$22.21 | 952 | 0.197 | 23.97 | 0.000 |
| PC Monitors | \$53.71 | 1,436 | \$22.74 | 4,523 | 0.216 | 37.74 | 0.000 |
| Digital Cameras | \$48.29 | 1,899 | \$29.86 | 1,971 | 0.117 | 30.41 | 0.000 |
| Notebook PCs | \$126.22 | 344 | \$62.61 | 311 | 0.254 | 19.52 | 0.000 |
| Total | \$33.13 | 9,718 | \$12.53 | 31,316 | 0.175 | 69.50 | 0.000 |
| Average | \$38.56 |  | \$19.58 |  |  |  |  |
| Median | \$26.00 |  | \$18.70 |  |  |  |  |
| Note: Categories with unsupportive results are indicated by * and italic. Corr. is the correlation between 9 -ending prices and the size of price change. The $\boldsymbol{p}$-value is a significance level derived from a paired-sample t-test. Cross-category paired t-tests showed that the price changes are of a larger magnitude when prices end with $9\left(t_{9}=2.598, p=.01\right)$. |  |  |  |  |  |  |  |

Table 13. Average Size of Price Change in Internet Data: \$9.99- vs. Non-\$9.99-Endings

| Category | \$9.99-Ending |  | Non-\$9.99-Ending |  | Corr. | t-Stat | $p$-Value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean Price Change | Sample Size | Mean Price Change | Sample Size |  |  |  |
| Music CDs | \$2.52 | 76 | \$1.13 | 4,544 | 0.118 | 52.01 | 0.000 |
| Movie DVDs | \$5.82 | 188 | \$1.93 | 8,513 | 0.143 | 47.19 | 0.000 |
| Video Games | \$9.62 | 433 | \$6.75 | 931 | 0.176 | 36.21 | 0.000 |
| Software | \$22.93 | 186 | \$13.39 | 5,343 | 0.047 | 27.82 | 0.000 |
| PDAs | \$26.86 | 170 | \$24.97 | 1,621 | 0.015 | 29.02 | 0.000 |
| Hard Drives | \$32.40 | 335 | \$16.27 | 6,617 | 0.062 | 25.32 | 0.000 |
| DVD Players | \$48.23 | 219 | \$27.40 | 1,374 | 0.144 | 24.16 | 0.000 |
| PC Monitors | \$72.98 | 247 | \$28.35 | 5,712 | 0.145 | 37.95 | 0.000 |
| Digital Cameras | \$53.91 | 566 | \$36.34 | 3,304 | 0.079 | 30.67 | 0.000 |
| Notebook PCs | \$110.03 | 47 | \$94.93 | 608 | 0.031 | 19.59 | 0.000 |
| Total | \$36.24 | 2,467 | \$16.20 | 38,567 | 0.095 | 70.15 | 0.000 |
| Average | \$38.53 |  | \$25.15 |  |  |  |  |
| Median | \$29.60 |  | \$20.60 |  |  |  |  |

Note: Categories with unsupportive results are indicated by * and italic. Corr. is the correlation between 9 -ending prices and the size of price change. The $\boldsymbol{p}$-value is a significance level derived from a paired-sample t-test. Cross-category paired t-tests showed that the price changes are of a larger magnitude when prices end with $9\left(t_{9}=3.224, p=.005\right)$.

Table 14. Average Size of Price Change in Internet Data: \$99- vs. Non-\$99-Endings

| Category | \$99-Ending |  | Non-\$99-Ending |  | Corr. | t-Stat | $p$-Value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean Price Change | Sample Size | Mean Price Change | Sample Size |  |  |  |
| Music CDs | N/A | 0 | \$1.15 | 4,620 |  |  | NA |
| Movie DVDs | \$6.47 | 60 | \$1.98 | 8,641 | 0.094 | 47.40 | 0.000 |
| Video Games | N/A | 0 | \$7.66 | 1,669 |  |  | NA |
| Software | \$23.09 | 251 | \$13.26 | 5,278 | 0.056 | 27.80 | 0.000 |
| PDAs | \$46.75 | 122 | \$23.57 | 1,669 | 0.160 | 29.08 | 0.000 |
| Hard Drives | \$26.70 | 137 | \$16.85 | 6,815 | 0.024 | 25.36 | 0.000 |
| DVD Players | \$58.98 | 132 | \$27.66 | 1,461 | 0.173 | 24.21 | 0.000 |
| PC Monitors | \$98.11 | 332 | \$26.19 | 5,627 | 0.269 | 37.96 | 0.000 |
| Digital Cameras | \$85.88 | 476 | \$32.32 | 3,394 | 0.224 | 30.70 | 0.000 |
| Notebook PCs | \$144.42 | 161 | \$80.24 | 494 | 0.221 | 19.57 | 0.000 |
| Total | \$71.83 | 1,671 | \$15.10 | 39,363 | 0.224 | 70.26 | 0.000 |
| Average | \$61.30 |  | \$23.09 |  |  |  |  |
| Median | \$52.90 |  | \$20.20 |  |  |  |  |

Note: Categories with unsupportive results are indicated by * and italic. Corr. is the correlation between 9 -ending prices and the size of price change. The $\boldsymbol{p}$-value is a significance level derived from a paired-sample t-test. Cross-category paired t-tests showed that the price changes are of a larger magnitude when prices end with $9\left(t_{7}=3.598, p=.004\right)$.

Table 15. Average Size of Price Change in Internet Data: \$99.99- vs. Non-\$99.99-Endings

| Category | \$99.99-Ending |  | Non-\$99.99-Ending |  | Corr. | t-Stat | $p$-Value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean Price Change | Sample Size | Mean Price Change | Sample Size |  |  |  |
| Music CDs | N/A | 0 | \$1.15 | 4,620 |  |  | NA |
| Movie DVDs | \$12.01 | 24 | \$1.99 | 8,677 | 0.133 | 47.50 | 0.000 |
| Video Games | N/A | 0 | \$7.66 | 1,364 |  |  | NA |
| Software | \$20.06 | 37 | \$13.66 | 5,492 | 0.014 | 27.87 | 0.000 |
| PDAs | \$30.80 | 24 | \$25.04 | 1,757 | 0.021 | 29.10 | 0.000 |
| Hard Drives | \$34.75 | 36 | \$16.95 | 6,916 | 0.023 | 25.38 | 0.000 |
| DVD Players | \$73.51 | 56 | \$28.68 | 1,537 | 0.166 | 24.24 | 0.000 |
| PC Monitors | \$112.24 | 64 | \$29.31 | 5,895 | 0.139 | 37.98 | 0.000 |
| Digital Cameras | \$83.74 | 139 | \$37.24 | 3,731 | 0.110 | 30.75 | 0.000 |
| Notebook PCs | \$144.12 | 17 | \$94.73 | 638 | 0.063 | 19.60 | 0.000 |
| Total | \$70.56 | 407 | \$16.87 | 40,627 | 0.106 | 70.33 | 0.000 |
| Average | \$63.90 |  | \$25.64 |  |  |  |  |
| Median | \$54.10 |  | \$20.00 |  |  |  |  |
| Note: Categories with unsupportive results are indicated by * and italic. Corr. is the correlation between 9 -ending prices and the size of price change. The $\boldsymbol{p}$-value is a significance level derived from a paired-sample t-test. Cross-category paired t-tests showed that the price changes are of a larger magnitude when prices end with $9\left(t_{7}=3.390, p=.006\right)$. |  |  |  |  |  |  |  |

Figure 1. Frequency Distribution of the Last Digit in the Dominick's Data


Figure 2. Frequency Distribution of the Last Two Digits in the Dominick's Data


Figure 3. Frequency Distribution of the Last Digit in the Internet Data


Figure 4. Frequency Distribution of the Last Two Digits in the Internet Data


Figure 5. Frequency Distribution of the Last Dollar Digit in the Internet Data


Figure 6. Frequency Distribution of the Last Two Dollar Digits in the Internet Data


Figure 7. Frequency Distribution of the Price Changes in the Dominick's Data


Figure 8. Price of Frozen Concentrate Orange Juice, Heritage House, 12 oz (UPC = 3828190029, Store No. 78), September 14, 1989 - May 8, 1997


Figure 9. Demand Curve under Consumer Inattention


# Not for Publication - Reviewer's Appendix 

## Contents


#### Abstract

A. Results on Price Endings for Individual Product Categories in Dominick's and Internet Data

Similar to the aggregate results reported in the paper, the following figures show that $9 \mathbb{4}$ and $99 \Phi$ are the most popular price-endings for most of the individual product categories in both Dominick's and the Internet data.


Figures R1a-R1c. Frequency Distribution of the Last Digit by Product Category - Dominick’s Figures R2a-R2c. Frequency Distribution of the Last Two Digits by Product Category Dominick's

Figure R3. Frequency Distribution of the Last Digit by Product Category - Internet Data Figure R4. Frequency Distribution of the Last Two Digits by Product Category - Internet Data Figure R5. Frequency Distribution of the Last Dollar Digit by Product Category - Internet Data Figure R6. Frequency Distribution of the Last Two Dollar Digits by Product Category - Internet Data

## B. Results on Price Changes for Individual Product Categories in Dominick's Data

 Similar to the aggregate results reported in the paper and the results for individual product categories in our Internet data in Table 5, the following figures show that price changes in multiples of dimes are most common among all price changes in Dominick's data.Figures R7a-R7c. Frequency Distribution of the Price Changes by Category - Dominick’s

## C. Sample Price Series for Our Internet data

The following figures provide sample price series for ten randomly selected products, one from each of the ten product categories in our Internet data. All data are for 743 days, from March 26, 2005 to April 15, 2005.

Figure R8a. Price of a CD (Product\# 3, Store\# 194)
Figure R8b. Price of a DVD (Product\# 23, Store\# 194)
Figure R8c. Price of a Notebook PC (Product\# 422, Store\# 258)
Figure R8d. Price of a Hard Drive (Product\# 71, Store\# 324)
Figure R8e. Price of a DVD Player (Product\# 262, Store\# 230)
Figure R8f. Price of a Digital Camera (Product\# 273, Store\# 108)
Figure R8g. Price of a PC Monitor (Product\# 189, Store\# 17)
Figure R8h. Price of a PDA (Product\# 490, Store\# 207)
Figure R8i. Price of a Software (Product\# 96, Store\# 292)
Figure R8j. Price of a Video Game (Product\# 205, Store\# 68)

## D. Proportion of Price Changes that Preserve 9-Endings

The following tables report in detail the proportion of 9-ending preserving price changes. I.e., price changes of $10 \Phi, \$ 1, \$ 10, \$ 100$, etc. For the Dominick's data, in all but one category (Front-End-Candies), there are considerably more price changes that are multiples of dimes and dollars for 9-ending prices. For the Internet data, in the low-priced product categories (Music CDs, Movie DVDs, Video Games), we find considerably more price changes that are multiples of dimes and dollars for 9-ending prices. For high-priced product categories (DVD Players, PC Monitors, Digital Cameras, Notebook PCs), we find considerably more price changes that are multiples of $\$ 10$ and $\$ 100$ for 9-ending prices.

Table R1: Price Changes in Multiples of Dimes in Dominick’s Data: 9\$-Ending vs. Non-9\$Ending Prices

Table R2: Price Changes in Multiples of Dollars in Dominick’s Data: 99\$-Ending vs. Non-99\$Ending Prices

Table R3. Price Changes in Multiples of Dimes in Internet Data: 9¢-endings vs. Non-9థ-endings
Table R4. Price Changes in Multiples of Dollars in Internet Data: 99\$-Endings vs. Non-99థEndings

Table R5. Price Changes in Multiples of \$10 in Internet Data: \$9-Endings vs. Non-\$9-Endings
Table R6. Price Changes in Multiples of \$10 in Internet Data: \$9.99-Endings vs. Non-\$9.99Endings

Table R7. Price Changes in Multiples of \$100 in Internet Data: \$99-Endings vs. Non-\$99Endings

Table R8. Price Changes in Multiples of \$100 in Internet Data: \$99.99-endings vs. Non-\$99.99endings

Figure R1a. Frequency Distribution of the Last Digit by Product Category - Dominick's











Figure R1b. Frequency Distribution of the Last Digit by Product Category - Dominick's

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Figure R1c. Frequency Distribution of the Last Digit by Product Category - Dominick's










Figure R2a. Frequency Distribution of the Last Two Digits by Product Category - Dominick's











Figure R2b. Frequency Distribution of the Last Two Digits by Product Category - Dominick's







Figure R2c. Frequency Distribution of the Last Two Digits by Product Category - Dominick's


Figure R3. Frequency Distribution of the Last Digit by Product Category - Internet Data










Figure R4. Frequency Distribution of the Last Two Digits by Product Category - Internet Data


Figure R5. Frequency Distribution of the Last Dollar Digit by Product Category - Internet Data











Figure R6. Frequency Distribution of the Last Two Dollar Digits by Product Category - Internet Data









Figure R7a. Frequency Distribution of the Price Changes by Category - Dominick's











Figure R7b. Frequency Distribution of the Price Changes by Category - Dominick's











Figure R7c. Frequency Distribution of the Price Changes by Category - Dominick's










Figure R8a. Price of a CD (Product \#3, Store \#194) 743 Days (March 26, 2003 -April 15, 2005)


Figure R8b. Price of a DVD (Product \#23, Store \#194)
743 Days (March 26, 2003 - April 15, 2005)


Figure R8c. Price of a Notebook PC (Product \#422, Store \#258) 743 Days (March 26, 2003 - April 15, 2005)


Figure R8d. Price of a Hard Drive (Product \#71, Store \#324)
743 Days (March 26, 2003 - April 15, 2005)


Figure R8e. Price of a DVD Player (Product \#262, Store \#230) 743 Days (March 26, 2003 - April 15, 2005)


Figure R8f. Price of a Digital Camera (Product \#273, Store \#108) 743 Days (March 26, 2003 - April 15, 2005)


Figure R8g. Price of a PC Monitor (Product \#189, Store \#17) 743 Days (March 26, 2003 - April 15, 2005)


Figure R8h. Price of a PDA (Product \#490, Store \#207)
743 Days (March 26, 2003 - April 15, 2005)


Figure R8i. Price of a Software Product (Product \#96, Store \#292)
743 Days (March 26, 2003 - April 15, 2005)


Figure R8j. Price of a Video Game (Product \#205, Store \#68)
743 Days (March 26, 2003 - April 15, 2005)


Table R1: Price Changes in Multiples of Dimes in Dominick's: 9¢- vs. Non-9థ-Ending Prices

| Category | 9¢-Ending |  | Non-9¢-Ending |  | $p$-Value |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Multiples of Dimes | Sample Size | Multiples of Dimes | Sample Size |  |
| Analgesics | 78.25\% | 367,969 | 5.60\% | 102,550 | . 0000 |
| Bath Soap | 74.93\% | 58,735 | 12.65\% | 18,298 | . 0000 |
| Bathroom Tissues | 47.97\% | 156,863 | 4.09\% | 184,414 | . 0000 |
| Bottled Juices | 42.10\% | 457,490 | 5.33\% | 583,025 | . 0000 |
| Canned Soup | 26.14\% | 304,439 | 4.12\% | 741,357 | . 0000 |
| Canned Tuna | 36.10\% | 170,023 | 6.15\% | 281,703 | . 0000 |
| Cereals | 37.21\% | 271,757 | 8.32\% | 494,597 | . 0000 |
| Cheeses | 46.49\% | 872,489 | 4.57\% | 1,039,738 | . 0000 |
| Cookies | 58.73\% | 1,135,112 | 9.01\% | 709,697 | . 0000 |
| Crackers | 46.99\% | 283,278 | 7.31\% | 279,353 | . 0000 |
| Dish Detergent | 56.10\% | 240,532 | 4.75\% | 183,222 | . 0000 |
| Fabric Softeners | 51.41\% | 212,288 | 5.96\% | 191,319 | . 0000 |
| Front-end-candies | 18.47\% | 137,453 | 11.66\% | 385,234 | . 0000 |
| Frozen Dinners | 32.72\% | 230,423 | 5.70\% | 336,201 | . 0000 |
| Frozen Entrees | 42.49\% | 883,284 | 5.93\% | 1,183,557 | . 0000 |
| Frozen Juices | 46.75\% | 301,114 | 5.40\% | 395,344 | . 0000 |
| Grooming Products | 71.30\% | 1,017,513 | 10.22\% | 287,969 | . 0000 |
| Laundry Detergents | 68.07\% | 446,767 | 4.68\% | 210,342 | . 0000 |
| Oatmeal | 36.27\% | 72,753 | 7.17\% | 107,971 | . 0000 |
| Paper Towels | 37.01\% | 109,596 | 4.26\% | 152,846 | . 0000 |
| Refrigerated Juices | 46.25\% | 405,144 | 4.59\% | 418,402 | . 0000 |
| Shampoos | 80.84\% | 1,916,061 | 29.23\% | 238,976 | . 0000 |
| Snack Crackers | 48.53\% | 488,341 | 4.61\% | 405,005 | . 0000 |
| Soaps | 48.23\% | 180,935 | 4.79\% | 190,632 | . 0000 |
| Soft Drinks | 76.54\% | 4,614,455 | 15.36\% | 1,219,151 | . 0000 |
| Tooth Brushes | 74.22\% | 350,705 | 2.46\% | 123,840 | . 0000 |
| Tooth Pastes | 61.64\% | 468,688 | 6.18\% | 291,045 | . 0000 |
| Total | 62.81\% | 16,154,207 | 7.64\% | 10,755,788 | . 0000 |
| Note: Categories with unsupportive results are indicated by * and italic. p-Value is an asymptotic significance level derived from Pearson $\chi^{2}$ test. |  |  |  |  |  |

Table R2: Price Changes in Multiples of Dollars in Dominick's: 99¢- vs. Non-99¢-Ending Prices

| Category | 99¢-Ending |  | Non-99¢-Ending |  | $p$-Value |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Multiples of Dollars | Sample Size | Multiples of Dollars | Sample Size |  |
| Analgesics | 17.09\% | 106,038 | 1.39\% | 364,481 | . 0000 |
| Bath Soap | 21.06\% | 15,608 | 3.11\% | 61,425 | . 0000 |
| Bathroom Tissues | 1.66\% | 36,944 | 0.04\% | 304,333 | . 0000 |
| Bottled Juices | 2.02\% | 104,451 | 0.27\% | 936,064 | . 0000 |
| Canned Soup | 0.19\% | 56,527 | 0.01\% | 989,269 | . 0000 |
| Canned Tuna | 2.96\% | 19,566 | 0.03\% | 432,160 | . 0000 |
| Cereals | 6.60\% | 56,437 | 0.99\% | 709,917 | . 0000 |
| Cheeses | 3.03\% | 160,237 | 0.16\% | 1,751,990 | . 0000 |
| Cookies | 5.41\% | 270,448 | 1.01\% | 1,574,361 | . 0000 |
| Crackers | 9.79\% | 62,297 | 0.06\% | 500,334 | . 0000 |
| Dish Detergent | 1.83\% | 52,117 | 0.22\% | 371,637 | . 0000 |
| Fabric Softeners | 10.67\% | 62,370 | 0.31\% | 341,237 | . 0000 |
| Front-end-candies* | 0.00\% | 11,923 | 0.01\% | 510,764 | . 1887 |
| Frozen Dinners | 3.38\% | 56,617 | 0.65\% | 510,007 | . 0000 |
| Frozen Entrees | 8.47\% | 188,496 | 0.53\% | 1,878,345 | . 0000 |
| Frozen Juices | 0.21\% | 67,862 | 0.04\% | 628,596 | . 0000 |
| Grooming Products | 5.21\% | 247,298 | 1.63\% | 1,058,184 | . 0000 |
| Laundry Detergents | 20.15\% | 158,974 | 2.53\% | 498,135 | . 0000 |
| Oatmeal | 1.28\% | 12,921 | 0.82\% | 167,806 | . 0000 |
| Paper Towels | 8.38\% | 15,137 | 0.03\% | 247,305 | . 0000 |
| Refrigerated Juices | 4.76\% | 101,063 | 0.25\% | 722,522 | . 0000 |
| Shampoos | 12.99\% | 503,157 | 5.86\% | 1,651,880 | . 0000 |
| Snack Crackers | 3.23\% | 97,690 | 0.13\% | 795,656 | . 0000 |
| Soaps | 4.43\% | 43,874 | 0.20\% | 327,693 | . 0000 |
| Soft Drinks | 12.87\% | 1,385,935 | 2.86\% | 4,447,671 | . 0000 |
| Tooth Brushes | 19.06\% | 108,407 | 0.89\% | 366,138 | . 0000 |
| Tooth Pastes | 4.85\% | 117,086 | 0.57\% | 642,647 | . 0000 |
| Total | 9.86\% | 4,119,480 | 1.39\% | 22,790,515 | . 0000 |
| Note: Categories with unsupportive results are indicated by * and italic. p-Value is an asymptotic significance level derived from Pearson $\chi^{2}$ test. |  |  |  |  |  |

Table R3. Price Changes in Multiples of Dimes in Internet: 9¢- vs. Non-94-endings

|  | 9世-Endings |  | Non-9¢-Endings |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Category | Multiples of <br> Dimes | Sample <br> Size | Multiples of <br> Dimes | Sample <br> Size | p-Value |
| Music CDs | $73.32 \%$ | 2,268 | $21.17 \%$ | 2,352 |  |
| Movie DVDs | $66.90 \%$ | 2,813 | $23.08 \%$ | 5,888 | .0000 |
| Video Games | $80.05 \%$ | 832 | $44.17 \%$ | 532 | .0000 |
| Software* | $57.32 \%$ | 778 | $60.43 \%$ | 4,751 | .1015 |
| PDAs | $66.76 \%$ | 355 | $59.40 \%$ | 1,436 | .0110 |
| Hard Drives | $74.36 \%$ | 1,435 | $57.39 \%$ | 5,517 | .0000 |
| DVD Players* | $57.18 \%$ | 383 | $59.83 \%$ | 1,210 | .3569 |
| PC Monitors* | $47.71 \%$ | 809 | $56.08 \%$ | 5,150 | .0000 |
| Digital Cameras* | $72.77 \%$ | 852 | $77.07 \%$ | 3,018 | .0093 |
| Notebook PCs* | $73.91 \%$ | 92 | $78.51 \%$ | 563 | .3250 |
| Total | $68.32 \%$ | 10,617 | $50.50 \%$ | 30,417 | .0000 |
| Note |  |  |  |  |  |

Note: Categories with unsupportive results are indicated by * and italic. $\boldsymbol{p}$-Value is an asymptotic significance level derived from Pearson $\chi^{2}$ test.

Table R4. Price Changes in Multiples of Dollars in Internet: 99¢- vs. Non-99¢-Endings

|  | 99థ-Endings |  | Non-99¢-Endings |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Category | Multiples of <br> Dollars | Sample <br> Size | Multiples of <br> Dollars | Sample <br> Size | p-Value |
| Music CDs | $62.43 \%$ | 1,142 | $5.69 \%$ | 3,478 |  |
| Movie DVDs | $72.19 \%$ | 1,532 | $6.89 \%$ | 7,169 | .0000 |
| Video Games | $77.69 \%$ | 744 | $33.71 \%$ | 620 | .0000 |
| Software | $56.42 \%$ | 553 | $50.18 \%$ | 4,976 | .0054 |
| PDAs | $70.33 \%$ | 300 | $52.45 \%$ | 1,491 | .0000 |
| Hard Drives | $84.95 \%$ | 1,083 | $45.14 \%$ | 5,869 | .0000 |
| DVD Players | $59.27 \%$ | 329 | $50.08 \%$ | 1,264 | .0030 |
| PC Monitors* | $47.98 \%$ | 544 | $47.17 \%$ | 5,415 | .7174 |
| Digital Cameras* | $65.02 \%$ | 852 | $74.12 \%$ | 3,018 | .0000 |
| Notebook PCs | $84.38 \%$ | 64 | $72.76 \%$ | 591 | .0444 |
| Total | $69.13 \%$ | 7,056 | $37.40 \%$ | 33,978 | .0000 |
| Nannnn |  |  |  |  |  |

Note: Categories with unsupportive results are indicated by * and italic. p-Value is an asymptotic significance level derived from Pearson $\chi^{2}$ test.

Table R5. Price Changes in Multiples of $\$ 10$ in Internet: $\$ 9$ - vs. Non-\$9-Endings

| Category | \$9-Endings |  | Non-\$9-Endings |  | $p$-Value |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \text { Multiples of } \\ \$ 10 \end{gathered}$ | Sample Size | $\begin{gathered} \text { Multiples of } \\ \$ 10 \end{gathered}$ | Sample Size |  |
| Music CDs* | 0.00\% | 587 | 0.25\% | 4,033 | . 2271 |
| Movie DVDs | 2.92\% | 926 | 0.35\% | 7,775 | . 0000 |
| Video Games | 32.78\% | 659 | 11.99\% | 705 | . 0000 |
| Software | 29.62\% | 1,347 | 3.25\% | 4,182 | . 0000 |
| PDAs | 43.38\% | 710 | 4.07\% | 1,081 | . 0000 |
| Hard Drives | 22.50\% | 1,169 | 2.11\% | 5,783 | . 0000 |
| DVD Players | 33.23\% | 641 | 7.35\% | 952 | . 0000 |
| PC Monitors | 33.43\% | 1,436 | 4.13\% | 4,523 | . 0000 |
| Digital Cameras | 48.98\% | 1,899 | 9.84\% | 1,971 | . 0000 |
| Notebook PCs | 74.13\% | 344 | 19.29\% | 311 | . 0000 |
| Total | 31.65\% | 9,718 | 2.76\% | 31,316 | . 0000 |
| Note: Categories with unsupportive results are indicated by * and italic. p-Value is an asymptotic significance level derived from Pearson $\chi^{2}$ test. |  |  |  |  |  |

Table R6. Price Changes in Multiples of $\mathbf{\$ 1 0}$ in Internet: \$9.99- vs. Non-\$9.99-Endings

| Category | \$9.99-Endings |  | Non-\$9.99-Endings |  | $p$-Value |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Multiples of \$10 | Sample Size | Multiples of \$10 | Sample Size |  |
| Music CDs* | 0.00\% | 76 | 0.22\% | 4,544 | . 6822 |
| Movie DVDs | 11.70\% | 188 | 0.38\% | 8,513 | . 0000 |
| Video Games | 42.26\% | 433 | 5.05\% | 931 | . 0000 |
| Software | 44.62\% | 186 | 8.46\% | 5,343 | . 0000 |
| PDAs | 38.82\% | 170 | 17.64\% | 1,621 | . 0000 |
| Hard Drives | 50.45\% | 335 | 3.26\% | 6,617 | . 0000 |
| DVD Players | 42.47\% | 219 | 13.83\% | 1,374 | . 0000 |
| PC Monitors | 34.41\% | 247 | 10.19\% | 5,712 | . 0000 |
| Digital Cameras | 55.48\% | 566 | 24.06\% | 3,304 | . 0000 |
| Notebook PCs | 78.72\% | 47 | 9.63\% | 608 | . 0000 |
| Total | 42.64\% | 2,467 | 7.49\% | 38,567 | . 0000 |

Note: Categories with unsupportive results are indicated by * and italic. p-Value is an asymptotic significance level derived from Pearson $\chi^{2}$ test.

Table R7. Price Changes in Multiples of \$100 in Internet: \$99- vs. Non-\$99-Endings

| Category | \$99-Endings |  | Non-\$99-Endings |  | $p$-Value |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \hline \text { Multiples of } \\ \$ 100 \\ \hline \end{gathered}$ | Sample Size | $\begin{gathered} \text { Multiples of } \\ \$ 100 \end{gathered}$ | Sample Size |  |
| Music CDs | $N / A$ |  |  |  |  |
| Movie DVDs |  |  |  |  |  |  |
| Video Games |  |  |  |  |  |  |
| Software | 1.59\% | 251 | 0.23\% | 5,278 | . 0000 |
| PDAs | 10.66\% | 122 | 0.30\% | 1,669 | . 0000 |
| Hard Drives* | 0.00\% | 197 | 0.06\% | 6815 | . 7993 |
| DVD Players | 6.06\% | 132 | 0.41\% | 1,461 | . 0000 |
| PC Monitors | 15.36\% | 332 | 0.32\% | 5,627 | . 0000 |
| Digital Cameras | 19.12\% | 476 | 0.77\% | 3,394 | . 0000 |
| Notebook PCs | 38.51\% | 161 | 6.07\% | 494 | . 0000 |
| Total | 13.70\% | 1,671 | 0.26\% | 39,363 | . 0000 |

Note: Categories with unsupportive results are indicated by * and italic. p-Value is an asymptotic significance level derived from Pearson $\chi^{2}$ test.

Table R8. Price Changes in Multiples of $\mathbf{\$ 1 0 0}$ in Internet: $\$ 99.99$ - vs. Non- $\$ 99.99$-endings

| Category | \$99.99-Endings |  | Non-\$99.99-Endings |  | $p$-Value |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Multiples of \$100 | Sample Size | Multiples of \$100 | Sample Size |  |
| Music CDs | $N / A$ |  |  |  |  |
| Movie DVDs |  |  |  |  |  |  |
| Video Games |  |  |  |  |  |  |
| Software* | 0.00\% | 37 | 0.29\% | 5,492 | . 7423 |
| PDAs* | 2.94\% | 34 | 0.97\% | 1,757 | . 2531 |
| Hard Drives* | 0.00\% | 36 | 0.06\% | 6,916 | . 8852 |
| DVD Players | 8.93\% | 56 | 0.59\% | 1,537 | . 0000 |
| PC Monitors | 12.50\% | 64 | 1.03\% | 5,895 | . 0000 |
| Digital Cameras | 14.39\% | 139 | 2.60\% | 3,731 | . 0000 |
| Notebook PCs | 41.18\% | 17 | 13.32\% | 638 | . 0011 |
| Total | 10.07\% | 407 | 0.71\% | 40,627 | . 0000 |

Note: Categories with unsupportive results are indicated by * and italic. p-Value is an asymptotic significance level derived from Pearson $\chi^{2}$ test.


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[^1]:    ${ }^{1}$ See, for example, Blinder, et al. (1998), Carlton (1986), Cecchetti (1986), Caplin (1993), Warner and Barsky (1995), Lach and Tsiddon (1996), Slade (1998), Ball and Romer (1990, 2003), Davis and Hamilton (2004), Fisher and Konieczny (2000, 2006), Konieczny and Skrzypacz (2005), and Rotemberg (1987, 2005). For recent surveys, see Willis (2003) and Wolman (2007).
    ${ }^{2}$ See also the studies of the European Central Bank's Inflation Persistence Network, e.g., Fabiani, et al. (2006).

[^2]:    ${ }^{3}$ See, e.g., Ball, Mankiw and Reis (2005), Mankiw and Reis (2002), Reis (2006a, 2006b), and Woodford (2003).

[^3]:    ${ }^{4}$ In case of missing data because the sellers' web sites were inaccessible or the price information was not available, then we used the following procedure. If $10 \%$ or more observations were missing for a product, then that series was excluded from the data altogether. If less than $10 \%$ of the data was missing, then we examined if the prices for the day before and the day after were the same. If they were the same, then the software agent automatically filled in for the missing data with that price. Otherwise, the agent filled in for the missing data with the price for the day after. We recognize that this is an arbitrary procedure. However, there are only $0.075 \%$ missing prices in the entire dataset, and thus missing data are unlikely to affect our results significantly.
    ${ }^{5}$ The categories were selected based on their popularity on the Internet. In addition, the products in these categories are sold by a large number of stores. For example, in the category of digital cameras, "Canon-EOS Digital Rebel XT" is sold by 63 stores. The selection of products was random. For example, in the category of DVDs, we chose products from multiple sub-categories (e.g., action, drama, comedy, etc.). Similarly, in the music CDs category, we chose from many different sub-categories (e.g., blues, jazz, country, etc.). However, in some categories (e.g., notebook PCs and hard drives), we included all the products available. In other categories (e.g., DVD players, digital cameras, PC monitors, software), we randomly chose products from all sub-categories. For example, in DVD players, we chose half of the products from standard DVD players while the other half came from the more expensive DVD/VCR combo players. In digital cameras and camcorders, we chose half from regular digital cameras while the other half came from digital camcorders. In PC monitors, we chose half from CRT and flat CRT models, and the other half from LCD and TFT. In the software category, we chose products from multiple genres of software (e.g., educational software, operating systems, programming software, utility software, etc.). Similarly, in video games, we included multiple genres (adventure, action, sports, etc.). See Figures R8a-R8i in the reviewer's appendix for sample price series.

[^4]:    ${ }^{6}$ Indeed, according to Dutta, et al. (1999) and Levy, et al. (1997, 1998), the average price of an item in large U.S. supermarket chains during 1991-92 was about \$1.70. According to Bergen, et al. (2008), the figure increased to $\$ 2.08$ by 2001.
    ${ }^{7}$ To save space, most of figures and tables on individual product categories are included in a separate reviewer's appendix to this paper. We shall note that the results for individual product categories are similar to the aggregate results we report here.
    ${ }^{8}$ The products in the Beer and Cigarettes categories are highly regulated that could potentially skew the results (Besley and Rosen, 1999, footnote 6). We, therefore, do not discuss the results of their analyses.
    ${ }^{9}$ Benford's Law, also known as the Significant Digit Law, predicts that in many naturally occurring settings such as tables, measurements, etc., the distribution of the leftmost digits is logarithmic, and not uniform as one would expect. See, for example, Varian (1972) and Hill (1995). For example, the probability of 1 occurring as the leftmost digit is $\log _{10} 2 \approx 0.301$, the probability of 2 occurring as the leftmost digit is $\log _{10}(3 / 2) \approx 0.176$, etc. This surprising fact was discovered in 1881 by Newcomb (1881), who noticed that the pages of logarithm tables containing numbers starting with 1 were more worn out than the other pages. In 1938 Benford (1938) studied over 20,000 different data sets, including areas of rivers, baseball statistics, numbers in magazine articles, and the street addresses of the first 342 people listed in the book American Men of Science and concluded that these indeed obeyed the Law. Under Benford's Law, the probabilities of the digits tend to being uniformly distributed as we move from left to right. For the second digits the skew is from 12 percent for the 0 down to 8.5 percent for the 9 . Nigrini (2002, Ch. 7) shows that as a first approximation, one can argue that the last-two digits are equally likely for each combination from 00 to 99 in three-digit and higher numbers. Therefore, the distribution of the rightmost digits that we find in our data cannot be explained by Benford's Law. We thank Mark Nigrini for his insight on this.
    ${ }^{10}$ With the exception of five categories (canned tuna, cigarettes, front-end-candies, oatmeal, and paper towels), the 99\& ending prices are the most common than other two-digit ending prices. Even in the five categories where the 994 ending is not the most popular, it is still very common and ranks in the top five price-endings among the 100 possible endings. See Figures R2a-R2c in the reviewer's appendix.

[^5]:    ${ }^{11}$ The results on the use of 9 for the last three and four digits in the Internet data, and some of the results in the section on price changes in the Internet data are presented only in tables as they are too numerous to be plotted.
    ${ }^{12}$ Note that there are 1,000 possible endings here.
    ${ }^{13}$ Note that there are 10,000 possible endings here.
    ${ }^{14}$ For results on individual product categories in the Internet data, see the reviewer's appendix.

[^6]:    ${ }^{15}$ Category level data indicate some cross-category variation, although in general they are consistent with the above finding. I.e., in most categories, price changes in multiples of 10 cents are more common than other price changes. See the reviewer appendix.

[^7]:    ${ }^{16}$ For hard drives, changes in multiples of ten dollars are the fourth most popular category. For CDs and DVDs, they are not ranked in the top 20, because the prices for both products are low and thus the price changes rarely reach $\$ 10$.

[^8]:    ${ }^{17}$ We should note, however, that the internet price series seem to have far fewer sales and promotions. Indeed, inspection of the internet price series suggests that there are not many cases of temporary price decreases which are reversed after two-three weeks. See, for example, the sample time series shown in Figures R8a-R8j, in the Referee Appendix.

[^9]:    ${ }^{18}$ Landsburg (1995) describes the historical origins of 9-ending prices. See also Ginzberg (1936).
    ${ }^{19}$ See Shugan (1980), Ball and Mankiw (1994), Ball (2000), Mankiw and Reis (2002), Ameriks, et al. (2003, 2004), Zbaracki, et al (2004), Ball, et al (2005), Rotemberg (2003, 2005, 2008), Reis (2006a, 2006b), Klenow and Willis (2007), and Knotek (2006).
    ${ }^{20}$ This is known as the place-value principle (Debaene, 1997). For example, each one of the three digits that make up number 999 signifies different magnitude because of their different location in the number, even though the three digits are identical. This principle applies only to Arabic numerals. It does not apply, for example to Roman numerals.
    ${ }^{21}$ The argument holds even if the marginal cost remains constant because marginal benefit declines as we move from left to right.
    ${ }^{22}$ This is consistent with recent laboratory experiment findings that people tend to drop the rightmost digit in processing price information (Bizer and Schindler, 2005). This kind of selective consumer inattention to price information is consistent with evidence from surveys of consumer behavior in this industry (Progressive Grocer, November 1974, p. 39 and Progressive Grocer, February 1964, pp. C104-C106, as cited by Gabor and Granger (1961) and Carlton and Perloff (2000)). This behavior is consistent also with the marketing literature on "just noticeable differences" in consumer behavior (Monroe, 1970, 2001), where consumers do not react to small price changes because they do not "notice" them (Kalyanaram and Little 1994).

[^10]:    ${ }^{23}$ The optimality of inattention to the last digit is not universal. For example, the strategy of ignoring the last digit as a timesaving device will not be very efficient in processing such numeric information as phone numbers, social security numbers, etc. ${ }^{24}$ Consistent with this idea is a recent study by Chen, et al. (2008), which also uses the Dominick's data and finds more frequent "small" price increases than decreases, for price changes of up to about 10¢. After ruling out standard models of price adjustment or inflation as explanation, they argue that the asymmetry might be due to consumer inattention.

[^11]:    ${ }^{25}$ The phenomenon of 9-ending prices has also received considerable attention in the marketing literature, where most studies explain the 9 -ending pricing phenomena on psychological grounds. Our explanation shares their emphasis on behavioral considerations for pricing phenomena. Most of these explanations, however, rely on some form of irrational behavior, making them more difficult to incorporate into economic analyses. For example, according to Nagle and Holden (1995, p. 300), buyers perceive the 9 -ending prices "... as significantly lower than the slightly higher round numbers that they approximate." As another example, Schindler and Kirby (1997) posit that consumers might perceive a 9-ending price as a round-number price with a small amount given back. Other theories argue that sellers like to give change or that buyers like to receive a change. It has been suggested also that 9 -ending prices may be interpreted as discount prices and thus are indicative of good bargains. Finally some authors note the cognitive accessibility of certain numbers, such as 0 and 5 , to explain pricing points. See Shapiro (1968) and Monroe (1990) for reviews of earlier literature. Basu (1997, 2006), Anderson and Simester (2003), and Ruffle and Shtudiner (2006) provide reviews of more recent literature. Rotemberg (2008) contains critical analyses of these and other related studies. ${ }^{26}$ We shall note an important caveat. While our data suggests that 9 -ending prices tend to be significantly more rigid than the non-9-ending prices, it is not clear that this rigidity, by itself, is necessarily an indicator of monetary non-neutrality. It could be that the retailers are actually charging the maximum flexible price and merely round the price up to the nearest 9 (if that price does not already end with 9). In such a world money would have no systematic effect on output, even though nominal prices are sticky. The average price level, in that case, would approximately equal the flexible price level plus half the gap between the price points (the average distance from the frictionless price to the next price point). In this sense, our finding of the 9-ending dominance suggests that retailers price on a grid (see, e.g., Genesove, 2003).
    ${ }^{27}$ USA Today has reported that "France, Spain and Britain quit producing low-denomination coins in recent decades because production costs kept going up while the coins' purchasing power went down" (Copeland 2001). More recently, it has been reported that in many EU countries which have adopted the Euro, the public seems to be exhibiting resistance to the use of 1-cent and 2-cent denomination coins. This is due to the inconvenience their use entails. The International Herald Tribune reports that these coins are "small, nearly valueless-and a nuisance to millions of Europeans. The tiny denomination 1- and 2-cent Euro coins are annoying shoppers and disrupting business from Paris to Milan" (Pfanner 2002, p. 1). In 2001, Rep. Jim Kolbe (RArizona) introduced the "Legal Tender Modernization Act," to make the U.S. penny obsolete. The bill was defeated. Previous attempts made in 1990 and 1996 also died in Congress (Copeland 2001). Recently, CBC featured an article on October 10, 2007, on "A 'penniless' Canada," and reported that Australia and New Zealand have already eliminated their pennies.
    ${ }^{28}$ See Bils and Klenow (2004), Levy and Young (2004), and Campbell and Eden (2005).

[^12]:    ${ }^{29}$ See, for example, Ehrmann (2005) and Hoffmann and Kurz-Kim (2009), and the studies cited therein.
    ${ }^{30}$ The use of price points might have a strong normative component. We are unlikely to see 9 -ending prices in certain settings. For example, imagine the patients' reaction if the dentist tells them that "A tooth filling costs $\$ 79.99$ - it's today's Special!" Or how about our reaction if we are considering corrective eye surgery, and the eye doctor's office manager tells us: "First eye - full price; the second eye - $50 \%$ off."
    ${ }^{31}$ See Fengler and Winter (2001), Ratfai (2003), Mostacci and Sabbatini (2003), and Konieczny and Rumler (2007).
    ${ }^{32}$ According to Heeler and Nguyen (2001), in the Chinese culture, numbers have special significance and symbolism. Even the sounds of the numbers can suggest good or bad luck. For example, the number 8 represents luck to Cantonese Chinese because it sounds like multiply or get rich ( $f a$ in Cantonese). In Japan, 8 also has great symbolic significance because the writing of the number 8 looks like a mountain ("八"), and thus the number 8 signifies growth and prosperity. Heeler and Nguyen (2001) find that close to $50 \%$ of restaurant menu prices sampled in Hong Kong had 8 -endings, which they refer to as "happy endings." Also, a Time Magazine article (Rawe, 2004) reports that at the casino of a recently-built $\$ 240$ million hotel, Sands Macao in Macao, China, the slot machines' winning trios of 7's have been replaced with trios of 8's. Consistent with these observations, note that the opening ceremony of the Beijing Olympic Games, held in the Beijing National Stadium, began exactly at 08:08:08pm on 8/8/2008. The cultural importance of numbers is not limited to "happy endings." For example, according to Mirhadi (2000), when the Masquerade Tower was added to Hotel Rio in Las Vegas in 1997, the architects decided to skip the $40^{\text {th }}$ to the $49^{\text {th }}$ floors because the Arabic numeral " 4 " in Chinese sounds similar to the word "death." The elevators in the building went directly from the $39^{\text {th }}$ floor to the $50^{\text {th }}$ floor.
    ${ }^{33}$ Additional analyses (not reported here to save space) show that 9 is indeed more rigid than any other digit in our datasets. Other popular digits in our data (e.g., 5), do not consistently lead to more price rigidity, and even when they do, the rigidity associated with them is considerably less compared to that associated with 9.

